# 2005 Stock Assessment Report for Atlantic Striped Bass: 

Catch-at-Age Based VPA \& Tag Release/Recovery Based Survival Estimation



A report prepared by the
Striped Bass Technical Committee for the Atlantic Striped Bass Management Board

November 2005


Healthy, self-sustaining populations of all Atlantic coast fish species or successful restoration well in progress by the year 2015

# 2005 Atlantic Striped Bass Advisory Report 

## Status of the Stock

Fishing Mortality Rates: Based on VPA results, average age 8-11 fishing mortality in 2004 is estimated at $\mathrm{F}=0.40$ which is below the Amendment 6 threshold of 0.41 but exceeds the target of 0.30. However, it is the consensus of the Technical Committee members that this is likely an overestimate of the 2004 F given the uncertainly with the terminal year estimate from the VPA and the systematic positive bias observed in the retrospective analysis. The 2003 value of F from this year's VPA is 0.29 , which is substantially lower than the terminal year F from last years VPA run of 0.62 . This is due not only to the addition of another years worth of data but also due to the modified suite of tuning indices used in the this years VPA and the inclusion of wave 1 (Jan./Feb.) estimates of recreational harvest mortality from NC and VA for 1996 - 2004 (see Data and Uncertainty section below).

The 2004 tagged based estimates of F using, stock-specific, model-based estimates of fishing mortality and a constant M of 0.15 , were as follows. For fish greater than 28 inches, the coastwide average F was estimated as 0.29 and specific tagging program values ranged from 0.02 in the New York ocean haul survey (NYOHS) to 0.31 in the Maryland (MD) tagging program. This value was similar to the VPA F weighted by N value for age 7-11 fish of 0.32 . For fish greater than 18 inches, the coast-wide average F was 0.29 and specific tagging program values ranging from 0.06 in the Virginia spawning stock (VARAP) program to 0.68 in the New Jersey Delaware Bay (NJDEL) program. This tag-based F estimate was greater than the VPA F weighted by N value for age 3-11 fish of 0.15 .

The 2004 tagged based estimates of F using, stock-specific, catch-equations for fish greater than 28 inches, indicated the coast-wide average F was 0.14 , and specific tagging program values ranged from 0.09 in the VARAP program to 0.26 in the Delaware and Pennsylvania (DE-PA) tagging program. These F estimates were less than the VPA F weight by N, for age 7-11 fish, of 0.32. For fish greater than 18 inches, the coast-wide average was 0.11 , and specific tagging program $F$ estimates that ranged from 0.05 in three different programs to 0.17 in the MD program. This tag-based F estimate is similar to the VPA F weighted by N value for age 3-11 fish of 0.15 .

Chesapeake Bay fishing mortality in 2004 is estimated as $\mathrm{F}=0.16$ by the direct enumeration study. This F represents mortality during the June 2003 - June 2004 period, so it is not directly comparable to the average, weighted (by N) VPA calendar-year F on age 3-8 striped bass that is equal to 0.12 .

Exploitation Rates: Based on the tagging programs, R/M estimates produced by 5 (VARAP, NCOOP, MD, NYHUD, and DE/PA) out of 8 programs have shown a decline in exploitation rates since the late 1990's. During the same period, the NYOHS and MA tagging programs have showed no trend and the NJDEL program has shown an increase in exploitation in recent years.

Stock Size: The estimate of total abundance for January 1, 2005 from the ADAPT VPA is 65.3 million age-1 and older fish. This estimate is about 1.2 million fish lower than the 2004 abundance but $10 \%$ higher than the average stock size for the previous five years. Population estimates were calculated for the first time this year from tag based F estimates using the catch equation. The 2004 population estimate of age $3+$ fish was 48.5 million fish that is roughly 8 million fish higher than the 2003 estimate. This estimate is higher than the ADAPT VPA estimate of 39.2 million age $3+$ fish at the beginning of 2004 .

The abundance of older fish (age 13+ from the ADAPT VPA) in the stock has also increased from 382,000 fish at the beginning of 2003 to 547,000 fish on January 1, 2005.

Spawning Stock Biomass (SSB): The female spawning stock biomass for 2004 is estimated at 55 million pounds which is above the recommended biomass threshold of 30.9 millions pounds $(13,956 \mathrm{mt})$ and the target SSB of 38.6 million pounds $(17,500 \mathrm{mt})$. SSB has declined by $9 \%$ since 2002 when it peaked at 60.6 million pounds.

Recruitment: Recruitment of the 2004 cohort for all stocks combined is 12.7 million age- 1 fish, which is close to the average age- 1 recruitment observed since the stocks were declared recovered in 1995

Catch: Total catch in numbers including landings and discards increased from 3.9 million fish in 2002 to 5.2 million fish in 2004, a $33.3 \%$ increase in losses since implementation of Amendment 6. The 2004 catch was also above the 1997-2003 average of 4.36 million fish. Ages 3 to 7 represented $59 \%$, and ages $8+$ represented $36 \%$ of the total catch in 2004. The strong 1996, 2000, and 2001 year-classes dominated the catch, accounting for $41 \%$ of total catch. Total catch of age 8+ fish increased from 1.4 million fish in 2002 to 1.8 million fish in 2004 (the highest level recorded in the time series) and the proportion of $8+$ fish in the catch increased to $36 \%$ in 2004 from $30 \%$ in 2003.

Recreational harvest ( 2.4 million fish) and discards ( 17.2 million fish) accounted for $72.5 \%$ of the total 2004 catch. Virginia recreational fisheries harvested $19.6 \%$ of total recreational landings, followed by New Jersey (17.7\%), Massachusetts (17.1\%), Maryland (13.2\%), North Carolina ( $13.2 \%$ ), and NY ( $10.2 \%$ ). The remaining states each landed $5 \%$ or less of the total recreational landings.

Estimates of Wave 1 (January-February) recreational harvest in North Carolina and Virginia from 1996-2004 were included in the catch at age for the first time this year. The estimates ranged from 7,544 in 2000 to 177,288 fish during 2004 in North Carolina and 5,985 fish in 1996 to 155,616 fish in 2004 in Virginia. These Wave 1 harvest estimates represented between 2\% and $14 \%$ of the total coast-wide recreational harvest during those years.

Commercial harvest ( 0.91 million fish) and discards ( 0.51 million fish) accounted for $27.5 \%$ of the total 2004 catch. Maryland commercial fisheries harvested $50.8 \%$ of the total commercial landings, followed by VA (16.3\%), PRFC (10.1\%), NY (7.8\%), and MA (6.7\%). The remaining states each landed $4 \%$ or less of the total commercial landings.

Data and Uncertainty: A formal review of abundance indices used in former assessments was initiated by ASMFC at a workshop in July of 2004. This workshop developed a set of evaluation criteria (Appendix A) and tasked states with a review of indices. The resulting review led to a revision and elimination of some indices formerly used in the ADAPT VPA. Both the Striped Bass Technical Committee and the Management Board approved of the criteria and of the review. The indices underwent further review based on residual patterns following initial model runs. This is a standard annual procedure that led to the elimination of additional indices for the 2005 analysis.

A winter fishery (January-February) for striped bass has developed off of North Carolina and Virginia since the mid-1990's. MRFSS estimates are not available from this time of year in Virginia and are only available for 2004 in North Carolina. Landings were estimated for these fisheries back to 1996 using observed relationships between landings and tag returns. These estimates were included in the catch at age matrix of the ADAPT VPA for the first time this year.

A variety of concerns were expressed by some members of the Technical Committee concerning input data for the assessment including the accuracy of aging older fish, the methods used to estimate commercial and recreational discards, the methods used to estimate NC and VA recreational harvest in Wave 1 dating back to 1996 and about the MRFSS estimates in general.

Uncertainties expressed by some members of the Technical Committee concerning the ADAPT VPA model include potential violations of some of the model assumptions such as the assumption that the catch at age is measured with out error. Concerns about the model output included the validity of bootstrap generated error estimates for terminal year F as calculated by ADAPT, the significant discrepancies between VPA estimates using old and new indices, and the retrospective bias (positive for F; negative for SSB) in the terminal year estimate that was apparent in most VPA runs for striped bass over the past few years. Some members felt that a correction to the terminal year estimate of F should be made using the average bias shown in this year's VPA run. However, other Technical Committee members were concerned about doing this because the direction and magnitude of the bias could change in next years VPA run.

Most Technical Committee members expressed the need for a more current estimate of the tag reporting rate used in the tag based estimates. The estimate currently being used is 0.43 and was based on a study in 1999 conducted on the Delaware River spawning stock. If the 1999 estimate is higher than the current tag reporting rate, the exploitation rate and the F estimate are underestimated. If the rate is lower than the current reporting rate, then $F$ estimates are overestimated. A research grant proposal is currently in submission to conduct a coast-wide high reward tagging study to develop a more current estimate of the reporting rate that applies to a wider geographical area. Some TC members suggested this type of study should be conducted at regular intervals (e.g. every 3 years).

Concerns mentioned about the survival estimates from the Brownie models included the variability of the year specific estimates of survival depending on the most recent year of reported tag returns that were included in the analysis. Some TC members mentioned concern that the assumption of mixing and dispersal was not being adequately met. Others felt that
concern had been addressed by an analysis of the Virginia Rappahannock tag data by John Hoenig that indicated only very minor violations of the assumption of complete mixing, which did not affect the results of the analysis.

There is concern expressed by some TC members about the use of a constant value of natural mortality ( M ) despite the presence of analyses suggesting an increase in $M$ in Chesapeake Bay in recent years. To address this concern, the Tag Committee used the catch equation method that allows for development of estimates of F without the use of a constant M value. The TC expressed the need for variance estimates for the F values from the catch equation method and this will be addressed in 2006. Some TC members expressed uncertainties about the recent reduction in the exploitation rate estimates used in the catch equation since the adoption of Amendment 6 in 2003 that showed a 10-25\% decline in exploitation despite a $33 \%$ increase in the total commercial and recreational losses (harvest plus discards) during the past two years. Others felt concerned about moving forward with the use of the catch equation method before further exploration concerning potential non-mixing of newly tagged animals was conducted.

## Management Advice

Based on the available assessment information, it is the consensus of the Technical Committee that overfishing is not occurring and that the population is not overfished. However, there are differing opinions within the Technical Committee concerning where the 2004 fish mortality rate was in relation to the Amendment 6 target of 0.30 . It is also the consensus of the Technical Committee that the abundance of older striped bass, age 13 and older, has increased since the adoption of Amendment 6 in 2003.

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## I. Introduction

This report summarizes results of catch-age based virtual population analyses (VPA) of Atlantic striped bass for 2004. The VPA analysis provides estimates of fishing mortality, stock abundance, and biomass for the mixed coastal stock.

The first analytical assessment of Atlantic striped bass stocks using VPA was conducted in 1997 for years 1982-1996 and reviewed by the $26^{\text {th }}$ Stock Assessment Review Committee at the Northeast Fisheries Science Center. The results of the review were reported in the proceedings of the $26^{\text {th }}$ Northeast Regional Stock Assessment Workshop ( $26^{\text {th }}$ SAW): SARC Consensus Summary of Assessments (NEFSC Ref. Document 98-03). The assessment methodology utilized NEFSC ADAPT version of VPA and remained unchanged until 2002. The stock status and assessment procedures were reviewed once more at the 36th SAW in December 2002.

A formal review of abundance indices used in former assessments was initiated by ASMFC at a workshop in July of 2004 (ASMFC 2004). This workshop developed a set of evaluation criteria (Appendix A) and tasked states with a review of indices. The resulting review led to a revision and elimination of some indices formerly used in ADAPT. Both the Striped Bass Technical Committee and the Management Board approved of the criteria and of the review. The indices underwent further review based on residual patterns following initial model runs. This is a standard annual procedure that led to the elimination of additional indices for the 2005 analysis.

## II. Catch-at-Age Virtual Population Analysis

## Data Summary

Catch at age was estimated using standard methods described in the previous assessment documents (ASMFC 2002). Commercial landings at age were estimated by applying corresponding length-frequency distributions and age-length keys to the reported number of fish landed by the commercial fishery in each state. Length-frequencies of recreational landings were based on a combination of MRFSS length samples and volunteer angler logbooks. Length frequencies of recreational discards were based on volunteer angler logbook and American Littoral Society data. State specific age-length keys were applied, where possible, to length frequencies to estimate number of fish at age landed or discarded by the recreational fishery. State specific methods for estimating age composition of commercial landings, recreational landings, and recreational discards are provided in individual state compliance reports to ASMFC for 2004. State specific data sources for estimates of recreational discard age composition are also summarized in Table 1a

Commercial Fishery in 2004
Commercial landings in 2004 totaled 907 thousand fish or 3.3 thousand MT ( 7.2 million lbs) (Table 1b). Landings increased $4.4 \%$ in numbers (38 thousand fish) and $2.2 \%$ in weight ( 70 MT ) compared to 2003. Overall, commercial harvest represented $18 \%$ of total losses in number in 2004 (Table 2, Figure 1). The greatest portion of the commercial harvest occurred in the Chesapeake Region (Maryland, PRFC, and Virginia). The harvest in these jurisdictions accounted for $77 \%$ by number (Table 3 ) and $58 \%$ by weight of the total commercial harvest in
2004. Harvest increased in all coastal states with commercial fisheries except Virginia and Delaware (Table 3). Age 4 made up the highest percentage of commercial landings ( $21 \%$ ) and ages $4-8$ comprised $68 \%$ of the harvest (Table 4). Most (77\%) of the harvest in the Chesapeake Region was ages 3-7 (Table 4, Figure 2). Peak harvest of fish in the rest of the coastal states was at age 8; more than half of the coastal harvest (54\%) was ages 8-10.

Direct measurements of commercial discards of striped bass in 2004 were only available for fisheries in the Hudson River Estuary. Discard estimates for fisheries in Chesapeake Bay and coastal locations since 1982 and for Delaware Bay in 2004, were based on the ratio of tags reported from discarded fish in the commercial fishery to tags reported from discarded fish in the recreational fishery, scaled by total recreational discards:

$$
\mathrm{CD}=\mathrm{RD}^{*}(\mathrm{CT} / \mathrm{RT})
$$

where:
$C D=$ unadjusted estimate of the number of fish discarded by commercial fishery, $R D=$ number of fish discarded by recreational fishery, estimates provided by the NOAA Marine Recreational Fisheries Survey (MRFSS).
$\mathrm{CT}=$ number of tags returned from discarded fish by commercial fishermen, $\mathrm{RT}=$ number of tags returned from discarded fish by recreational fishermen.

Total discards are allocated to fishing gears based on the relative number of tags recovered by each gear. Discards by fishing gear were multiplied by gear specific release mortalities and summed to estimate total number of fish killed in a given year. Tag return data and release mortality by gear for 2004 are given in Table 5. Starting in 1998, the Technical Committee attempted to improve the estimate of commercial discards by calculating tag return ratios and discards separately for Chesapeake Bay and the coast. A separate estimate for Delaware Bay was added in 2004. The ratio of tags from fish discarded by commercial fishermen to tags returned from fish discarded by recreational fishermen in 2004 was 0.47 in Chesapeake Bay, 0.12 in Delaware Bay, and 0.04 along the coast (ME - NC)(Table 6).

Expanding recreational discards to commercial discards based on reported tag returns assumes equal reporting tag rates in commercial and recreational fisheries. To evaluate this assumption we examined the ratio of tags recovered by commercial and recreational fisheries for landed fish. If the availability of tagged fish to commercial and recreational fisheries is equal, the ratio of tags recovered by commercial and recreational fisheries should be close to the ratio of landings. This was not the case suggesting a lower reporting rate by the commercial fishery in some locations and years (Table 6). To correct for this bias, we calculated a correction factor by dividing the three-year mean of ratios of commercial to recreational landings by the three-year mean of ratios of tags returned by the two fisheries. Since only one year of data for Delaware Bay was available, we used the mean of the correction factors for the coast and Chesapeake Bay. The correction factors for 2004 were 1.41 for Chesapeake Bay, 1.77 for the coast (Table 6), and 1.59 for Delaware Bay.

In summary, we calculated commercial discard losses for all fisheries except those in the Hudson River by multiplying recreational discards by the commercial/recreational tag ratio from discarded fish, then by the corresponding correction factor, apportioning discards among gears,
and finally by multiplying by appropriate gear specific discard mortalities. Total commercial discards for 2004 were estimated as 519 thousand fish, representing $10.0 \%$ of total removals in number (Table 2, Figure 1).

Commercial discard proportions at age were obtained by applying age distributions from fishery dependent sampling or independent surveys using comparable gear (Table 7a). Gear specific proportions at age were applied to discard estimates by gear and expanded estimates summed across all gears. Most commercial discards were fish of ages 3-8 (Table 7b). Discards were higher in 2004 than in 2003 and the third highest since 1982 (Figure 3).

Total commercial striped bass removals (landings and discards) were 1.43 million fish in 2004 (Table 2). Although total removals in 2004 exceeded those in 2003, they remain below the peak in 1997 (Figure 4). Landings have generally exceeded discards since the mid 1990's (Figure 3). Commercial losses in 2004 was dominated by age 4 (2000 year class) fish (Figure 5).

## Recreational Fishery in 2004

Recreational statistics were collected as part of the MRFSS (Marine Recreational Fishery Statistics Survey) program. Details of the assessment methodology can be found on the MRFSS web site (http://www.st.nmfs.gov/st1/recreational/the_mrfss.html). MRFSS did not sample in January and February (wave-1) prior to 2004 when sampling began in North Carolina waters. Therefore, there was little information for the winter fishery (Jan, Feb) that has developed off of North Carolina and Virginia since the mid 1990's. We estimated landings for these fisheries back to 1996 using observed relationships between landings and tag returns (Appendix B). For North Carolina, we used the ratio of estimated landings to tag returns in wave-1 of 2004 and annual tag returns in wave- 1 to estimate annual landings in January and February. For Virginia waters, we used the 1996-2004 mean ratio of landings and tag returns in wave-6 and annual tag returns in wave-1 to estimate landings in January and February. Methods and results are summarized in Appendix C.

We estimated age composition of the January/February recreational fishery in North Carolina and Virginia from length-frequency data collected by MRFSS and appropriate state age-length keys. Length-frequencies for the North Carolina winter harvest of 2004 came from data in wave6 of 2003 and wave-1 of 2004. That for the winter harvests of 1996-2003 came from wave-6 of year t-1. We converted lengths to age for North Carolina with a combined age-length key from New York and North Carolina. Length-frequencies for the Virginia winter harvest in 1996-2004 came from MRFSS data in wave-6 of year $\mathrm{t}-1$. We converted the Virginia lengths to age with a Virginia age-length key. Estimates of wave-1 harvest at age for North Carolina and Virginia were added to the existing CAA matrix for 1996 through 2004. We did not estimate discards for the winter recreational fishery in North Carolina or Virginia.

Total landings in 2004 (MRFSS A+B1 and estimated winter landings) were estimated at 2.4 million fish totaling 11.9 thousand MT ( 26.1 million pounds) (Table 1b). Landings decreased slightly compared to 2003 (Table 1b). Overall, recreational harvest represented $46.0 \%$ by number of all losses in 2004 (Table 2, Figure 1). The states with the highest landings were Massachusetts, New York, New Jersey, Maryland, Virginia, and North Carolina (Table 8). Landings in Virginia made up 19.6 \% of the total and were the highest of all states. Striped bass
ages six through 10 comprised 62.6\% of landings (Table 9, Figure 6). Highest landings occurred for age eight (1996 year class) which made up $16.8 \%$ of the total (Figure 6). Fish harvested in the recreational fishery were generally larger than those harvested in the commercial fishery (Figure 7).

Recreational discards (B2) increased in 2004 to 17.2 million fish (Table 2) compared to 14.6 million fish in 2003. Discard losses due to hooking mortality ( $0.08 *$ released fish) were estimated at 1.4 million fish in 2004 (Table 2). The states with the greatest number of discards were Massachusetts and Maryland (Table 10). Recreational discards represented $27 \%$ by number of total losses (Table 2). Discard losses of the 2001 year class (Age 3) were the highest ( $38.5 \%$ ) among all cohorts in 2004 (Table 10, Figure 6)

Total recreational striped bass removals (landings and discards) in 2004 were 3.76 million fish (Table 11). Total removals were highest in Massachusetts, New Jersey, Maryland, and Virginia. The catch was dominated by ages 3,4 , and $8(41.8 \%$ of total) (Figure 8$)$. Total recreational discard and landings losses have generally increased since 1982, with intermittent declines in 1998-1999 and 2001-2002 (Figure 9). Recreational removals in 2004 were the highest of the time series. The proportion of recreational removals caused by discards has remained relatively stable since 2001(Figure 10).

## Total Catch at Age

The above components were totaled by year to produce the overall catch at age matrix for VPA input (Table 12). The total removal of striped bass in 2004 was 5.2 million fish and reflects a 7.2 $\%$ and a $33 \%$ increase from 2003 and 2002. More importantly, removals of fish age $8+$ increased in 2004 by $28.5 \%$ and a 32.9 \% compared to 2003 and 2002. Total removals in 2004 were the highest since 1982 (Figure 11). Ages three, four, and eight sustained the highest losses in 2004 (Figure 12). Ages 5 and 7 comprised the greatest proportion of the catch in 2003.

## Weight at Age

Catch weight at age information was updated for the period 1998-2002 using all available weight data from MA, NY, MD, VA, NH, and CT (1998-2001) and adding data from RI and DE in 2002 (Appendix D). Mean weights at age for the 2003 and 2004 striped bass catches were determined as a result of the expansion of catch and weight at age. Data came from Maine and New Hampshire recreational harvest and discards; Massachusetts recreational and commercial catch; Rhode Island recreational and commercial catch, Connecticut recreational catch, New York recreational catch and commercial landings; New Jersey recreational catch; and Delaware, Maryland, Virginia, and North Carolina recreational and commercial catch. Weighted mean weights at age were calculated as the sum of weight at age multiplied by the catch at age in numbers, divided by the sum of catch at age in numbers. Details of developing weights at age for 1982 to 1996 can be found in NEFSC Lab Ref. 98-03. Weights at age for 1982-2004 are presented in Table 13.

## Survey Indices

The ADAPT model requires indices of abundance to be measured either at the beginning or the middle of the year. Consequently, indices from surveys conducted in the spring were assigned
sampling date of January 1. Indices measured in summer were assigned to the middle of the year, and those collected in the fall were assigned to the January 1 of the following year with their age increased by one. All juvenile survey indices were advanced forward to the January 1 of the following year and the index was assigned age 1

Based on criteria developed at the VPA indices workshop and the recommendations by the Striped Bass Technical Committee, we made significant changes to many survey indices available for input into the VPA in 2004. The NEFSC spring inshore survey was reduced from age-specific indices to an aggregate index, and was truncated at 1991 due to missed sampling of inshore survey strata prior to 1991. The Massachusetts commercial age-specific harvest-per-trip indices were redeveloped as age-specific (ages 5-13+) total catch-per-hour indices. The New Jersey trawl aggregate index was further apportioned into age-specific geometric mean indices for age 2-9+. Due to large proportional standard errors of the New York ocean haul seine survey indices for age $>9$, the $13+$ age-specific index was aggregated to a $9+$ group. The Virginia pound net survey was eliminated from the input because few analyses conducted could support its continued use as an index that reflected striped bass abundance. Two new indices were added to the input: age-specific (ages 2-10+) Delaware River spawning stock indices and a coast-wide MRFSS aggregate index. The MRFSS index was based on data only from private boats that fished in the ocean during waves 3-6 (Appendix E). There were no changes made to the Connecticut aggregate trawl index, Connecticut age-specific recreational catch indices, the Maryland spawning stock age-specific indices or any indices for YOY (age 0) in Maryland, Virginia, New Jersey, and New York, or for juveniles (age 1) in Maryland and Long Island, New York. The changes resulted in a total of 62 indices for use in initial runs of ADAPT (Table 14).

Among the fisheries-dependent indices, trends in the MA Commercial indices and CT Recreational CPUE suggest steady population levels since the mid 90s, while the coast-wide MRFSS index suggests a decline since 1998 (Figure 13).

The fishery-independent indices for combined ages generally show a stable, high level of population abundance punctuated by strong year classes (Figure 13). The strong 1993, 1996 and 2001 year classes contributed to the annual variability in the NY, DE, NJ and NEFSC survey results. There was fair correspondence between the NJ and DE trawl surveys (Figure 14).

Indices of young-of-the-year show low to moderate recruitment in the Chesapeake Bay, Delaware Bay, and the Hudson River in 2004 (Figure 15). The high 2003 MD and VA index continues as age one in 2004. The high numbers of age one striped bass in the Western Long Island survey in 2004 suggests the possibility that there was high survival of the 2003 year class in New York coastal waters (Figure 15).

## ADAPT Virtual Population Analysis

Catch at Age and Indices
Initial runs of ADAPT for the 2005 assessment used a combination of 62 age-specific and age aggregated fishery independent and fishery dependent indices discussed above and in Table 14. Residual plots showed systematic trends in residuals for some survey indices and this led to a
rejection of the MA commercial catch per hour indices for ages 8-13, MD spawning stock indices for ages 3 and 4, the DE trawl index, and the DE spawning stock index for age 2. Furthermore, the MA commercial indices failure to track strong year classes provided additional justification for exclusion from analysis. The remaining 52 indices were used in the final run of ADAPT. Indices included the MA commercial catch per hour indices ages 5-7, MD SSB index for ages 7-13+, NY Ocean Haul seine ages 3-8 and aggregated for 9=13, CT CPUE for ages 2-9, NEFSC aggregated for ages 2-9, young-of-year (age 0) in Maryland, Virginia, New York and New Jersey, age 1 index for Maryland and Long Island, New York, CT trawl aggregated for ages 4-6, DE spawning stock for ages 4-9, and aggregated for 10-13, the NJ trawl index for ages 2-8 and aggregated for 9-13, and a MRFSS index for aggregated ages 2-13 (Table 14).

The 2003 assessment (through fishing year 2002) concluded that the 13+ age configuration of the ADAPT model produced the most accurate estimates of $F$ and stock size in the presence of age error/bias in the catch-at-age and survey indices (Striped Bass Stock Assessment Committee 2003). This configuration was continued for the 2004 and 2005 assessments.

An iterative re-weighting of the survey indices was applied to the model.

## Partial Recruitment Vector

A flat top partial recruitment vector was assumed for the ADAPT model. Initial PR values were calculated using the three year geometric mean fishing mortality for each age from the previous ADAPT model scaled to the highest value of F among all ages.

## Model Configuration

This year's ADAPT run used the same input options as last year's assessment: full F in terminal year was calculated using classic method; F at oldest true age for all years, including terminal year was calculated using Heincke's method and ages 8 through 11 were used to calculate the oldest true age. Plus group abundance was calculated using the backward method and the model assumed a flat topped partial recruitment.

## ADAPT Results

## Fishing Mortality

The 2004 average fishing mortality rate ( F ) for fully recruited ages 8 through 11 equaled 0.40 and was above the current target ( 0.30 )(Table 15, Figure 16). This represents a dramatic drop in F on fully recruited ages from that reported for 2003 (reported as $F=0.62$ in 2004, SBSASC 2004). However, this may reflect the shift in model indices or the addition of winter harvest estimates for NC and VA. The 2003 value of F in the current run was 0.29 suggesting an increase in 2004. Fishing mortality in 2004 on ages $3-8$, which are generally targeted in producer areas, was $\mathrm{F}=0.16$. Among the individual age groups, the highest value of $\mathrm{F}(0.50)$ was estimated for 9 year old fish (1996 year class) (Table 16, Figure 17). Estimates of F in 2004 were generally higher for ages eight and above than for younger ages. We did not include bootstrap generated error estimates for terminal year F values because we have concerns about validity of such estimates as calculated by ADAPT. An F weighted by N was calculated for comparison to tagging results in 2004 since the tag releases and recaptures are weighted by abundance as part of
the experimental design. The 2004 VPA F weighted by N for ages $7-11$ (age 7 to compare with tagged fish $>28$ ") was 0.32 (Table 15). An F weighted by N for ages 3-8, comparable to the direct enumeration estimate for Chesapeake Bay, was equal to 0.12 (Table 15).

The iterative re-weight option used in ADAPT applies extra weight to those indices which have the best model fit. The indices assigned the highest weights were the CT CPUE ages $4-9$, the CT trawl aggregate index, the Delaware spawning stock indices, the MA commercial indices, and the MRFSS index (Table 17).

## Population Abundance (January 1)

Striped bass abundance increased steadily from 1982 through 1997 when it reached a level around 60 million fish (Table 18a, Figure 18). Total abundance declined to 54 million fish in 2000 and then increased to 65 million fish in 2005. The 2003 cohort remained strong at 19 million fish at age 2 in 2005 and exceeded the size of the strong 1993 and 2001year classes at 2. Estimates of abundance obtained this year were higher than those reported in 2004 (SBSAC 2004). Error estimates for abundance at age for 2005 were lowest for ages 7-9 (Table 18b).

Abundance of striped bass age $8+$ increased steadily through 2002 to 6.6 million. It has since fluctuated without obvious trend around 6.2 million fish through the present (Table 18a, Figure 18). The 1 Jan 2005 estimate was 5.9 million fish.

## Spawning Stock Biomass

Female spawning stock biomass (SSB) grew steadily from 1982 through 2002 when it peaked at about 27 thousand metric tons (Table19, Figure 19). Female SSB has declined since then and was estimated at 24.9 thousand metric tons in 2004, assuming 1:1 male- female ratio. The estimated SSB remained above the threshold level of 1995 , which was estimated as 14.6 thousand metric tons. Again, values obtained in the 2005 analysis exceeded those obtained in 2004 (SBSAC 2004).

## Retrospective Patterns

A retrospective analysis was conducted on the VPA results extending back to 1999, in order to determine trends in estimation of F, total abundance, recruitment, and female SSB in the terminal year. The retrospective evaluation was made using the iterative re-weighting option, which assumes the chi-weights from the terminal year estimate are equivalent in all subsequent years. The analysis revealed that average fishing mortality estimates for ages $8-11$ were overestimated prior to 2003 (Figure 20a). However, the terminal year estimate for 2003 was identical to that obtained for 2003 made the next year. There was no significant bias in terminal year estimates of total abundance (Figure 20b) or recruitment (Figure 21a). A slightly negative bias occurred in terminal year estimates of female SSB (Figure21b)

## Sensitivity Runs

Sensitivity runs made in the 2004 assessment (ASMFC 2004) indicated that the model was relatively insensitive to the inclusion or exclusion of indices. This year however, the use of revised indices led to a dramatic change in estimates of population parameters compared to those
made in 2004. For comparative purposes, we made an ADAPT run using last year's indices updated for 2004. Use of last year's configuration of indices resulted in higher estimates of F and lower estimates of age 8+ abundance from the mid 1990's through the present (Figure 22). The estimate of F for 2004 using last year's configuration of indices was 0.67 suggesting an increase in F over 2003. Divergence in estimates using new and old configuration of indices increased through the time period.

## Sources of Uncertainty

The ADAPT virtual population analysis model used in this assessment assumes that the catch at age input data are measured without error, the recruitment vector is constant after the age of full recruitment, and that changes in abundance indices reflect changes in population abundance. All of these assumptions may be violated to some degree as used for striped bass.

Accurate estimates of catch at age require that we know the total loss in number and that we apportion this loss correctly to age. The best data on loss comes from the directed recreational and commercial fisheries. The exception in this year's assessment was estimates of harvest in the winter fishery that has developed off of North Carolina and Virginia. MRFSS data were generally not available for this time of year and we estimated harvest for these fisheries using relationships between harvest and tag returns. There is less confident in estimates of discard losses in commercial and recreational fisheries because little of the data is measured directly. Moreover, gear specific release mortalities are assumed to be constant even though mortalities may vary with season and with changes in gear specifics such as increased use of circle hooks. The quality of data on age composition varies among fisheries and region. In most cases, fish in catches or discards are measured and length frequencies are converted to age frequencies with age length keys. States with large harvests usually sample fisheries directly and develop age length keys from the fishery and time of year of the fishery. However, states with small fisheries must often rely on length data from small samples or fishery independent collections and use age length keys developed by neighboring jurisdictions. Finally, the assignment of age to samples becomes less certain with increasing fish age. The ADAPT runs made last year (ASMFC 2004) were sensitive to large changes ( $40 \%$ ) in the catch at age input. The addition the winter harvest in this year's analyses also affected the outcome.

The abundance indices used this year's analysis were improved through a reasoned and objective evaluation process described in ASMFC 2004 and in Appendix A. The review reduced the number of indices and the number of indices at age, especially for fish age eight and older. This year's ADAPT VPA analysis was highly sensitive to the selection of indices, especially to those for the older ages. There is clearly a need to develop additional indices of abundance for older fish in the fished subset of the population.

Estimates of F and population size from the catch at age analyses employed for striped bass are most uncertain for the terminal year. Retrospective analyses conducted in prior striped bass assessments usually suggested a positive bias in the terminal year estimates of $F$ and a negative bias in terminal estimates of population size. Although similar results were obtained this year, bias was less, especially for the 2003 terminal year estimate. It is possible that the bias has become less of a problem with improved accounting of losses to the population and improved abundance indices.

## Summary

The striped bass population remains at high level of abundance due, in part, to strong incoming cohorts. The fully exploited population abundance (age 8+) decreased since last year, but has been relatively stable since about 2001. Average fishing mortality for fully recruited ages (8-11) in 2004 was estimated at 0.40 . The F estimate for 2003 was 0.29 which is much lower than the F for the same year (0.62) estimated in the 2004 assessment (SBSASC 2004). However, this difference is due, in part, to a change in tuning indices and the addition of winter harvest in NC and VA. Estimates of F increased from 2003 to 2004 in ADAPT outputs for both the new and the old indices. The 2004 fully recruited fishing mortality estimate is above the target of 0.3. Average fishing mortality for ages 7-11 weighted by N was 0.32 and for ages $3-8$ weighted by N was 0.12 . Spawning stock biomass has decreased from levels in 2002 and 2003, but remains well above the 1995 threshold level.

## III. Tagging Program Analyses

## Introduction

This report summarizes the results of analysis by the ASMFC Striped Bass Tagging Subcommittee (SBTS) of tagging data from the U.S.F.W.S. Cooperative Striped Bass Tagging Program through the 2004 tagging year. These results now include two different sets of estimates of instantaneous fishing mortality ( F ) rates, one of which is based on the protocol previously employed by the SBTS (Smith et al. 2000), where we employ tag recovery models to estimate annual survival; survival is then converted to total instantaneous mortality, Z. Estimates of survival are corrected for bias due to live release of striped bass, because the tag recovery models assume all recoveries are of dead animals (Smith et al. 2000). The final step is subtraction of an assumed constant value of natural mortality, M, to estimate F.

The new protocol, introduced into our report for the fist time, is based on a formulation of Baranov's catch equation in Ricker (1975) and was proposed by Pollock et al. (1991). Crecco (2003) first applied this method to the striped bass tag results, as well as to combinations of tag and virtual population estimates. In this protocol, we do not assume a constant value of M. Instead, F is estimated as a function of both Z and the exploitation rate, $\mu$. Following F estimation, M is estimated by subtraction of F from Z . Also presented are length structure of tagged striped bass, age structure of recaptures, geographic distributions of recaptures by month, and estimates of catch and exploitation rates by program.

A second change in the report is that we have added a new regulatory period to our period models, extending them from 3 periods to 4 periods. The new period is based on Addendum 1 to Amendment 6, which began in 2000, with a goal of reducing F on larger fish. Analysis of this change was conducted in advance by V. Vecchio, NY DEC, for the SBTS. The new period provided generally better fits to the tag-recovery data than the previous 3 period models.

Finally, we present two time series of Atlantic coastwide total abundance estimates for age 3+ striped bass, and two time series of estimates of age 7+ striped bass. These are based on the form of the catch equation: Kill $=\mathrm{F}$ * (average N ). One series is produced using the F estimates
generated assuming constant M , and the other set of estimates was based on the F series produced via the catch equation.

## Description of Tagging Programs

Eight tagging programs provided information for this report, and have been in progress for at least 11 years. Most producer area and coastal programs tag striped bass (mostly $>=18$ inches total length) during routine state monitoring programs. Producer area tagging programs operate mainly during spring spawning, and use many capture gears, such as pound nets, gill nets, seines and electroshocking. Producer area programs are as follows: 1. Delaware and Pennsylvania (DE-PA) with fish tagged primarily in April and May, 2. Hudson River (HUDSON) with fish tagged in May, 3. Maryland (MDDNR) with fish tagged primarily in April and May, and 4. Virginia spawning stock program (VARAP) with fish tagged in the Rappahannock River during April and May. Coastal programs tag striped bass from mixed stocks during fall, winter, or early spring and use several gears including hook \& line, seine, gill net, and otter trawl. The coastal tagging programs are as follows: 1. Massachusetts (MADFW) with fish tagged during fall months, 2. North Carolina winter trawl survey (NCCOOP) with fish tagged primarily in January, 3. New Jersey Delaware Bay (NJDEL) with fish tagged in March and April, and 4. New York ocean haul survey (NYOHS) with fish tagged during fall months.

Tag release and recapture data are exchanged between the U.S. Fish and Wildlife Service (USFWS) office in Annapolis, MD, and the cooperating tagging agencies. The USFWS maintains the tag release/recovery database and provides rewards to fishermen who report the recapture of tagged fish. Through July of 2004, a total of 426,576 striped bass have been tagged and released, with 75,930 recaptures reported and recorded in the USFWS database (Tina McCrobie, personal comm.).

## Data Analysis

The Striped Bass Tagging Committee's analysis protocol is based on assumptions described in Brownie et al. (1985) and elaborated for striped bass in Smith et al. (2000). The tag recovery data is analyzed in program MARK (White, 1999). Important assumptions of the tagging programs (as reported in Brownie 1985) are as follows:

1. The sample is representative of the target population.
2. There is no tag loss.
3. Survival rates are not affected by the tagging itself.
4. The year of tag recoveries is correctly tabulated.

Other assumptions related to the modeling component of the analyses include:
5. The fate of each tagged fish is independent of the fate of other tagged fish.
6. The fate of a given tagged fish is a multinomial random variable.
7. All tagged individuals of an identifiable class (age, sex) in the sample have the same annual survival and recovery rates.

The analysis protocol follows an information-theoretic approach based on Kullback-Leibler information theory and Akaike's information criterion (Burnham and Anderson 2003), and involves the following steps. First, a set of biologically-reasonable candidate models are
identified prior to analysis (Table 20; see section on Justification of candidate models). Various patterns of survival and recovery are used to parameterize the candidate models. These models allow parameters to be constant, time specific, or allow time to be modeled as a continuous variable. Other models allow time periods to coincide with changes in regulatory regimes.

## Justification of candidate models

Candidate models (selected before analysis) are based on biologically-reasonable hypotheses. The global model $\{\mathrm{S}(\mathrm{t}) \mathrm{r}(\mathrm{t})$, i.e., full parameterized model $\}$ is a time saturated model, and is used to estimate over-dispersion and model fit statistics (see section on Diagnostic procedures). Models that parameterize survival as constant within time periods \{S(p)r(p), $S(p) r(t), S(d) r(p)$, and $\mathrm{S}(\mathrm{v}) \mathrm{r}(\mathrm{p})$ \} are based on regulatory changes within the time series (1987-2004). Four regulatory periods are defined as follows: moratorium years (1987-1989), an interim fishery (1990-1994), a full fishery under Amendment 5 (1995-1999) and the recent changes introduced in 2000, which were designed to reduce F on older fish (2000-2004). Given the importance of recent years for management, we also model the terminal year separately $\{\mathrm{S}(\mathrm{d}) \mathrm{r}(\mathrm{p})\}$ and the most recent two years separately $\{\mathrm{S}(\mathrm{v}) \mathrm{r}(\mathrm{p})\}$. The Virginia tagging program models an additional period-specific model (1990-1992, 1993-1994, 1995-2003). Although changes within the striped bass fishery are addressed with time and period-specific models, we believe that constant models are also reasonable. Selection of a constant model $\{\mathrm{S}() .\mathrm{r}(),. \mathrm{S}() .\mathrm{r}(\mathrm{p}), \mathrm{S}() .\mathrm{r}(\mathrm{t})\}$ does not mean "no" variation in survival across the time series, but suggests that year-to-year variation in annual survival is "...relatively small in relation to the information contained in the sample data" (Burnham and Anderson 2003).

Models parameterized with covariates are also included within the candidate set. Selection of models with time as a covariate within regulatory periods $\{\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp}), \mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{t}), \mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})\}$ support increasing or decreasing monotonic trends in survival within survival. These models are reasonable given increases in fishing effort during the time series. There is a concern that trend models may over or underestimate the terminal year estimate of survival, and analyses of simulated data are needed to address this issue.

## Diagnostic procedures

Model adequacy is a major concern when deriving inference from a model or a suite of models. Over-dispersion, inadequate data (such as low sample size), or poor model structure may cause a lack of model fit. Over-dispersion is expected in striped bass tagging data, given that a lack of independence may result from schooling behavior. If over-dispersion is detected, then an estimate of the variance inflation factor (i.e., c-hat) is used to adjust AICc (after adjustment, AICc is called QAICc; Anderson et al (1994)). We estimate c-hat by dividing the observed Pearson Chi-square value (goodness-of-fit statistic of the global model) by the expected Pearson Chi-square value (derived from a bootstrap analysis of the global model). The goodness-of-fit probability of the global model is examined with a bootstrap-derived p-value based on model deviance (Burnham and Anderson 2003). A low p-value ( $<0.15$ ) and a large estimate of c-hat ( $>$ 4), in part, imply inadequate model structure (Burnham and Anderson 2003). A low bootstrapderived p-value ( $<0.15$ ) combined with a moderate estimate of c-hat ( $>1$ and $<4$ ) supports overdispersion (and not inadequate model structure). Over-dispersion is corrected with c-hat adjustment (as described above).

## Estimates of survival

The tagging committee calculates maximum likelihood estimates of the multinomial parameters of survival and recovery based on an observed matrix of recaptures (using Program MARK). Candidate models are fit to the tag recovery data and arranged in order of fit by the second-order adjustment to Akaike's information criterion (AICc) (Akaike, 1973; Burnham and Anderson, 1992). Annual survival rates are estimated for two size groups (fish $>=18$ inches TL and fish $>=$ 28 inches TL). Annual survival is calculated as a weighted average across all models, where weight is a function of model fit (Buckland et al. 1997). Model averaging eliminates the need to select the single "best" model, allowing the uncertainty of model selection to be incorporated into the variance of parameter estimates (Burnham and Anderson 2003). Survival is inestimable for the terminal year in the fully time saturated $\{\mathrm{S}(\mathrm{t}) \mathrm{r}(\mathrm{t})\}$ model, so the time saturated model is excluded from the model averaged survival estimate for the terminal year. A weighted average of unconditional variances (conditional on the set of models) is estimated for the model-averaged estimates of survival (Buckland et al. 1997).

## Bias-adjusted estimates of survival

Because we model dead recoveries, survival estimates are adjusted by annual estimates of liverelease bias (Smith et al. 2000),

$$
\text { bias }=-\left[\frac{\theta \cdot P_{L} \cdot \frac{f}{\lambda}}{\left(1-\left(1-\theta \cdot P_{L}\right) \frac{f}{\lambda}\right.}\right] \text {, }
$$

where $\theta=0.92$ (based on an $8 \%$ hook-and-release mortality rate, Diodati and Richards 1996), $P_{L}=$ annual proportion of tagged striped bass released alive, $f=$ annual recovery rate estimated with a Brownie recovery model (Brownie et al. 1985), and $\lambda=$ reporting rate. Annual and geographic-based reporting rates are desirable, but unavailable; consequently we use a constant reporting rate of 0.43 based on a high-reward tag study of the recreational fishery in Delaware Bay (Kahn and Shirey 2000). Gear-specific tagging mortality is not included in bias adjustment because estimates are unavailable for most gears types, such as trawls, pound nets, gill nets, and electrofishing. Estimates of tag-induced mortality are low (0\%, Goshorn et al. 1998; 1.3\% Rugolo and Lange 1993) and excluded from bias adjustments. Additionally, we do not correct for tag loss given low estimates of 0\% (Goshorn et al.1998), $2 \%$ (Dunning et al. 1987), and 2.6\% (Sprankle et al. 1996).

## Estimates of F based on constant M

For each tagging program, instantaneous fishing mortality (F) is estimated by converting the adjusted survival $(S)$ to total mortality $(Z)$ and subtracting a constant value $(M=0.15)$ for natural mortality, where $\mathrm{F}=-\mathrm{LN}(\mathrm{S})-0.15$. Using this technique, natural mortality is held fixed, and any change in total mortality ( Z ) results in an equal change in fishing mortality (F). Uncertainty in estimates of F ( $95 \%$ confidence intervals) are calculated from model-averaged unconditional variances of the adjusted survival estimates. We estimate an average F for coastal programs, and
a weighted-average of F for producer area programs. Weights for producer area averages (based on the estimated proportion of fish contributed to the coast-wide stock, G. Shepherd, pers. comm. and D. Kahn, pers. comm.) are as follows: Hudson (0.13); Delaware (0.09); and Chesapeake Bay (0.78), with MD (0.67) and VA (0.33).

## Estimates of $F$ based on exploitation rate and the catch equation

Ricker (1975, p. 11) presents a formulation of Baranov's catch equation which he recommends for Type 2 fisheries, in which fishing and natural mortality occur concurrently. This is the case for striped bass, where the fishery operates over much of the year. The equation is set up to solve for the exploitation rate, $\mu$. Pollock et al. (1991) solve the same formulation for F as follows:

$$
\mathrm{F}=\mu / \mathrm{A} * \mathrm{Z}
$$

where $A=(1-S)$, the annual total mortality rate. We obtain $Z$ from the bias-corrected survival rates developed from the MARK tag-recovery models described above. Instead of assuming that $M$ is constant and subtracting it from $Z$, however, we rely on the catch equation, which shows that F is a function of both the exploitation rate and Z . Essentially, this formulation is a ratio equation, showing that the ratio of $\mu$ to A equals the ratio of $F$ to $Z$. We have estimates of the exploitation rate (see below), Z and A , with the latter two simple functions of the survival rate estimates obtained via the tag-recovery models. Once F is estimated, we can estimate M by subtracting F from Z. This is the approach used by Crecco (2003).

## Encounter and exploitation rates

In addition to estimates of S and F , we estimated annual catch rates and annual exploitation rates for two length groups ( $>=18$ inches and $>=28$ inches) with tag recoveries of striped bass released by eight agencies (1987-2004) of the Cooperative Striped Bass Tagging Program. Each time series of annual catch rates and annual exploitation rates reflects trends in total catch rate (including releases) and harvest rate, respectively. Estimates of annual catch rates and annual exploitation rates are independent among years; fish at large after the first recovery-year are not used in the analysis. All of the estimates are calculated using a tag reporting rate estimate of 0.43 from a 2000 study conducted on the Delaware River stock, but employing tag returns from the whole Atlantic coast. This estimate is identical to one developed independently and presented in Smith et al. (2000). The reporting rate is the proportion of tagged, recaptured fish whose tag is actually reported to the U. S. F and W. S. Thus we assume that the same tag reporting rate was operative along the whole coast. Annual catch rates and annual exploitation rates are adjusted $\mathrm{R} / \mathrm{M}$ ratios as described below, where R is the number of tags reported as recaptured over the year from the number tagged at the beginning of the year (the recovery rate) and M is the number of fish tagged or marked at the beginning of the year (reporting rate $=0.43$, hooking mortality rate $=0.08, \mathrm{R}_{\mathrm{k}}=$ killed recaptures, $\mathrm{R}_{\mathrm{L}}=$ recaptures released alive):
(1) Annual catch rate $=(R / 0.43) / M$
(2) Annual exploitation rate $=\left(\left(\mathrm{R}_{\mathrm{k}}+\mathrm{R}_{\mathrm{L}} * 0.08\right) / 0.43\right) / \mathrm{M}$

## Stock size estimation

Using the form of Baranov's catch equation, catch $=\mathrm{F}$ * (average stock size), we were able to estimate stock size since we have estimates of total kill and estimates of F. Note that the total kill includes discards, which are generally of the same magnitude as the total landings in number. These estimates were developed for 18 inch plus fish, which in practice is usually fish 18 " or above, corresponding roughly to 3 year old and older striped bass.

Two separate time series of stock sizes were developed. The first was based on the F estimates that assumed constant M for 18 inch plus fish, while the second was based on the estimates generated via the catch equation. Since the F estimates are based on total survival for the constant M estimates, and in the case of the catch equation estimates, exploitation rate that includes discard mortality from released fish that were recaptured, the total kill is the correct variable to employ here.

## Tagging Assessment Results

## Exploitation Rate and Total Catch Rate

The exploitation rate estimates for 28 inch fish are presented by program and as an unweighted mean (Table 20a). For 2004, the two Chesapeake Bay programs, Virginia and Maryland, had the lowest estimates of 0.08 . The highest estimate was from the Delaware River stock, but it was only 0.22 . For the whole time series, coastwide average exploitation rates peaked from 1997 and 1998 at 0.24 and have declined substantially since then. The coastwide, unweighted average for 2004 was 0.13 .

The total catch rates on the coast averaged 0.19 for 2004 (Table 20b). This is continues a declining trend since a peak of 0.34 in 1991. The catch rate estimate for 1997 was 0.31 , over $50 \%$ of the 2004 estimate. The difference between the total catch rate and the exploitation rate suggests that the live release rate was 0.06 . This estimate could be biased low because anglers may be less likely to notice tags on fish they have released. They could also be less likely to recover tags they do notice, since they are releasing the fish. This value of 0.06 is the estimate of release rate since 1997. Prior to 1997, the release rate estimate was substantially higher, as high as 0.21 in 1991.

For 18 inch plus fish, exploitation rate was lower than for 28 inch fish, and declined to 0.09 in 2004 (Table 20c). Catch rate for 18 inch plus fish was also slightly lower than for 28 inch fish, and for $2004=0.17$ (Table 20d). These two values for 2004 were again part of a continuing decline.

Fish >= 28 inches: Estimates of $F$ assuming constant $M$ and stock size of fish aged 7+
Uncorrected survival and F estimates, together with the bias-corrected estimates and confidence intervals for bias-corrected fishing mortality are presented by program in Table 21. The models receiving the higher weights in the final estimate are shown by program in Table 22.

Summaries of the F estimates assuming constant M are in Table 23. The 2004 estimates of F for the four mixed-stock coastal programs (Massachusetts, New York Ocean Haul, New Jersey, and

North Carolina winter trawl) were $0.10,0.02,0.72$, and 0.26 , respectively, with an unweightedmean F of 0.27 (Table 23). The New Jersey Delaware Bay estimates are very erratic among years and the 2004 value of 0.72 has a large influence on the coastal average and the coastwide average. This is the highest F estimate in the time series for the coastal programs. The 2004 estimates for producer area programs Hudson River, Delaware River, and Chesapeake Bay (HUDSON, DE/PA, MDDNR, VARAP) were $0.27,0.32,0.34$, and 0.25 , respectively, with a weighted mean fishing mortality (F) of 0.31, again the highest in the time series. The Delaware River and Maryland Chesapeake estimates were relatively high, with the Hudson and Virginia estimates at a lower level. The average of the coastal and producer programs is the coastwide 2004 estimate for the fully-recruited fish, assuming constant $\mathrm{M}, \mathrm{F}=0.29$. While this estimate is the highest in the time series, it is still slightly below the target $\mathrm{F}=0.30$. Variation in these F rates as additional data has been added is portrayed in Figure 23.

Stock size estimates of 7+ fish developed using this series of F estimates increased to 10.5 million in 2002, then declined slightly to 8.2 million in 2004 (Table 23).

## Fish $>=28$ inches: Estimates of $F$ from the catch equation

Estimates of fully-recruited F for 2004 from the catch equation average only about half the level of the constant M estimates. The 2004 estimates of F for the four mixed-stock coastal programs (Massachusetts, New York Ocean Haul, New Jersey, and North Carolina winter trawl) were $0.10,0.10,0.23$, and 0.15 , respectively, with an unweighted-mean F of 0.15 (Tables 24, 25). The New Jersey estimate was the highest for 2004, as in the constant M estimates, because the survival estimate was low (Table 21). The 2004 estimates for producer area programs Hudson River, Delaware River, and Chesapeake Bay were $0.22,0.26,0.11$ and 0.09 , respectively, with a weighted mean fishing mortality ( F ) of 0.13 (Table 25). The average of the coastal and producer programs is the coastwide 2004 estimate for the fully-recruited fish, assuming constant M , and equals 0.14 (Table 25). The estimates of total abundance obtained with the catch equation F estimates are higher than those obtained with the constant M estimates, because the F estimates are lower, so if the same kill occurs with a lower F, it implies the total stock is larger. These estimates peak in 2004 at 17 million age 7+ fish (Table 25).

## 18 inch plus fish: Estimates of F assuming constant M and stock size estimates from 1990-2004

Estimates of uncorrected survival and fishing mortality by program, assuming constant M , with bias-corrected estimates of these parameters are in Table 26. The F estimates produced under this method were almost as high as those for fully recruited fish for the producer areas, while the coastal program estimates were actually slightly higher than those for the 28 inch coastal F estimates. These estimates were much higher than F estimates produced for 18 inch plus fish using the catch equation method.

Table 27 presents the weights used in averaging the models into the final survival estimates presented in Table 26. F estimates are summarized and stock size estimates are presented in Table 28. The 2004 estimates for producer area programs of Hudson River, Delaware River, Maryland Chesapeake Bay, and Virginia Rappahannock River are 0.25, 0.29, 0.36 and 0.06, respectively, with the average $\mathrm{F}=0.26$. Among producer areas, the Maryland estimates were the highest, at 0.36 for 2004 . The coastal program F estimates are erratic, except for the

Massachusetts results, which are very low and stable. Results of coastal programs in terms of F estimates for 2004 for Massachusetts, New York Long Island, New Jersey Delaware Bay, and North Carolina winter trawl are $0.10,0.30,0.68$ and 0.18 , with an average of 0.31 . The coastwide averages, including both coastal and producer areas, have the highest F estimate in 2000 at $\mathrm{F}=0.40$, then decline to $\mathrm{F}=0.29$ for 2004 , identical to that for the 28 inch fish.

Estimated stock sizes using these F estimates are somewhat erratic, with the peak year in 1995 at 48 million fish (Table 28), due to the very low estimate of $F=0.06$ in 1995. Since 1997, the estimates were more stable and lower, ranging between 12.5 and 15.8 million. The 2004 estimate is the largest of this recent period at 18 million.

18 inch plus fish: estimates of $F$ using the catch equation and stock size estimates from 19902004

Using the catch equation method, the estimates of fishing mortality were lower for 18 inch plus fish than they were assuming constant natural mortality (Tables 29, 30). The producer area average for 2004 was 0.13 and the coastal average was 0.08 . These results are lower than those made with an assumption of constant $M$ because they do not assume that $M=0.15$ and they are dependent on the exploitation rate estimates, which are generally relatively low for 18 inch plus fish. With this method, the coastwide F estimate for 18 inch plus fish in 2004 was only $\mathrm{F}=0.11$. The peak F estimates in this time series occurred in 1997-1998 at $\mathrm{F}=0.16$. F has declined in recent years Table 30 .

The stock size estimates computed from this F series are more reasonable than the previous set of estimates in that they exhibit more stability (Table 30). A moderately consistent stock growth is apparent, with some declines, until about 2000, when stock growth becomes rapid. The estimate for 2004 is the highest in the time series at 48.5 million fish.

## Length frequency, age, and geographic distribution of recaptures

Total length frequencies of fish tagged in 2004 and age distributions of fish recaptured in 2004 were tabulated by program (Tables 31 and 32). Total length frequencies represent the length of fish at the time of tagging. Age distributions are based on a subsample of the total number of tagged fish, because not 18 inch plus fish are aged. Ages (from scales) estimated at the time of tagging are adjusted to the recovery date. For each tagging program, geographic distributions of all recaptures during 2004 (from fish tagged and released during the full time series) were depicted by state and month (Table 33).

## Sources of uncertainty in the tag-based estimates of fishing mortality rate

Confidence intervals have not yet been developed for the estimates based on the catch equation, but will be implemented for next years report.

Violations of the basic assumptions have been investigated in detail for the Virginia tag data set and only very minor violations of the assumption of complete mixing was detected, which did not affect the results of the analysis (J. Hoenig, personal communication). The major concern is that the tagged fish be representative of the stocks.

The estimate of reporting rate employed needs to be re-estimated. This is the rate at which tags are reported in, once tagged fish are recaptured. The current estimate of 0.43 was based on a study in 1999 conducted on the Delaware River spawning stock, employing tag returns from the Atlantic Coast (Kahn and Shirey 2000). If the estimate is too high, then the exploitation rate and the F estimate would be underestimated.. If the rate is too low, then F would be overestimated. A research grant proposal is currently in submission to conduct a high reward tagging study to develop a more current estimate of the reporting rate.

The assumption of constant natural mortality has been contradicted for the Chesapeake Bay stock by two alternative analyses (Crecco 2003, Hoenig, personal communication), both of which found that natural mortality had increased for the resident stock. The catch equation method introduced in this report was used to avoid the assumption of a constant value of M.

Finally, the estimates of F vary somewhat from year-to-year as additional tag returns are added in subsequent years. This variation primarily occurs during the most recent years F estimates. Some of the coastal programs results for 18 inch plus fish are fairly erratic and seem to lack some credibility, but the Tagging Subcommittee has not been able to determine a cause or violation of assumptions as the source of the erratic estimates.

## IV. Discussion

Two major modifications were made this year to the input data for the 2005 VPA which resulted in considerable modifications to the estimates of fishing mortality, spawning stock biomass and population estimates compared to last years VPA run. For the first time, estimates of recreational harvest in NC and VA during Wave 1 (January and February) were included the catch at age. In addition, the Technical Committee modified the suite of tuning indices used in the VPA this year following a comprehensive review of the various indices over the past two years. With these changes plus the inclusion of 2004 data, the 2003 estimate of F declined from 0.62 in last years VPA to 0.29 in this years run. In addition, SSB and total abundance estimates were higher in this year's VPA run.

Another major change to this year's assessment is the utilization of Baranov's catch equation with the tagging data to develop estimates of $F$. By using the $Z$ values from the Brownie models and $\mu$ from $\mathrm{R} / \mathrm{M}, \mathrm{F}$ estimates could be developed without the assumption of a constant M . Additionally, abundance estimates could be calculated using the tagging data for comparison with the VPA abundance estimates.

Coastwide fishing mortality estimates of F for 2004 from tagging estimates and the VPA were all bellow the Amendment 6 threshold value of $F=0.41$. Therefore, it was the consensus of the Technical Committee that overfishing was not occurring on the coastal migratory population of striped bass. However, there were differing opinions within the Technical Committee concerning where the 2004 fish mortality rate was in relation to the Amendment 6 target of $\mathbf{0 . 3 0}$. The differing opinions were the result of the wide distribution in the 2004 coastwide estimates of F from the tagging program and the VPA. The tagging program
included a coastwide average of $\mathrm{F}=0.14$ for fish over 28 inches using the new catch equation method and 0.29 using the traditional method that used a constant value of M. The 2004 estimate of F from the VPA was 0.40 . However, retrospective analysis showed a positive bias with terminal year F estimates from the VPA suggesting that the 2004 value may prove to be lower in subsequent years.

There is currently not a consensus on the Technical Committee as to the reason for the discrepancies between the estimates of $F$. The most prominent reasons discussed are the evidence of an increase in natural mortality (M) in Chesapeake Bay and the potential for a decrease in the tag reporting rate in recent years. An increase in natural mortality above the assumed value of $M=0.15$ would result in overestimates of F's from the VPA and the tag based method using a constant value of M . A significant decrease in the tag reporting rate below the current estimate of 0.43 could result in an under estimate of $F$ with the two tag based estimates.

Comparison of the time series of F estimates from the VPA and the two tagging program methods showed similarities and differences depending on the method used, time frame looked at and size range of fish included in the analysis. Estimates of F on fully recruited fish using assumed age $7+$ fish (fish $\geq 28$ inches) from the tagging program and age $7-11$ weighted by N from the VPA were very similar for all three methods from 1990 to 2001. However, they began to diverge after that with the VPA and constant M tagging method showing increases in F while the catch equation tagging method showed a slight decrease (Figure 24). Estimates of F on age 3 and older fish were very similar when comparing the VPA and catch equation tagging estimates throughout the time series while estimates using the constant $M$ tagging method were more variable as well as consistently higher in value subsequent to 1997 (Figure 25).

Comparison of the time series of population estimates from the three methods presented in this assessment also showed similarities and differences depending on the method and age range of fish. For fish age 3 and older, the VPA and catch equation tagging method showed a general trend of increasing abundance through the 1990's followed by a short term decline and a subsequent increase in abundance in 2003 and 2004 (Figure 26). Using the constant M tagging method, population estimates for age 3+ fish were erratic through the early to mid-1990's followed by a leveling off since 1997 at levels considerably below the estimates from the catch equation tagging method and the VPA. For the age 7+ segment of the striped bass population, abundance estimates from all three methods were similar from 1990 to 2001. Since then, the trend for VPA and constant M tag method estimates have been relatively flat while estimates using the catch equation method with the tagging data have increase substantially (Figure 27).

SSB estimates are available from the VPA and the 2004 estimate of 24.9 thousand metric tons is above both the threshold and target values from Amendment 6 . Therefore, it was the consensus of the Technical Committee that the population is not overfished.

Finally, VPA estimates of age 12+ and 13+ striped bass show an increasing trend in abundance since 2003 (Figure 28). It was the consensus of the Technical Committee that the abundance of older striped bass has increased since the adoption of Amendment 6 in 2003.

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## V. Figures



Figure 1. Proportions of 2004 striped bass mortalities by fishery component.


Figure 2. Commercial harvest of striped bass at age in Chesapeake Bay and in non-Bay states in 2004.


Figure 3. Commercial removals of Atlantic striped bass, 1982-2004.


Figure 4. Total commercial removals of Atlantic striped bass (landings and discards), 1982-2004.


Figure 5. Total commercial removals at age in 2004.


Figure 6. Recreational removals (landings and discard) of Atlantic striped bass at age in 2004.


Figure 7. Length composition of total recreational and commercial striped bass landings, 2004.


Figure 8. Total Recreational removals at age of Atlantic striped bass in 2004.


Figure 9. Total recreational removals of Atlantic striped bass (landings and discards), 1982-2004.


Figure 10. Recreational removals of Atlantic striped bass, 1982-2004


Figure 11. Recreational and commercial removals (landings and discard) in number of Atlantic striped bass, 1982-2004.


Figure 12. Recreational and commercial removals (landings and discard) at age in number for 2003 and 2004.


Figure 13. Fishery-dependent striped bass indices, combined ages.








Figure 14. Fishery-independent surveys of striped bass abundance, combined ages.



Figure 15. Young-of-the-year and yearling indices, 1982-2004.


Figure 16. Striped bass fishing mortality estimates from ADAPT model.


Figure 17. 2004 Fishing mortality at age estimated from ADAPT VPA.


Figure 18. Striped bass population abundance estimates from 2004 ADAPT model


Figure 19. Striped bass female spawning stock biomass (000s mt) and Jan. 1 total biomass (000s mt ) from 2004 assessment.

## Fishing Mortality



Figure 20. Retrospective analysis of fishing mortality and abundance from the 2004 striped bass VPA.


Figure 21. Retrospective analysis of recruitment and female SSB from the 2004 striped bass VPA.

## Fishing Mortality



8+ Abundance


Figure 22. Differences in assessment results resulting from changes in tuning indices used in VPA. Old indices as used in 2004 assessment. New indices as used in 2005 assessment.


Figure 23. Year-to-year variation in F estimates for 28 inch Fish \& Wildlife generated with the constant M method. Labels refer to the year that estimates were reported. Note variation in the vertical axis among charts.


Figure 24. Estimates of F for Atlantic coast striped bass for age $7-11$ fish from the VPA and age $7+$ fish from the tagging program using the catch equation and constant M methods.


Figure 25. Estimates of F for Atlantic coast striped bass for age 3-11 fish from the VPA and age $3+$ fish from the tagging program using the catch equation and constant M methods.


Figure 26. Population estimates of age $3+$ striped bass from the VPA as well as the tagging program using the catch equation and constant M methods.


Figure 27. Population estimates of age $7+$ striped bass from the VPA as well as the tagging program using the catch equation and constant M methods.


Figure 28. Estimates of age 12+ and 13+ abundance of Atlantic coast striped bass from the VPA.

## VI. Tables

Table 1a. Data sources by state for estimating striped bass age structure of recreational discards in 2004.

| State | Number | Length Frequency | Age Length Key |
| :--- | :--- | :--- | :--- |
| ME | MRFSS B2 * 0.08 | ME Volunteer angler log books | MA key from volunteer angler, tagging, and commercial monitoring programs |
| NH | MRFSS B2 * 0.08 | NH Volunteer angler log books | MA key from volunteer angler, tagging, and commercial monitoring programs |
| MA | MRFSS B2 * 0.08 | MA Volunteer angler log data | MA key from volunteer angler, tagging, and commercial monitoring programs |
| RI | MRFSS B2 * 0.08 | ALS Tag length data | NY Ocean haul seine and MA hook and line monitoring data |

Table 1b. Total Atlantic Coast harvest of striped bass in metric tons and numbers from 1982 through 2004.

|  | Commercial |  | Recreational |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | MT | N | MT | N | MT | N |
| 1982 | 998 | 428,630 | 1,146 | 217,256 | 2,145 | 645,886 |
| 1983 | 649 | 357,541 | 1,220 | 299,444 | 1,869 | 656,985 |
| 1984 | 1,110 | 870,871 | 581 | 114,463 | 1,691 | 985,334 |
| 1985 | 437 | 174,621 | 373 | 133,522 | 810 | 308,143 |
| 1986 | 68 | 17,681 | 502 | 114,623 | 570 | 132,304 |
| 1987 | 63 | 13,552 | 388 | 43,755 | 452 | 57,307 |
| 1988 | 117 | 33,310 | 572 | 86,705 | 689 | 120,015 |
| 1989 | 91 | 7,402 | 333 | 37,562 | 424 | 44,964 |
| 1990 | 314 | 115,636 | 1,012 | 163,242 | 1,326 | 278,878 |
| 1991 | 669 | 153,798 | 1,656 | 262,469 | 2,325 | 416,267 |
| 1992 | 652 | 230,714 | 1,830 | 300,180 | 2,482 | 530,894 |
| 1993 | 796 | 312,860 | 2,569 | 428,719 | 3,364 | 741,579 |
| 1994 | 807 | 307,443 | 3,081 | 565,167 | 3,888 | 872,610 |
| 1995 | 1,559 | 534,914 | 5,648 | $1,089,223$ | 7,207 | $1,624,137$ |
| 1996 | 1,227 | 766,518 | 6,091 | $1,199,253$ | 7,318 | $1,965,771$ |
| 1997 | 2,685 | $1,058,181$ | 7,769 | $1,646,971$ | 10,454 | $2,705,152$ |
| 1998 | 2,943 | $1,223,828$ | 6,098 | $1,468,542$ | 9,041 | $2,692,370$ |
| 1999 | 2,994 | $1,103,783$ | 6,713 | $1,448,800$ | 9,707 | $2,552,583$ |
| 2000 | 3,029 | $1,057,711$ | 8,214 | $2,012,685$ | 11,243 | $3,070,396$ |
| 2001 | 3,124 | 941,733 | 9,016 | $2,085,126$ | 12,140 | $3,026,859$ |
| 2002 | 2,692 | 654,062 | 9,023 | $1,970,495$ | 11,715 | $2,624,557$ |
| 2003 | 3,227 | 868,987 | 10,936 | $2,536,272$ | 14,163 | $3,405,259$ |
| 2004 | 3,297 | 907,328 | 11,874 | $2,381,823$ | 15,171 | $3,289,151$ |

Table 2. Total 2004 striped bass harvest and discard in numbers (A) and percent of total by fishery (B).

| A |  |  |  |  |
| :--- | ---: | :--- | ---: | ---: |
| Fishery <br> component | Harvest | Bycatch | Discards | Total <br> Removals |
| Recreational | $2,381,823$ | $17,167,874$ | $1,373,430$ | $3,755,253$ |
| Commercial | 907,328 | $4,108,753$ | 518,847 | $1,426,175$ |
| Total | $3,289,151$ | $21,276,627$ | $1,800,051$ | $5,181,428$ |

B

| Fishery <br> component | Harvest | Discards | Total <br> Removals |
| :--- | :---: | :---: | :---: |
| Recreational | 46.0 | 26.5 | 72.5 |
| Commercial | 17.5 | 10.0 | 27.5 |
| Total | 63.5 | 36.5 | 100.0 |

Table 3. Commercial landings in number of Atlantic striped bass by state, 1982-2004.


## For Technical Committee Review Only

Table 4. Atlantic Coast striped bass commercial harvest in numbers at age by state in 2004.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  | total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ |  |
| MA |  |  |  |  |  | 242 | 2,368 | 12,414 | 13,139 | 12,824 | 8,224 | 5,131 | 6,291 | 60,633 |
| RI |  | 5 | 104 | 678 | 1,618 | 1,887 | 3,337 | 2,690 | 2,147 | 1,685 | 683 | 534 | 499 | 15,867 |
| NY |  |  | 298 | 5,802 | 8,182 | 11,306 | 10,116 | 19,340 | 8,629 | 3,868 | 1,488 | 1,190 | 149 | 70,367 |
| DE |  |  | 1 | 365 | 3,292 | 7,894 | 2,527 | 6,909 | 2,929 | 3,099 | 1,024 | 290 | 75 | 28,406 |
| MD |  | 10 | 56,997 | 143,502 | 94,104 | 62,309 | 59,185 | 26,758 | 10,122 | 3,687 | 3,433 | 329 | 628 | 461,064 |
| VA |  |  | 2,150 | 6,158 | 7,442 | 11,992 | 9,323 | 20,255 | 25,205 | 20,959 | 20,728 | 12,751 | 11,033 | 147,998 |
| PRFC |  |  | 23,615 | 29,486 | 26,923 | 7,259 | 1,562 | 1,930 | 827 | 368 |  |  |  | 91,980 |
| NC |  |  |  |  |  |  | 790 | 3,971 | 7,664 | 9,119 | 5,699 | 2,319 | 1,452 | 31,014 |
| Total |  | 15 | 83,165 | 186,002 | 141,562 | 102,891 | 89,208 | 94,267 | 70,661 | 55,610 | 41,279 | 22,544 | 20,126 | 907,329 |

Table 5. Recovery of tagged striped bass by commercial gear in 2004 and assumed gear specific release mortalities.

|  | Commercial Gear |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Anchor Gill Net | Drift <br> Gill Net | Hook\&Line | Other | Pound Net | Seine | Trawl |  |
| Number |  |  |  |  |  |  |  |  |
| Chesapeake Bay | 46 | 42 | 28 | 1 | 334 |  |  | 451 |
| Coast | 11 | 1 | 65 | 1 | 8 |  | 11 | 97 |
| Delaware Bay | 3 | 4 |  |  |  |  |  | 7 |
| Percent |  |  |  |  |  |  |  |  |
| Chesapeake Bay | 0.10 | 0.09 | 0.06 | 0.00 | 0.74 |  |  |  |
| Coast | 0.11 | 0.01 | 0.67 | 0.01 | 0.08 |  | 0.12 |  |
| Delaware Bay | 0.43 | 0.57 |  |  |  |  |  |  |
| Release Mortality | 0.43 | 0.08 | 0.08 | 0.20 | 0.05 | 0.15 | 0.35 |  |

Table 6. Ratios of commercial and recreational landings and tag recaptures from released and kept fish in Chesapeake Bay and the Atlantic Coast.

|  |  |  | Ches |  |  | Coast |  |  | DE Bay |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Com | Rec | Ratio | Com | Rec | Ratio | Com | Rec | Ratio |
| 2002 | Landings | 504,146 | 603,250 | 0.84 | 116,847 | 1,193,019 | 0.10 |  |  |  |
|  | Landed tags | 181 | 609 | 0.30 | 48 | 636 | 0.08 |  |  |  |
|  | Discard tags | 41 | 316 | 0.13 | 25 | 600 | 0.04 |  |  |  |
| 2003 | Landings | 662,518 | 886,330 | 0.75 | 203,171 | 1,519,377 | 0.13 |  |  |  |
|  | Landed tags | 407 | 523 | 0.78 | 34 | 774 | 0.04 |  |  |  |
|  | Discard tags | 79 | 279 | 0.28 | 13 | 649 | 0.02 |  |  |  |
| 2004 | Landings | 677,662 | 730,222 | 0.93 | 228,003 | 1,443,282 | 0.16 | 28,406 | 179,657 | 0.16 |
|  | Landed tags | 348 | 497 | 0.70 | 74 | 731 | 0.10 | 2 | 59 | 0.03 |
|  | Discard tags | 104 | 221 | 0.47 | 23 | 600 | 0.04 | 5 | 42 | 0.12 |
| Three year mean of landings ratios (02-04) |  |  |  | 0.84 |  |  | 0.13 |  |  |  |
| Three year mean of landings tags (02-04) |  |  |  | 0.59 |  |  | 0.07 |  |  |  |
| Correction factor |  |  |  | 1.41 |  |  | 1.77 |  |  | 1.59 |

Table 7a. Data sources for estimating striped bass age structure of commercial discards in 2004.

| Gear | Data source | Data type | conversion to <br> age |
| :--- | :--- | :--- | :--- |
| Coastal gill net | NEFC Observer Program for 2004 <br> (N=52 fish) | len-freq | state age-len <br> key |
| Coastal Hook and Line | H\&L discards <br> MA 2004 compliance report | age structure |  |
| Coastal Pound Net | Trap net discards <br> RI 2004 compliance report | age structure |  |
| Coastal Otter Trawl | NEFC Observer Program for 2004 <br> (N=666 fish) | len-freq | state age-len <br> key |
| Ches Bay Anchor Gill <br> Net | FI sampling, James \& Rappahannock <br> R. <br> VA 2004 compliance report | age structure |  |
| Ches Bay Drift Gill Net | Drift gill net havest <br> MD 2004 compliance report | age structure |  |
| Ches Bay Hook and Line | H\&L and pound net harvest <br> MD 2004 compliance report | age structure |  |
| Ches Bay Pound Net | FI sampling, Rappahannock R. <br> VA 2004 compliance report | age structure |  |
| Del Bay Gill Net | NJ Del. Bay tagging program <br> USFWS coastwide tagging database | len-freq | state age-len <br> key |

Table 7b. Atlantic Coast striped bass commercial discards in numbers at age, 2004.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ |  |
| Coast | 28 | 1,211 | 3,966 | 7,788 | 11,164 | 17,768 | 18,957 | 25,146 | 13,615 | 4,664 | 4,937 | 1,657 | 748 | 111,650 |
| Ches. | 2,956 | 51,833 | 75,072 | 66,281 | 40,487 | 24,198 | 30,617 | 15,307 | 9,838 | 5,195 | 5,527 | 719 | 328 | 328,358 |
| Bay Del. Bay Hudson |  |  | 341 | $3,709$ | 7,866 | $\begin{array}{r} 19,303 \\ 144 \end{array}$ | $\begin{array}{r} 18,275 \\ 62 \end{array}$ | $\begin{array}{r} 17,063 \\ \hline 37 \end{array}$ | $\begin{aligned} & 8,362 \\ & 27 \end{aligned}$ | $\begin{array}{r} 2,378 \\ 5 \end{array}$ | $\begin{array}{r} 645 \\ 3 \end{array}$ | $\begin{array}{r} 193 \\ 3 \end{array}$ | 123 1 | $\begin{array}{r} 78,258 \\ 581 \end{array}$ |
| Total | 2,984 | 53,044 | 79,379 | 77,906 | 59,688 | 61,413 | 67,911 | 57,553 | 31,842 | 12,242 | 11,112 | 2,572 | 1,200 | 518,847 |

Table 8. Total Atlantic Coast striped bass recreational harvest in numbers by state, 1982-2004.

| Year | State |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | ME | NH | MA | RI | CT | NY | NJ | DE | MD | VA | NC | Total |
| 1982 | 929 |  | 83,933 | 1,757 | 50,081 | 21,278 | 58,294 |  | 984 |  |  | 217,256 |
| 1983 | 7,212 | 4,576 | 39,316 | 1,990 | 42,826 | 43,731 | 127,912 | 135 | 31,746 |  |  | 299,444 |
| 1984 |  |  | 3,481 | 1,230 | 5,678 | 57,089 | 13,625 | 16,571 | 16,789 |  |  | 114,463 |
| 1985 | 11,862 |  | 66,019 | 670 | 15,350 | 23,107 | 13,145 |  | 2,965 | 404 |  | 133,522 |
| 1986 |  |  | 29,434 | 3,291 | 1,760 | 27,477 | 36,999 |  | 14,077 | 1,585 |  | 114,623 |
| 1987 |  | 90 | 10,807 | 2,399 | 522 | 14,191 | 9,279 |  | 4,025 | 2,442 |  | 43,755 |
| 1988 |  | 647 | 21,050 | 5,226 | 2,672 | 20,230 | 12,141 |  | 133 | 24,259 | 347 | 86,705 |
| 1989 | 738 |  | 13,044 | 4,303 | 5,777 | 12,388 | 1,312 |  |  |  |  | 37,562 |
| 1990 | 2,912 | 617 | 20,515 | 4,677 | 6,082 | 24,799 | 44,878 | 2,009 | 736 | 56,017 |  | 163,242 |
| 1991 | 3,265 | 274 | 20,799 | 17,193 | 4,907 | 54,502 | 38,300 | 2,741 | 77,873 | 42,224 | 391 | 262,469 |
| 1992 | 6,357 | 2,213 | 57,084 | 14,945 | 9,154 | 45,162 | 41,426 | 2,400 | 99,354 | 21,118 | 967 | 300,180 |
| 1993 | 612 | 1,540 | 58,511 | 17,826 | 19,253 | 78,560 | 64,935 | 4,055 | 104,682 | 78,481 | 264 | 428,719 |
| 1994 | 3,771 | 3,023 | 74,538 | 5,915 | 16,929 | 87,225 | 34,877 | 4,140 | 199,378 | 127,945 | 7,426 | 565,167 |
| 1995 | 2,189 | 3,902 | 73,806 | 29,997 | 38,261 | 155,821 | 254,055 | 15,361 | 355,237 | 149,103 | 11,491 | 1,089,223 |
| 1996 | 1,893 | 6,461 | 68,300 | 60,074 | 62,840 | 225,428 | 127,952 | 22,867 | 337,415 | 250,731 | 35,291 | 1,199,252 |
| 1997 | 35,259 | 13,546 | 199,373 | 62,162 | 64,639 | 236,287 | 67,800 | 19,706 | 334,068 | 518,483 | 95,648 | 1,646,971 |
| 1998 | 38,094 | 5,929 | 207,952 | 44,890 | 64,215 | 181,031 | 88,973 | 18,758 | 391,824 | 383,786 | 43,089 | 1,468,541 |
| 1999 | 21,102 | 4,641 | 126,755 | 56,320 | 55,805 | 197,672 | 237,010 | 8,772 | 263,191 | 411,873 | 65,659 | 1,448,800 |
| 2000 | 62,186 | 4,262 | 181,295 | 95,496 | 53,191 | 259,085 | 402,302 | 39,543 | 506,462 | 389,126 | 19,737 | 2,012,685 |
| 2001 | 59,947 | 15,291 | 288,032 | 80,125 | 54,165 | 189,710 | 560,208 | 41,195 | 382,557 | 355,020 | 58,876 | 2,085,126 |
| 2002 | 71,907 | 12,857 | 308,749 | 78,190 | 51,060 | 200,547 | 416,455 | 29,149 | 282,429 | 411,248 | 107,904 | 1,970,495 |
| 2003 | 56,871 | 24,878 | 407,100 | 115,471 | 94,361 | 313,761 | 391,842 | 29,522 | 525,191 | 455,812 | 121,463 | 2,536,272 |
| 2004 | 36,091 | 10,057 | 406,590 | 73,964 | 72,368 | 242,840 | 421,009 | 23,884 | 313,914 | 467,389 | 313,717 | 2,381,823 |

Table 9. Total Atlantic Coast striped bass recreational harvest in numbers at age by state in 2004.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ |  |
| ME |  |  | 7,070 | 15,219 | 8,177 | 4,570 | 426 | 374 | 10 | 34 | 58 | 42 | 111 | 36,091 |
| NH |  |  | 24 | 81 | 1,028 | 2,475 | 2,575 | 2,070 | 577 | 434 | 376 | 150 | 267 | 10,057 |
| MA |  |  | 0 | 1,853 | 17,009 | 51,155 | 84,664 | 104,031 | 52,101 | 37,072 | 27,774 | 10,565 | 20,366 | 406,590 |
| RI |  |  | 32 | 410 | 2,308 | 6,148 | 9,072 | 16,606 | 12,320 | 9,121 | 6,484 | 3,451 | 8,013 | 73,964 |
| CT |  |  | 0 | 550 | 4,862 | 9,717 | 11,246 | 17,548 | 13,810 | 6,824 | 5,687 | 1,432 | 693 | 72,368 |
| NY |  |  | 0 | 3,017 | 14,609 | 35,953 | 32,689 | 54,090 | 38,947 | 17,285 | 27,113 | 16,336 | 2,802 | 242,840 |
| NJ |  |  | 7,435 | 44,277 | 71,076 | 59,623 | 48,676 | 89,110 | 46,245 | 24,100 | 12,234 | 6,670 | 11,564 | 421,009 |
| DE |  |  | 81 | 449 | 1,506 | 3,347 | 4,481 | 7,515 | 2,774 | 2,202 | 625 | 331 | 573 | 23,884 |
| MD | 184 | 504 | 37,187 | 64,959 | 30,130 | 45,539 | 37,143 | 23,820 | 20,816 | 20,323 | 18,580 | 7,464 | 7,266 | 313,914 |
| VA |  |  | 50,531 | 64,144 | 54,412 | 49,684 | 45,462 | 57,838 | 49,270 | 30,262 | 31,735 | 14,323 | 19,729 | 467,389 |
| NC |  |  |  |  |  |  | 4,907 | 28,392 | 63,094 | 92,888 | 64,496 | 30,320 | 29,619 | 313,717 |
| Total | 184 | 504 | 102,358 | 194,959 | 205,118 | 268,209 | 281,342 | 401,393 | 299,965 | 240,546 | 195,160 | 91,084 | 101,002 | 2,381,823 |

Table 10. Total Atlantic Coast striped bass recreational discards in numbers at age by state, 2004.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ |  |
| ME |  |  | 640 | 31,084 | 14,537 | 5,440 | 3,953 | 1,395 | 1,373 | 356 | 216 | 99 | 42 | 70 | 59,207 |
| NH |  |  | 178 | 7,101 | 3,764 | 1,961 | 1,518 | 487 | 405 | 94 | 71 | 46 | 22 | 51 | 15,699 |
| MA |  |  | 3,723 | 172,783 | 99,295 | 55,749 | 50,651 | 30,496 | 31,946 | 11,096 | 7,046 | 3,814 | 1,404 | 2,280 | 470,284 |
| RI |  | 23 | 5,179 | 24,557 | 10,705 | 5,170 | 3,405 | 823 | 361 | 54 |  |  |  |  | 50,278 |
| CT |  | 1,409 | 12,014 | 27,652 | 14,401 | 6,978 | 7,225 | 3,623 | 5,195 | 3,728 | 1,795 | 1,570 | 539 | 1,026 | 87,156 |
| NY |  | 102 | 17,352 | 49,495 | 19,127 | 8,246 | 6,811 | 3,413 | 4,635 | 3,441 | 1,524 | 2,337 | 1,320 | 0 | 117,804 |
| NJ | 15 | 2,220 | 7,202 | 27,683 | 28,580 | 12,690 | 8,628 | 8,179 | 7,352 | 3,543 | 1,727 | 881 | 416 | 638 | 109,755 |
| DE |  |  |  | 1,630 | 3,229 | 1,861 | 1,711 | 832 | 1,083 | 578 | 493 | 251 | 114 | 199 | 11,981 |
| MD |  | 41,446 | 38,828 | 120,428 | 59,545 | 9,200 | 8,785 | 8,026 | 3,773 | 1,636 | 79 |  |  | 18 | 291,764 |
| VA |  | 19,969 | 19,283 | 58,195 | 28,040 | 4,121 | 3,714 | 3,419 | 1,459 | 692 | 22 |  |  | 5 | 138,919 |
| NC |  | 85 | 1,047 | 7,697 | 4,306 | 2,176 | 1,859 | 1,095 | 1,159 | 505 | 280 | 198 | 85 | 92 | 20,584 |
| Total | 15 | 65,254 | 105,446 | 528,305 | 285,531 | 113,594 | 98,259 | 61,789 | 58,742 | 25,724 | 13,253 | 9,197 | 3,942 | 4,380 | 1,373,430 |

Table 11. Total Atlantic Coast striped bass recreational harvest and discards in number at age by state, 2004.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ |  |
| ME |  |  | 640 | 38,153 | 29,757 | 13,617 | 8,523 | 1,821 | 1,747 | 366 | 250 | 157 | 85 | 181 | 95,298 |
| NH |  |  | 178 | 7,124 | 3,845 | 2,988 | 3,992 | 3,063 | 2,475 | 671 | 506 | 422 | 173 | 318 | 25,756 |
| MA |  |  | 3,723 | 172,783 | 101,148 | 72,758 | 101,806 | 115,160 | 135,977 | 63,197 | 44,118 | 31,588 | 11,969 | 22,647 | 876,874 |
| RI |  | 23 | 5,179 | 24,589 | 11,115 | 7,479 | 9,553 | 9,894 | 16,967 | 12,374 | 9,121 | 6,484 | 3,451 | 8,013 | 124,242 |
| CT |  | 1,409 | 12,014 | 27,652 | 14,951 | 11,840 | 16,942 | 14,869 | 22,743 | 17,538 | 8,619 | 7,257 | 1,971 | 1,719 | 159,524 |
| NY |  | 102 | 17,352 | 49,495 | 22,144 | 22,855 | 42,764 | 36,102 | 58,725 | 42,388 | 18,809 | 29,450 | 17,656 | 2,802 | 360,644 |
| NJ | 15 | 2,220 | 7,202 | 35,118 | 72,858 | 83,766 | 68,250 | 56,855 | 96,461 | 49,788 | 25,827 | 13,115 | 7,085 | 12,202 | 530,764 |
| DE |  |  |  | 1,711 | 3,678 | 3,367 | 5,058 | 5,313 | 8,597 | 3,353 | 2,695 | 876 | 445 | 772 | 35,865 |
| MD |  | 41,630 | 39,332 | 157,614 | 124,504 | 39,330 | 54,324 | 45,169 | 27,593 | 22,452 | 20,402 | 18,580 | 7,464 | 7,284 | 605,679 |
| VA |  | 19,969 | 19,283 | 108,726 | 92,184 | 58,533 | 53,397 | 48,882 | 59,297 | 49,962 | 30,285 | 31,735 | 14,323 | 19,734 | 606,308 |
| NC |  | 85 | 1,047 | 7,697 | 4,306 | 2,176 | 1,859 | 6,002 | 29,552 | 63,599 | 93,168 | 64,694 | 30,405 | 29,711 | 334,301 |
| TOTAL | 15 | 65,438 | 105,950 | 630,663 | 480,490 | 318,711 | 366,468 | 343,131 | 460,135 | 325,689 | 253,799 | 204,357 | 95,025 | 105,382 | 3,755,253 |

Table 12. Total Atlantic Coast striped bass catch at age, including recreational and commercial harvest and discards, 1982-2004. Numbers in thousands.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ |  |
| 1982 | 1.8 | 105.6 | 256.7 | 220.8 | 58.4 | 19.2 | 24.2 | 16.8 | 11.7 | 10.6 | 11 | 13.7 | 15.7 | 766.2 |
| 1983 | 3.6 | 110.3 | 178.2 | 193.1 | 150 | 39.3 | 18.7 | 4.1 | 2.9 | 3.7 | 4.6 | 5.6 | 13.6 | 727.7 |
| 1984 | 5.6 | 542.8 | 302.7 | 82.4 | 60.4 | 51.7 | 18.3 | 4.7 | 2.1 | 2.1 | 0.7 | 0.3 | 11.1 | 1084.9 |
| 1985 | 1.3 | 72.5 | 102 | 40.5 | 58.7 | 43.1 | 43.5 | 17.3 | 6.4 | 3.4 | , | 0.8 | 10.3 | 400.8 |
| 1986 | 11.3 | 21 | 63.8 | 132.9 | 49.9 | 32 | 20.4 | 24 | 9.2 | 5.3 | 3.4 | 1.6 | 10.1 | 384.9 |
| 1987 | 1.4 | 10.9 | 37.6 | 51.4 | 67.3 | 25 | 13.2 | 6.5 | 6.4 | 3 | 1.5 | 2 | 12.9 | 239.1 |
| 1988 | 2.6 | 30.9 | 41.8 | 63.2 | 107.1 | 97.9 | 40.6 | 24.4 | 14 | 5.8 | 3.7 | 3.3 | 9.6 | 444.9 |
| 1989 | 0.7 | 36 | 79.7 | 68.2 | 104.9 | 95.4 | 45.7 | 21 | 10.4 | 3.8 | 3.2 | 2 | 8.9 | 479.9 |
| 1990 | 2.1 | 46.2 | 124.5 | 187.8 | 173.2 | 165.2 | 104.1 | 67.9 | 20.7 | 7.3 | 5.1 | 3.5 | 13.7 | 921.3 |
| 1991 | 1.8 | 72.8 | 145.3 | 208.7 | 162 | 101.4 | 91.3 | 82.9 | 58.8 | 24.1 | 14.2 | 2.8 | 22.3 | 988.4 |
| 1992 | 2.9 | 45.8 | 199.7 | 189.2 | 177.1 | 109.5 | 62.4 | 67.8 | 58.4 | 44.8 | 9.3 | 4.1 | 15.9 | 986.9 |
| 1993 | 0.3 | 69.6 | 185.3 | 327.3 | 288.5 | 185.4 | 86.6 | 67.3 | 82.6 | 76.2 | 41.1 | 9.3 | 17.5 | 1437 |
| 1994 | 5.7 | 145.4 | 348.8 | 290.6 | 367.8 | 232.4 | 135.4 | 86.7 | 99.9 | 81 | 36 | 22.3 | 14.6 | 1866.6 |
| 1995 | 4.1 | 433.5 | 470.8 | 456.1 | 405.3 | 489.9 | 214.5 | 196.0 | 153.8 | 90.6 | 53.4 | 17.5 | 14.2 | 2999.7 |
| 1996 | 1.0 | 98.8 | 649.4 | 650.1 | 542.9 | 468.7 | 442.2 | 209.6 | 136.8 | 68.9 | 42.5 | 46.3 | 19.0 | 3376.2 |
| 1997 | 3.3 | 291.5 | 602.0 | 971.2 | 685.3 | 655.7 | 458.6 | 415.7 | 223.5 | 140.6 | 70.0 | 34.0 | 28.7 | 4580.2 |
| 1998 | 26.4 | 183.4 | 485.4 | 706.7 | 1125.0 | 510.9 | 280.4 | 265.0 | 215.5 | 113.8 | 95.1 | 45.2 | 65.5 | 4118.3 |
| 1999 | 8.4 | 108.3 | 419.6 | 648.8 | 642.2 | 730.2 | 351.8 | 238.9 | 205.4 | 148.4 | 104.5 | 48.6 | 49.2 | 3704.4 |
| 2000 | 38.0 | 323.9 | 419.9 | 989.2 | 1021.2 | 780.4 | 738.1 | 311.9 | 160.6 | 141.5 | 59.6 | 29.3 | 30.8 | 5044.4 |
| 2001 | 34.7 | 161.9 | 431.5 | 605.4 | 830.6 | 696.7 | 576.8 | 480.4 | 205.8 | 119.6 | 103.0 | 49.6 | 48.0 | 4344.0 |
| 2002 | 24.5 | 207.0 | 226.9 | 254.0 | 450.2 | 651.5 | 668.6 | 497.7 | 341.5 | 260.2 | 109.6 | 86.5 | 111.3 | 3889.5 |
| 2003 | 28.4 | 254.6 | 476.4 | 601.6 | 707.8 | 604.3 | 708.1 | 494.1 | 374.2 | 284.2 | 127.9 | 80.9 | 93.7 | 4836.2 |
| 2004 | 68.5 | 159.1 | 793.9 | 745.3 | 520.6 | 531.1 | 500.5 | 612.2 | 428.3 | 321.7 | 256.8 | 120.2 | 126.8 | 5184.8 |

Table 13. Mean weight at age (kg) of harvested and discarded Atlantic striped bass, 1982-2004.

| Year | Age |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13/13+ |
| 1982 | 0.13 | 0.64 | 1.09 | 1.54 | 2.42 | 3.75 | 4.83 | 5.79 | 6.20 | 8.68 | 10.80 | 11.20 | 14.05 |
| 1983 | 0.20 | 0.55 | 0.94 | 1.37 | 2.37 | 3.29 | 3.77 | 5.36 | 6.01 | 8.10 | 9.57 | 10.39 | 11.11 |
| 1984 | 0.24 | 0.60 | 1.69 | 1.62 | 2.67 | 3.39 | 5.07 | 5.65 | 6.76 | 7.76 | 8.41 | 12.65 | 12.38 |
| 1985 | 0.06 | 0.61 | 1.07 | 1.66 | 2.19 | 3.59 | 4.91 | 5.46 | 6.77 | 7.45 | 9.00 | 10.69 | 13.91 |
| 1986 | 0.14 | 0.57 | 1.27 | 2.40 | 2.44 | 3.12 | 3.95 | 5.05 | 5.44 | 6.09 | 7.75 | 9.16 | 12.78 |
| 1987 | 0.20 | 0.77 | 1.41 | 2.11 | 2.50 | 2.91 | 3.61 | 4.74 | 5.52 | 6.49 | 7.77 | 9.78 | 13.15 |
| 1988 | 0.31 | 0.91 | 1.10 | 1.98 | 3.12 | 4.02 | 4.38 | 4.70 | 5.24 | 5.62 | 8.58 | 10.40 | 13.27 |
| 1989 | 0.16 | 0.83 | 1.22 | 2.23 | 3.06 | 4.53 | 5.37 | 6.23 | 6.04 | 8.68 | 8.94 | 9.74 | 13.36 |
| 1990 | 0.08 | 0.89 | 1.14 | 2.05 | 2.35 | 3.83 | 4.91 | 5.96 | 5.70 | 5.97 | 7.44 | 9.08 | 12.60 |
| 1991 | 0.21 | 0.92 | 1.29 | 2.17 | 2.62 | 3.17 | 4.81 | 5.64 | 6.46 | 6.24 | 9.46 | 8.30 | 14.22 |
| 1992 | 0.10 | 0.69 | 1.31 | 1.93 | 2.81 | 3.67 | 4.90 | 5.79 | 6.96 | 8.15 | 9.77 | 12.44 | 13.97 |
| 1993 | 0.07 | 0.76 | 1.31 | 1.99 | 2.77 | 3.58 | 4.80 | 6.11 | 7.03 | 8.01 | 9.53 | 10.76 | 14.55 |
| 1994 | 0.24 | 1.05 | 1.69 | 2.21 | 2.85 | 3.50 | 4.94 | 6.20 | 6.80 | 7.53 | 9.73 | 10.69 | 12.73 |
| 1995 | 0.28 | 0.70 | 1.35 | 2.18 | 2.77 | 3.65 | 5.38 | 6.16 | 7.27 | 8.86 | 7.57 | 9.73 | 16.66 |
| 1996 | 0.14 | 1.05 | 1.47 | 2.32 | 3.23 | 4.52 | 6.39 | 7.11 | 7.81 | 9.20 | 9.31 | 10.10 | 13.70 |
| 1997 | 0.13 | 0.62 | 1.18 | 2.46 | 2.81 | 3.64 | 4.51 | 5.07 | 6.73 | 9.17 | 9.94 | 10.24 | 14.78 |
| 1998 | 0.39 | 0.77 | 1.20 | 1.62 | 2.25 | 2.95 | 4.69 | 5.66 | 6.82 | 7.03 | 7.76 | 9.87 | 11.87 |
| 1999 | 0.62 | 0.90 | 1.11 | 1.44 | 1.91 | 2.51 | 3.36 | 5.03 | 6.56 | 7.85 | 8.69 | 9.76 | 11.98 |
| 2000 | 0.37 | 0.55 | 1.10 | 1.45 | 1.96 | 2.79 | 3.89 | 5.09 | 7.11 | 7.37 | 9.70 | 10.70 | 13.55 |
| 2001 | 0.16 | 0.38 | 1.12 | 1.75 | 2.21 | 3.25 | 4.12 | 5.02 | 6.36 | 7.79 | 8.65 | 8.29 | 10.87 |
| 2002 | 0.12 | 0.31 | 1.06 | 1.51 | 2.18 | 3.17 | 4.19 | 5.48 | 6.03 | 7.56 | 9.09 | 9.75 | 11.52 |
| 2003 | 0.10 | 0.60 | 1.00 | 1.40 | 2.20 | 3.20 | 4.10 | 5.20 | 6.10 | 7.20 | 8.50 | 9.40 | 11.00 |
| 2004 | 0.23 | 0.33 | 0.84 | 1.40 | 2.43 | 3.11 | 4.14 | 5.17 | 6.07 | 7.12 | 8.18 | 9.03 | 10.71 |

Table 14. Indices of abundance for Atlantic striped bass adjusted to appropriate 1 January measurement time, 1982-2005. Indicesused in the final ADAPT run are highlighted in bold.

| Index | MACOM | MACOM | MACOM | MACOM | MACOM | MACOM | MACOM | MACOM | MACOM | MDSSN | MDSSN |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ | 3 | 4 |
| Date Tuned | mean | mean | mean | mean | mean | mean | mean | mean | mean | 1-Jan | 1-Jan |
| Index Type | number | number | number | number | number | number | number | number | number | number | number |
| Year |  |  |  |  |  |  |  |  |  |  |  |
| 1982 | - | - | - | - | - | - | - | - | - | - | - |
| 1983 | - | - | - | - | - | - | - | - | - | - | - |
| 1984 | - | - | - | - | - | - | - | - | - | - | - |
| 1985 | - | - | - | - | - | - | - | - | - | 303.6 | 31.9 |
| 1986 | - | - | - | - | - | - | - | - | - | 260.0 | 495.8 |
| 1987 | - | - | - | - | - | - | - | - | - | 251.7 | 111.1 |
| 1988 | - | - | - | - | - | - | - | - | - | 73.6 | 70.7 |
| 1989 | - | - | - | - | - | - | - | - | - | 152.5 | 80.4 |
| 1990 | - | - | - | - | - | - | - | - | - | 158.1 | 120.3 |
| 1991 | 0.064 | 0.064 | 0.07 | 0.11 | 0.12 | 0.07 | 0.02 | 0.01 | 0.03 | 191.1 | 62.2 |
| 1992 | 0.039 | 0.05 | 0.07 | 0.14 | 0.17 | 0.11 | 0.02 | 0.01 | 0.03 | 218.7 | 152.6 |
| 1993 | 0.048 | 0.054 | 0.05 | 0.08 | 0.16 | 0.15 | 0.07 | 0.01 | 0.04 | 132.0 | 186.0 |
| 1994 | 0.044 | 0.065 | 0.05 | 0.07 | 0.14 | 0.14 | 0.08 | 0.03 | 0.03 | 103.5 | 97.3 |
| 1995 | 0.023 | 0.046 | 0.08 | 0.14 | 0.14 | 0.13 | 0.08 | 0.05 | 0.03 | 117.2 | 67.3 |
| 1996 | 0.026 | 0.05 | 0.09 | 0.17 | 0.16 | 0.14 | 0.08 | 0.04 | 0.03 | 368.3 | 102.2 |
| 1997 | 0.032 | 0.055 | 0.09 | 0.21 | 0.10 | 0.11 | 0.07 | 0.03 | 0.03 | 46.3 | 134.6 |
| 1998 | 0.06 | 0.068 | 0.09 | 0.15 | 0.14 | 0.08 | 0.07 | 0.03 | 0.03 | 142.8 | 32.7 |
| 1999 | 0.037 | 0.067 | 0.06 | 0.08 | 0.15 | 0.15 | 0.08 | 0.05 | 0.04 | 174.2 | 80.1 |
| 2000 | 0.037 | 0.073 | 0.12 | 0.11 | 0.17 | 0.15 | 0.08 | 0.03 | 0.03 | 50.7 | 107.6 |
| 2001 | 0.095 | 0.085 | 0.15 | 0.13 | 0.09 | 0.05 | 0.05 | 0.02 | 0.02 | 39.1 | 52.3 |
| 2002 | 0.065 | 0.173 | 0.11 | 0.12 | 0.09 | 0.05 | 0.05 | 0.03 | 0.04 | 41.5 | 38.5 |
| 2003 | 0.072 | 0.079 | 0.17 | 0.13 | 0.12 | 0.07 | 0.04 | 0.04 | 0.02 | 110.0 | 47.8 |
| 2004 | 0.073 | 0.118 | 0.15 | 0.20 | 0.10 | 0.07 | 0.04 | 0.02 | 0.04 | 179.1 | 121.7 |

Table 14 cont'd.

| Index | MDSSN | MDSSN | MDSSN | MDSSN | MDSSN | MDSSN | MDSSN | MDSSN | MDSSN | NYOHS | NYOHS |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13+ | 3 | 4 |
| Date Tuned | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan |
| Index Type | number | number | number | number | number | number | number | number | number | number | number |
|  |  |  |  |  |  |  |  |  |  |  |  |
| Year |  |  |  |  |  |  |  |  |  |  |  |
| 1982 | - | - | - | - | - | - | - | - | - | - | - |
| 1983 | - | - | - | - | - | - | - | - | - | - | - |
| 1984 | - | - | - | - | - | - | - | - | - | - | - |
| 1985 | 4.5 | 1.2 | 1.7 | 0 | 0.3 | 0.1 | 0 | 0.2 | 0.7 | - | - |
| 1986 | 4.1 | 5.3 | 2.1 | 3 | 3 | 0 | 0 | 0 | 0.9 | - | - |
| 1987 | 188.8 | 1.8 | 1.6 | 4.2 | 0.2 | 0 | 0 | 0 | 10.8 | - | - |
| 1988 | 57.7 | 77.4 | 1.4 | 0 | 0 | 4.3 | 0 | 0 | 0.4 | 1.13 | 6.93 |
| 1989 | 45.5 | 48.9 | 33.3 | 0.2 | 0.1 | 0 | 0 | 0 | 0 | 6.41 | 7.64 |
| 1990 | 48.3 | 34.3 | 32 | 29.8 | 0.9 | 0.1 | 0.1 | 0.5 | 1 | 1.86 | 2.73 |
| 1991 | 47.1 | 26.7 | 26.1 | 19.2 | 10.7 | 0.4 | 1.5 | 0 | 2.3 | 1.89 | 9.19 |
| 1992 | 58.8 | 70.1 | 43.2 | 29.4 | 13.9 | 7.3 | 3.3 | 0 | 2.4 | 5.23 | 9.26 |
| 1993 | 88.5 | 51.2 | 52.2 | 37.5 | 23 | 7.7 | 3.2 | 0.8 | 3 | 1.49 | 7.84 |
| 1994 | 118 | 59.6 | 34.1 | 43.1 | 17.8 | 8.7 | 3.1 | 1.3 | 0.3 | 3.81 | 9.43 |
| 1995 | 60.9 | 51.8 | 40.2 | 25.1 | 19.8 | 11.6 | 9.7 | 3.5 | 4.7 | 2.22 | 4.26 |
| 1996 | 34.7 | 69.5 | 64.4 | 42.3 | 35.4 | 16.7 | 15.2 | 4.7 | 1.6 | 3.2 | 3.52 |
| 1997 | 46 | 21.7 | 19.7 | 25.8 | 22.3 | 12.3 | 12 | 3.7 | 1.8 | 11.75 | 105.61 |
| 1998 | 149.3 | 32.3 | 13.2 | 18.5 | 17.3 | 15 | 9.1 | 9.9 | 2.5 | 20.24 | 23.79 |
| 1999 | 56.8 | 35.3 | 11.4 | 6.6 | 11.1 | 5.2 | 5.1 | 2.7 | 1.2 | 19.6 | 31.02 |
| 2000 | 50.3 | 58.2 | 27.2 | 14.1 | 8.1 | 7.9 | 7.8 | 4.9 | 5.5 | 1.97 | 17.75 |
| 2001 | 51.6 | 23.2 | 28.5 | 38 | 13.2 | 11.9 | 9.8 | 5.5 | 4.7 | 7.79 | 11.81 |
| 2002 | 83.3 | 34 | 29.9 | 31.6 | 22.8 | 7.4 | 4.1 | 5.4 | 5.5 | 1.49 | 12.94 |
| 2003 | 37.1 | 61.5 | 56.8 | 30.8 | 27.5 | 34.4 | 9.9 | 10.6 | 10.9 | 7.33 | 5.14 |
| 2004 | 41 | 32.9 | 43.9 | 46.5 | 37.2 | 26.4 | 27.3 | 8.1 | 15.5 | 11.51 | 20.76 |
| 2005 | 77.3 | 39.4 | 38.5 | 45.1 | 34.4 | 34.9 | 34.2 | 9.9 | 9 | 5.46 | 62.09 |

Table 14 cont'd

| Index | NYOHS | NYOHS | NYOHS | NYOHS | NYOHS | NEFSC | YOYNY | YOYNJ | YOYMD | YOYVA | YRLLI |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | 5 | 6 | 7 | 8 | 9:13 | 2:09 | 1 | 1 | 1 | 1 | 2 |
| Date Tuned | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan |
| Index Type | number | number | number | number | number | number | number | number | number | number | number |
| Year |  |  |  |  |  |  |  |  |  |  |  |
| 1982 | - | - | - | - | - | - | 8.86 | - | 0.59 | 1.56 | - |
| 1983 | - | - | - | - | - | - | 14.17 | 0.12 | 3.57 | 2.71 | - |
| 1984 | - | - | - | - | - | - | 16.25 | 0.03 | 0.61 | 3.4 | - |
| 1985 | - | - | - | - | - | - | 15 | 0.29 | 1.64 | 4.47 | - |
| 1986 | - | - | - | - | - | - | 1.92 | 0.18 | 0.91 | 2.41 | 0.61 |
| 1987 | - | - | - | - | - | - | 2.92 | 0.28 | 1.34 | 4.74 | 0.3 |
| 1988 | 12.77 | 9.91 | 3.14 | 1.24 | 0.42 | - | 15.9 | 0.41 | 1.46 | 15.74 | 0.21 |
| 1989 | 5.53 | 4.72 | 2.42 | 0.62 | 0.99 | - | 33.46 | 0.35 | 0.73 | 7.64 | 0.81 |
| 1990 | 1.5 | 1.62 | 1.04 | 0.95 | 0.43 | - | 21.35 | 1.03 | 4.87 | 11.23 | 1.78 |
| 1991 | 9.52 | 3.54 | 3.06 | 1.73 | 2.15 | 0.235 | 19.08 | 1 | 1.03 | 7.34 | 0.37 |
| 1992 | 6.16 | 1.31 | 0.42 | 0.64 | 2.22 | 0.237 | 3.6 | 0.47 | 1.52 | 3.76 | 1.26 |
| 1993 | 4.85 | 2.28 | 0.62 | 0.27 | 1.47 | 0.481 | 11.43 | 1.19 | 2.34 | 7.35 | 1.34 |
| 1994 | 7.09 | 1.71 | 0.8 | 0.23 | 1.25 | 1.394 | 12.59 | 1.78 | 13.97 | 18.11 | 0.75 |
| 1995 | 2.46 | 2.12 | 1.31 | 0.86 | 2.48 | 0.952 | 17.64 | 0.96 | 6.4 | 10.48 | 1.43 |
| 1996 | 3.32 | 0.94 | 0.86 | 0.46 | 0.69 | 0.602 | 16.23 | 1.98 | 4.41 | 5.45 | 1.29 |
| 1997 | 16.13 | 4.64 | 1.33 | 1.03 | 0.78 | 1.182 | 8.93 | 1.7 | 17.61 | 23 | 1.54 |
| 1998 | 44.23 | 6.56 | 1.81 | 0.36 | 1.13 | 0.729 | 22.3 | 1.01 | 3.91 | 9.35 | 1 |
| 1999 | 17.91 | 29.83 | 3.82 | 0.95 | 1.31 | 0.448 | 13.39 | 1.31 | 5.5 | 13.25 | 2.1 |
| 2000 | 4.87 | 1.68 | 1.24 | 0.14 | 0.58 | 1.274 | 26.64 | 1.9 | 5.34 | 2.8 | 2.05 |
| 2001 | 26.54 | 9.43 | 2.23 | 2.25 | 1.03 | 0.623 | 3.16 | 1.77 | 7.42 | 16.18 | 1.56 |
| 2002 | 4.19 | 6.05 | 2.09 | 0.78 | 0.85 | 0.981 | 22.98 | 1.07 | 12.57 | 14.17 | 2.16 |
| 2003 | 4.19 | 1.83 | 1.67 | 1.3 | 1.21 | 0.774 | 12.32 | 0.51 | 2.2 | 3.98 | 2.53 |
| 2004 | 7.12 | 5.25 | 2.31 | 3.68 | 6.23 | 0.335 | 17.36 | 2.43 | 10.83 | 22.89 | 1.19 |
| 2005 | 29.79 | 6.84 | 2.42 | 0.83 | 1.54 | 0.404 | 8.81 | 1.13 | 4.85 | 12.7 | 2.41 |

Table 14 cont'd.

| Index | YRLMD | DETRWL | CTTRL | DESSN | DESSN | DESSN | DESSN | DESSN | DESSN | DESSN | DESSN |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | 2 | 2:08 | 4:06 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| Date Tuned | 1-Jan | mean | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan |
| Index Type | number | number | number | number | number | number | number | number | number | number | number |
| Year |  |  |  |  |  |  |  |  |  |  |  |
| 1982 | 0.02 | 0.19 | 0 | - | - | - | - | - | - | - | - |
| 1983 | 0.02 | 0.01 | 0 | - | - | - | - | - | - | - | - |
| 1984 | 0.28 | 0.01 | 0.02 | - | - | - | - | - | - | - | - |
| 1985 | 0 | 0.01 | 0 | - | - | - | - | - | - | - | - |
| 1986 | 0.15 | 0 | 0 | - | - | - | - | - | - | - | - |
| 1987 | 0.03 | 0 | 0.05 | - | - | - | - | - | - | - | - |
| 1988 | 0.06 | 0 | 0.04 | - | - | - | - | - | - | - | - |
| 1989 | 0.07 | 0 | 0.06 | - | - | - | - | - | - | - | - |
| 1990 | 0.18 | 0 | 0.16 | - | - | - | - | - | - | - | - |
| 1991 | 0.28 | 1.17 | 0.15 | - | - | - | - | - | - | - | - |
| 1992 | 0.18 | 0.23 | 0.22 | - | - | - | - | - | - | - | - |
| 1993 | 0.14 | 0.89 | 0.27 | - | - | - | - | - | - | - | - |
| 1994 | 0.18 | 1.96 | 0.3 | - | - | - | - | - | - | - | - |
| 1995 | 0.58 | 2.59 | 0.59 | - | - | - | - | - | - | - | - |
| 1996 | 0.12 | 15.65 | 0.63 | 0.1 | 7.7 | 3.5 | 1.1 | 1.6 | 1.4 | 1.2 | 1.1 |
| 1997 | 0.08 | 7.2 | 0.85 | 2.0 | 1.6 | 8.6 | 3 | 1.1 | 1.4 | 1.6 | 0.7 |
| 1998 | 0.23 | 2.73 | 0.97 | 1.1 | 2.4 | 2.7 | 9.6 | 2.5 | 1.7 | 2.9 | 2.6 |
| 1999 | 0.16 | 2.04 | 1.1 | 0.0 | 1.6 | 2.2 | 2.7 | 3.6 | 1.1 | 0.8 | 1.2 |
| 2000 | 0.31 | 10.05 | 0.84 | 0.9 | 0.9 | 5.2 | 4.3 | 3.4 | 5.6 | 1.6 | 0.7 |
| 2001 | 0.23 | 6.03 | 0.61 | 0.1 | 2.3 | 2 | 3.7 | 2.2 | 2.8 | 4 | 1 |
| 2002 | 0.28 | 4.17 | 1.3 | 0.7 | 1.4 | 3.8 | 3.6 | 3.2 | 2.3 | 1.8 | 1.9 |
| 2003 | 0.58 | 7.21 | 0.87 | 0.5 | 2.4 | 2.4 | 3.3 | 2.2 | 2.7 | 3.1 | 2.6 |
| 2004 | 0.07 | 1.45 | 0.56 | 0.2 | 4.9 | 6.8 | 2.9 | 2 | 1.6 | 3.3 | 2.3 |
| 2005 | 0.55 | 2.14 | 0 | 0.0 | 0.0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 14 cont'd.

| Index | DESSN | NJTRL | NJTRL | NJTRL | NJTRL | NJTRL | NJTRL | NJTRL | NJTRL | MRFSS | CTCPUE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | 10:13 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9:13 | 2:13 | 2 |
| Date Tuned | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | 1-Jan | mean | mean |
| Index Type | number | number | number | number | number | number | number | number | number | number | number |
| Year |  |  |  |  |  |  |  |  |  |  |  |
| 1982 | - | - | - | - | - | - | - | - | - | - | 0.33 |
| 1983 | - | - | - | - | - | - | - | - | - | - | 0.4 |
| 1984 | - | - | - | - | - | - | - | - | - | - | 0.12 |
| 1985 | - | - | - | - | - | - | - | - | - | - | 0.06 |
| 1986 | - | - | - | - | - | - | - | - | - | - | 0.08 |
| 1987 | - | - | - | - | - | - | - | - | - | - | 0.04 |
| 1988 | - | - | - | - | - | - | - | - | - | 0.36 | 0.02 |
| 1989 | - | 0.04 | 0.06 | 0 | 0.02 | 0.01 | 0 | 0 | 0.01 | 0.27 | 0.27 |
| 1990 | - | 0.02 | 0.06 | 0.01 | 0.12 | 0.04 | 0.01 | 0.03 | 0.02 | 0.23 | 0.17 |
| 1991 | - | 0.17 | 0.17 | 0.01 | 0.01 | 0.09 | 0.08 | 0.01 | 0.06 | 0.39 | 0.15 |
| 1992 | - | 0.07 | 0.13 | 0.02 | 0.02 | 0.02 | 0.02 | 0 | 0 | 0.7 | 0.17 |
| 1993 | - | 0.08 | 0.11 | 0.05 | 0.03 | 0.04 | 0.01 | 0.01 | 0 | 0.58 | 0.07 |
| 1994 | - | 0.15 | 0.19 | 0.08 | 0.05 | 0.03 | 0.02 | 0.01 | 0.01 | 0.88 | 0.21 |
| 1995 | - | 0.48 | 0.13 | 0.05 | 0.03 | 0.02 | 0.01 | 0.01 | 0.01 | 1.17 | 0.6 |
| 1996 | 0.75 | 0.24 | 0.87 | 0.24 | 0.09 | 0.03 | 0.02 | 0.01 | 0 | 1.31 | 0.47 |
| 1997 | 1.89 | 0.27 | 0.59 | 0.34 | 0.16 | 0.05 | 0.02 | 0.01 | 0.01 | 1.37 | 0.18 |
| 1998 | 2.15 | 0.62 | 0.27 | 0.06 | 0.18 | 0.12 | 0.08 | 0.04 | 0.04 | 1.73 | 0.21 |
| 1999 | 2.09 | 0.06 | 0.26 | 0.15 | 0.21 | 0.09 | 0.03 | 0.01 | 0.01 | 1.61 | 0.38 |
| 2000 | 2.28 | 0.22 | 0.2 | 0.23 | 0.4 | 0.24 | 0.11 | 0.04 | 0.01 | 1.5 | 0 |
| 2001 | 1.9 | 0.1 | 0.11 | 0.13 | 0.2 | 0.07 | 0.03 | 0.02 | 0.01 | 1.26 | 0.89 |
| 2002 | 1.42 | 0.01 | 0.02 | 0.07 | 0.21 | 0.36 | 0.23 | 0.08 | 0.03 | 1.07 | 1.41 |
| 2003 | 5.44 | 0.58 | 0.46 | 0.08 | 0.24 | 0.19 | 0.24 | 0.1 | 0.04 | 0.91 | 1.33 |
| 2004 | 5.5 | 0.3 | 0.84 | 0.15 | 0.07 | 0.1 | 0.07 | 0.06 | 0.05 | 1.01 | 1.07 |
| 2005 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 14 cont'd.

| Index | CTCPUE | CTCPUE | CTCPUE | CTCPUE | CTCPUE | CTCPUE | CTCPUE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| Date Tuned | mean | mean | mean | mean | mean | mean | mean |
| Index Type | number | number | number | number | number | number | number |


| Year |  |  |  |  |  | 0.09 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1982 | 0.21 | 0.11 | 0.09 | 0.08 | 0.04 | 0.02 | 0.01 |
| 1983 | 0.19 | 0.08 | 0.04 | 0.03 | 0.01 | 0 |  |
| 1984 | 0.33 | 0.23 | 0.14 | 0.05 | 0.04 | 0.01 | 0 |
| 1985 | 0.32 | 0.22 | 0.12 | 0.09 | 0.04 | 0.03 | 0.01 |
| 1986 | 0.2 | 0.47 | 0.45 | 0.18 | 0.05 | 0.01 | 0.05 |
| 1987 | 0.24 | 0.34 | 0.2 | 0.14 | 0.06 | 0.04 | 0.03 |
| 1988 | 0.52 | 0.28 | 0.18 | 0.15 | 0.12 | 0.05 | 0.03 |
| 1989 | 0.48 | 0.47 | 0.16 | 0.18 | 0.13 | 0.09 | 0.03 |
| 1990 | 0.58 | 0.56 | 0.27 | 0.12 | 0.13 | 0.15 | 0.13 |
| 1991 | 0.67 | 0.43 | 0.35 | 0.14 | 0.07 | 0.09 | 0.13 |
| 1992 | 0.48 | 0.57 | 0.29 | 0.23 | 0.11 | 0.1 | 0.16 |
| 1993 | 0.7 | 0.62 | 0.49 | 0.28 | 0.22 | 0.1 | 0.08 |
| 1994 | 0.61 | 0.88 | 0.46 | 0.57 | 0.36 | 0.23 | 0.16 |
| 1995 | 1.2 | 1.34 | 0.59 | 0.59 | 0.32 | 0.18 | 0.19 |
| 1996 | 1.09 | 2.39 | 0.9 | 0.84 | 0.38 | 0.6 | 0.37 |
| 1997 | 1.11 | 1.28 | 1.64 | 0.58 | 0.31 | 0.23 | 0.21 |
| 1998 | 2.29 | 1.53 | 0.74 | 1.59 | 0.43 | 0.21 | 0.17 |
| 1999 | 0.43 | 1.28 | 0.37 | 0.39 | 0.6 | 0.62 | 0.41 |
| 2000 | 0.01 | 0.65 | 1.04 | 1.11 | 2.46 | 0.55 | 0.3 |
| 2001 | 0.67 | 0.56 | 2.24 | 1.12 | 0.67 | 0.65 | 0.41 |
| 2002 | 1.13 | 0.58 | 1.61 | 0.22 | 0.2 | 0.26 | 0.19 |
| 2003 | 1.36 | 0.63 | 0.75 | 0.41 | 0.39 | 0.38 | 0.34 |
| 2004 | 2.45 | 1.75 | 0.62 | 0.65 | 0.32 | 0.5 | 0.32 |
| 2005 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 15. Average F estimates (ages 8-11, 3-11) weighted by N.

| Year | Avg. F <br> $\mathbf{8 - 1 1}$ | Avg. F <br> $\mathbf{3 - 1 1}$ | wtd 7-11 | wtd 3-8 |
| :---: | :---: | :---: | :---: | :---: |
| $\mathbf{1 9 8 2}$ | 0.54 | 0.36 | 0.41 | 0.36 |
| $\mathbf{1 9 8 3}$ | 0.36 | 0.37 | 0.23 | 0.37 |
| $\mathbf{1 9 8 4}$ | 0.11 | 0.29 | 0.14 | 0.30 |
| $\mathbf{1 9 8 5}$ | 0.16 | 0.11 | 0.26 | 0.11 |
| $\mathbf{1 9 8 6}$ | 0.21 | 0.09 | 0.24 | 0.09 |
| $\mathbf{1 9 8 7}$ | 0.10 | 0.04 | 0.09 | 0.04 |
| $\mathbf{1 9 8 8}$ | 0.20 | 0.06 | 0.18 | 0.06 |
| $\mathbf{1 9 8 9}$ | 0.11 | 0.06 | 0.09 | 0.06 |
| $\mathbf{1 9 9 0}$ | 0.15 | 0.09 | 0.15 | 0.09 |
| $\mathbf{1 9 9 1}$ | 0.22 | 0.08 | 0.14 | 0.07 |
| $\mathbf{1 9 9 2}$ | 0.14 | 0.07 | 0.11 | 0.06 |
| $\mathbf{1 9 9 3}$ | 0.22 | 0.08 | 0.14 | 0.07 |
| $\mathbf{1 9 9 4}$ | 0.21 | 0.09 | 0.14 | 0.08 |
| $\mathbf{1 9 9 5}$ | 0.24 | 0.12 | 0.20 | 0.11 |
| $\mathbf{1 9 9 6}$ | 0.21 | 0.12 | 0.21 | 0.12 |
| $\mathbf{1 9 9 7}$ | 0.30 | 0.15 | 0.28 | 0.14 |
| $\mathbf{1 9 9 8}$ | 0.26 | 0.13 | 0.21 | 0.12 |
| $\mathbf{1 9 9 9}$ | 0.28 | 0.11 | 0.20 | 0.10 |
| $\mathbf{2 0 0 0}$ | 0.19 | 0.15 | 0.21 | 0.15 |
| $\mathbf{2 0 0 1}$ | 0.20 | 0.14 | 0.20 | 0.13 |
| $\mathbf{2 0 0 2}$ | 0.27 | 0.13 | 0.27 | 0.12 |
| $\mathbf{2 0 0 3}$ | 0.29 | 0.15 | 0.26 | 0.14 |
| $\mathbf{2 0 0 4}$ | 0.40 | 0.15 | 0.32 | 0.12 |

Table 16. Estimated fishing mortality (F) at age.

| AGE | $\mathbf{1 9 8 2}$ | $\mathbf{1 9 8 3}$ | $\mathbf{1 9 8 4}$ | $\mathbf{1 9 8 5}$ | $\mathbf{1 9 8 6}$ | $\mathbf{1 9 8 7}$ | $\mathbf{1 9 8 8}$ | $\mathbf{1 9 8 9}$ | $\mathbf{1 9 9 0}$ | $\mathbf{1 9 9 1}$ | $\mathbf{1 9 9 2}$ | $\mathbf{1 9 9 3}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{1}$ | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| $\mathbf{2}$ | 0.12 | 0.09 | 0.24 | 0.04 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 |
| $\mathbf{3}$ | 0.38 | 0.29 | 0.38 | 0.06 | 0.04 | 0.02 | 0.02 | 0.03 | 0.04 | 0.04 | 0.04 | 0.03 |
| $\mathbf{4}$ | 0.37 | 0.52 | 0.20 | 0.07 | 0.10 | 0.04 | 0.03 | 0.04 | 0.09 | 0.07 | 0.06 | 0.07 |
| $\mathbf{5}$ | 0.28 | 0.44 | 0.29 | 0.20 | 0.12 | 0.06 | 0.10 | 0.06 | 0.14 | 0.10 | 0.08 | 0.12 |
| $\mathbf{6}$ | 0.19 | 0.29 | 0.25 | 0.32 | 0.15 | 0.07 | 0.12 | 0.12 | 0.13 | 0.11 | 0.08 | 0.10 |
| $\mathbf{7}$ | 0.33 | 0.26 | 0.20 | 0.33 | 0.24 | 0.08 | 0.16 | 0.07 | 0.18 | 0.09 | 0.09 | 0.08 |
| $\mathbf{8}$ | 0.59 | 0.08 | 0.09 | 0.28 | 0.29 | 0.10 | 0.20 | 0.11 | 0.14 | 0.20 | 0.08 | 0.12 |
| $\mathbf{9}$ | 0.67 | 0.18 | 0.05 | 0.17 | 0.22 | 0.11 | 0.32 | 0.12 | 0.14 | 0.16 | 0.20 | 0.13 |
| $\mathbf{1 0}$ | 0.71 | 0.44 | 0.18 | 0.10 | 0.19 | 0.10 | 0.13 | 0.13 | 0.11 | 0.23 | 0.16 | 0.40 |
| $\mathbf{1 1}$ | 0.20 | 0.73 | 0.13 | 0.11 | 0.14 | 0.07 | 0.16 | 0.09 | 0.23 | 0.30 | 0.12 | 0.21 |
| $\mathbf{1 2}$ | 0.64 | 0.14 | 0.09 | 0.20 | 0.25 | 0.10 | 0.21 | 0.11 | 0.13 | 0.18 | 0.12 | 0.16 |
| $\mathbf{1 3 +}$ | 0.64 | 0.14 | 0.09 | 0.20 | 0.25 | 0.10 | 0.21 | 0.11 | 0.13 | 0.18 | 0.12 | 0.16 |


| AGE | $\mathbf{1 9 9 4}$ | $\mathbf{1 9 9 5}$ | $\mathbf{1 9 9 6}$ | $\mathbf{1 9 9 7}$ | $\mathbf{1 9 9 8}$ | $\mathbf{1 9 9 9}$ | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 1}$ | $\mathbf{2 0 0 2}$ | $\mathbf{2 0 0 3}$ | $\mathbf{2 0 0 4}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{1}$ | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.01 | 0.00 |
| $\mathbf{2}$ | 0.02 | 0.03 | 0.01 | 0.03 | 0.01 | 0.01 | 0.04 | 0.02 | 0.02 | 0.02 | 0.04 |
| $\mathbf{3}$ | 0.06 | 0.06 | 0.06 | 0.07 | 0.06 | 0.04 | 0.06 | 0.06 | 0.04 | 0.05 | 0.07 |
| $\mathbf{4}$ | 0.06 | 0.10 | 0.11 | 0.12 | 0.10 | 0.10 | 0.12 | 0.11 | 0.04 | 0.14 | 0.09 |
| $\mathbf{5}$ | 0.10 | 0.11 | 0.16 | 0.16 | 0.18 | 0.12 | 0.21 | 0.13 | 0.11 | 0.16 | 0.16 |
| $\mathbf{6}$ | 0.13 | 0.18 | 0.16 | 0.27 | 0.16 | 0.16 | 0.20 | 0.20 | 0.14 | 0.19 | 0.16 |
| $\mathbf{7}$ | 0.10 | 0.16 | 0.23 | 0.23 | 0.17 | 0.15 | 0.23 | 0.21 | 0.28 | 0.21 | 0.23 |
| $\mathbf{8}$ | 0.10 | 0.19 | 0.21 | 0.33 | 0.19 | 0.20 | 0.19 | 0.22 | 0.26 | 0.33 | 0.26 |
| $\mathbf{9}$ | 0.25 | 0.26 | 0.18 | 0.35 | 0.27 | 0.21 | 0.19 | 0.17 | 0.22 | 0.31 | 0.50 |
| $\mathbf{1 0}$ | 0.18 | 0.35 | 0.17 | 0.28 | 0.28 | 0.29 | 0.20 | 0.20 | 0.32 | 0.27 | 0.44 |
| $\mathbf{1 1}$ | 0.31 | 0.16 | 0.26 | 0.24 | 0.29 | 0.43 | 0.17 | 0.21 | 0.27 | 0.24 | 0.40 |
| $\mathbf{1 2}$ | 0.16 | 0.23 | 0.19 | 0.33 | 0.23 | 0.22 | 0.19 | 0.20 | 0.26 | 0.31 | 0.35 |
| $\mathbf{1 3 +}$ | 0.16 | 0.23 | 0.19 | 0.33 | 0.23 | 0.22 | 0.19 | 0.20 | 0.26 | 0.31 | 0.35 |

Table 17. 2004 Iterative re-weighting factors.

| index | age | chi wt. | rank | index | age | chi wt. | rank |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CT CPUE | 2 | 0.92 | 49 | MD SSN | 11 | 1.61 | 36 |
| CT CPUE | 3 | 0.87 | 50 | MD SSN | 12 | 3.69 | 15 |
| CT CPUE | 4 | 4.51 | 11 | MD SSN | 13+ | 1.50 | 37 |
| CT CPUE | 5 | 2.41 | 25 | MRFSS | 2-13+ | 11.62 | 2 |
| CT CPUE | 6 | 2.16 | 26 | NEFSC | 2-13+ | 3.54 | 17 |
| CT CPUE | 7 | 2.82 | 24 | NJ TRL | 2 | 0.99 | 45 |
| CT CPUE | 8 | 3.46 | 18 | NJ TRL | 3 | 1.84 | 30 |
| CT CPUE | 9 | 3.62 | 16 | NJ TRL | 4 | 1.80 | 31 |
| CT TRL | 4-6 | 3.06 | 20 | NJ TRL | 5 | 1.41 | 39 |
| DE SSN | 3 | 4.53 | 10 | NJ TRL | 6 | 1.68 | 34 |
| DE SSN | 4 | 6.59 | 6 | NJ TRL | 7 | 1.35 | 40 |
| DE SSN | 5 | 5.33 | 8 | NJ TRL | 8 | 1.85 | 29 |
| DE SSN | 6 | 17.45 | 1 | NJ TRL | 9-13+ | 0.97 | 48 |
| DE SSN | 7 | 8.08 | 4 | NY OHS | 3 | 1.86 | 28 |
| DE SSN | 8 | 7.28 | 5 | NY OHS | 4 | 2.03 | 27 |
| DE SSN | 9 | 4.47 | 12 | NY OHS | 5 | 1.47 | 38 |
| DE SSN | 10-13+ | 4.56 | 9 | NY OHS | 6 | 1.12 | 44 |
| MA COM | 5 | 3.43 | 19 | NY OHS | 7 | 1.33 | 41 |
| MA COM | 6 | 5.83 | 7 | NY OHS | 8 | 0.83 | 51 |
| MA COM | 7 | 8.80 | 3 | NY OHS | 9-13+ | 1.24 | 42 |
| MD SSN | 5 | 1.73 | 33 | MD YOY | 1 | 2.87 | 22 |
| MD SSN | 6 | 1.63 | 35 | NJ YOY | 1 | 2.87 | 23 |
| MD SSN | 7 | 1.74 | 32 | NY YOY | 1 | 1.24 | 43 |
| MD SSN | 8 | 0.98 | 47 | VA YOY | 1 | 4.02 | 14 |
| MD SSN | 9 | 0.98 | 46 | LI YRL | 2 | 4.16 | 13 |
| MD SSN | 10 | 0.81 | 52 | MD YRL | 2 | 2.89 | 21 |

Table 18a. Estimated population size at age in thousands.

| AGE | 1982 | 1983 | 1984 | 1985 | 1986 | 1987 | 1988 | 1989 | 1990 | 1991 | 1992 | 1993 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1,534 | 3,181 | 2,401 | 3,579 | 2,763 | 3,944 | 5,219 | 5,609 | 8,419 | 8,644 | 8,706 | 11,065 |
| 2 | 997 | 1,318 | 2,735 | 2,062 | 3,080 | 2,368 | 3,393 | 4,489 | 4,827 | 7,244 | 7,438 | 7,491 |
| 3 | 866 | 760 | 1,033 | 1,852 | 1,707 | 2,631 | 2,028 | 2,892 | 3,831 | 4,112 | 6,168 | 6,360 |
| 4 | 758 | 508 | 490 | 610 | 1,500 | 1,410 | 2,230 | 1,707 | 2,415 | 3,182 | 3,405 | 5,123 |
| 5 | 259 | 449 | 260 | 345 | 487 | 1,168 | 1,166 | 1,861 | 1,406 | 1,905 | 2,545 | 2,755 |
| 6 | 121 | 169 | 248 | 168 | 243 | 373 | 943 | 905 | 1,504 | 1,050 | 1,490 | 2,027 |
| 7 | 92 | 86 | 109 | 166 | 105 | 180 | 298 | 721 | 690 | 1,142 | 810 | 1,181 |
| 8 | 40 | 57 | 57 | 77 | 102 | 71 | 142 | 219 | 578 | 498 | 898 | 639 |
| 9 | 25 | 19 | 45 | 45 | 50 | 66 | 55 | 100 | 169 | 435 | 352 | 710 |
| 10 | 22 | 11 | 14 | 37 | 33 | 35 | 51 | 35 | 76 | 126 | 320 | 249 |
| 11 | 65 | 9 | 6 | 10 | 29 | 23 | 27 | 38 | 26 | 59 | 86 | 234 |
| 12 | 31 | 46 | 4 | 5 | 8 | 22 | 19 | 20 | 30 | 18 | 38 | 66 |
| 13+ | 35 | 111 | 145 | 60 | 49 | 140 | 54 | 89 | 118 | 142 | 146 | 124 |
| Total | 4,846 | 6,727 | 7,547 | 9,016 | 10,155 | 12,430 | 15,625 | 18,685 | 24,090 | 28,556 | 32,401 | 38,023 |
| 8+ | 218 | 253 | 271 | 234 | 271 | 357 | 348 | 501 | 997 | 1,278 | 1,840 | 2,022 |


| AGE | 1994 | 1995 | 1996 | 1997 | 1998 | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 | 2005 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 16,562 | 13,338 | 12,932 | 15,586 | 10,625 | 10,982 | 8,261 | 15,490 | 18,024 | 5,976 | 22,275 | 12,721 |
| 2 | 9,523 | 14,250 | 11,476 | 11,130 | 13,412 | 9,121 | 9,444 | 7,075 | 13,300 | 15,490 | 5,117 | 19,107 |
| 3 | 6,383 | 8,061 | 11,863 | 9,786 | 9,309 | 11,373 | 7,750 | 7,829 | 5,940 | 11,256 | 13,097 | 4,240 |
| 4 | 5,302 | 5,171 | 6,502 | 9,609 | 7,866 | 7,563 | 9,401 | 6,282 | 6,339 | 4,902 | 9,247 | 10,537 |
| 5 | 4,107 | 4,294 | 4,028 | 4,995 | 7,372 | 6,116 | 5,909 | 7,176 | 4,846 | 5,220 | 3,663 | 7,280 |
| 6 | 2,104 | 3,194 | 3,321 | 2,965 | 3,665 | 5,305 | 4,669 | 4,142 | 5,408 | 3,755 | 3,838 | 2,687 |
| 7 | 1,573 | 1,596 | 2,296 | 2,425 | 1,946 | 2,682 | 3,890 | 3,298 | 2,921 | 4,051 | 2,673 | 2,813 |
| 8 | 936 | 1,228 | 1,175 | 1,568 | 1,663 | 1,416 | 1,983 | 2,666 | 2,305 | 1,896 | 2,833 | 1,833 |
| 9 | 488 | 726 | 876 | 818 | 966 | 1,187 | 998 | 1,418 | 1,851 | 1,524 | 1,176 | 1,871 |
| 10 | 535 | 327 | 482 | 627 | 498 | 632 | 831 | 710 | 1,031 | 1,277 | 966 | 617 |
| 11 | 144 | 386 | 198 | 351 | 410 | 323 | 407 | 585 | 501 | 647 | 837 | 535 |
| 12 | 163 | 91 | 282 | 131 | 238 | 265 | 182 | 295 | 408 | 330 | 438 | 482 |
| 13+ | 107 | 74 | 116 | 111 | 345 | 269 | 191 | 286 | 525 | 382 | 462 | 547 |
| Total | 47,926 | 52,736 | 55,550 | 60,102 | 58,315 | 57,233 | 53,917 | 57,251 | 63,397 | 56,707 | 66,622 | 65,270 |
| 8+ | 2,373 | 2,832 | 3,129 | 3,606 | 4,120 | 4,092 | 4,592 | 5,960 | 6,621 | 6,056 | 6,712 | 5,885 |

Table 18b. Estimates of stock size (abundance in thousands), standard error, and coefficients of variation (CV) for Atlantic striped bass on 1 Jan 2005.

| Age | Abundance <br> (thousands) | Standard Error | CV |
| ---: | ---: | ---: | ---: |
|  |  |  |  |
| 1 | 12,721 | 3622.0 | 0.285 |
| 2 | 19,107 | 4071.1 | 0.213 |
| 3 | 4,240 | 895.7 | 0.211 |
| 4 | 10,537 | 1734.6 | 0.165 |
| 5 | 7,280 | 993.2 | 0.136 |
| 6 | 2,687 | 395.9 | 0.147 |
| 7 | 2,813 | 262.1 | 0.093 |
| 8 | 1,833 | 182.0 | 0.099 |
| 9 | 1,871 | 213.0 | 0.114 |
| 10 | 617 | 121.7 | 0.197 |
| 11 | 535 | 121.7 | 0.228 |

Table 19. Estimated female spawning stock population biomass at age (metric tons).

| AGE | 1982 | 1983 | 1984 | 1985 | 1986 | 1987 | 1988 | 1989 | 1990 | 1991 | 1992 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 4 | 17 | 11 | 12 | 20 | 46 | 44 | 71 | 51 | 72 | 95 | 102 |  |
| 5 | 33 | 51 | 30 | 40 | 60 | 176 | 183 | 282 | 197 | 271 | 386 |  |
| 6 | 95 | 100 | 147 | 108 | 134 | 212 | 633 | 720 | 1,089 | 607 | 981 |  |
| 7 | 174 | 134 | 186 | 277 | 163 | 253 | 444 | 1,409 | 1,355 | 2,057 | 1,340 |  |
| 8 | 97 | 129 | 117 | 177 | 222 | 137 | 257 | 506 | 1,444 | 1,149 | 2,103 |  |
| 9 | 62 | 53 | 130 | 130 | 128 | 164 | 127 | 251 | 473 | 1,264 | 1,029 |  |
| 10 | 82 | 36 | 45 | 124 | 98 | 98 | 133 | 110 | 216 | 351 | 1,087 |  |
| 11 | 335 | 38 | 24 | 40 | 103 | 76 | 96 | 128 | 98 | 205 | 317 |  |
| 12 | 152 | 228 | 21 | 21 | 33 | 89 | 78 | 87 | 127 | 66 | 192 |  |
| 13+ | 222 | 581 | 848 | 393 | 289 | 864 | 334 | 562 | 696 | 947 | 959 |  |
| Total | 1,266 | 1,361 | 1,556 | 1,326 | 1,273 | 2,110 | 2,353 | 4,103 | 5,765 | 7,009 | 8,494 |  |
| AGE | 1993 | 1994 | 1995 | 1996 | 1997 | 1998 | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 |
| 4 | 157 | 171 | 187 | 217 | 344 | 205 | 188 | 225 | 164 | 156 | 112 | 206 |
| 5 | 390 | 599 | 651 | 651 | 777 | 1,054 | 658 | 602 | 784 | 580 | 580 | 408 |
| 6 | 1,363 | 1,386 | 2,167 | 2,475 | 2,119 | 2,224 | 2,657 | 2,263 | 2,195 | 3,023 | 2,084 | 2,085 |
| 7 | 2,082 | 2,774 | 2,888 | 4,589 | 4,533 | 3,349 | 3,523 | 5,032 | 4,638 | 4,438 | 6,059 | 3,974 |
| 8 | 1,546 | 2,261 | 2,975 | 3,184 | 3,860 | 3,689 | 3,015 | 3,601 | 5,158 | 4,772 | 3,832 | 5,581 |
| 9 | 2,128 | 1,460 | 2,259 | 2,839 | 2,601 | 2,629 | 3,371 | 2,786 | 3,776 | 4,740 | 4,067 | 2,943 |
| 10 | 850 | 1,820 | 1,168 | 1,847 | 2,459 | 1,584 | 2,138 | 2,696 | 2,466 | 3,295 | 3,897 | 2,875 |
| 11 | 961 | 586 | 1,363 | 835 | 1,562 | 1,600 | 1,153 | 1,661 | 2,176 | 1,952 | 2,409 | 2,880 |
| 12 | 316 | 772 | 410 | 1,153 | 591 | 1,096 | 1,075 | 819 | 1,235 | 1,738 | 1,407 | 1,726 |
| 13+ | 843 | 638 | 570 | 741 | 755 | 1,903 | 1,498 | 1,210 | 1,449 | 2,805 | 1,939 | 2,260 |
| Total | 10,633 | 12,465 | 14,637 | 18,529 | 19,598 | 19,332 | 19,272 | 20,894 | 24,038 | 27,497 | 26,384 | 24,936 |

Table 20. Candidate models used in the analyses of striped bass tag recoveries.

| $\mathrm{S}()$.r (.) | Constant survival and reporting |
| :---: | :---: |
| $\mathrm{S}(\mathrm{t}) \mathrm{r}(\mathrm{t})$ | Time specific survival and reporting |
| $S() r.(t)$ | Constant survival and time specific reporting |
| $S(p) r(t)$ | *Regulatory period based survival and time specific reporting |
| $S(p) r(p)$ | *Regulatory period based survival and reporting |
| $S() r.(p)$ | *Constant survival and regulatory period based reporting |
| $S(t) r(p)$ | *Time specific survival and regulatory period reporting |
| $S(d) r(p)$ | **Regulatory period based survival with unique terminal year and regulatory period based reporting |
| S (v) r(p) | ***Regulatory period based survival with 2 terminal years unique and regulatory period based reporting |
| $\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp})$ | *Linear trend within regulatory period for both survival and reporting |
| $\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})$ | *Linear trend within regulatory period survival and regulatory period based reporting (no trend) |
| $\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{t})$ | *Linear trend within regulatory period survival and time specific reporting (no trend) |
| $\mathrm{S}(\mathrm{Va}) \mathrm{r}(\mathrm{Va})$ | Three period model for VA program (1990-1992, 1993-1994, 1995-2003) |
| * Periods (p) | $1=\{1987-1989\}, 2=\{1990-1994\}, 3=\{1995-1999\} 4=\{2000-2004\}$ |
| ** Periods | $1=\{1987-1989\}, 2=\{1990-1994\}, 3=\{1995-1999\}, 4=\{2000-2002\}, 5=2004$ |
| (d) *** Periods <br> (v) | $\begin{aligned} & 1=\{1987-1989\}, 2=\{1990-1994\}, 3=\{1995-1999\}, 4=\{2000-2002\}, 5= \\ & \{2003-2004\} \end{aligned}$ |

Table 20a. R/M estimates of exploitation rates of 28 inch fish from tagging programs. Exploitation rate is the proportion of tagged fish that were harvested or killed (with reporting rate adjustment of 0.43 , and hooking mortality rate adjustment of 0.08 ).

| Year | NJDB | NYOHS** | NCCOOP | MA** | VA Rap | MDCB | DE/PA | NYHUD | MEAN |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 | $*$ | $*$ | $*$ | $*$ | $*$ | $*$ | $*$ | $*$ |  |
| 1988 | $*$ | 0.05 | 0.06 | $*$ | $*$ | 0.07 | $*$ | 0.05 | $\mathbf{0 . 0 6}$ |
| 1989 | 0.02 | 0.04 | 0.05 | $*$ | $*$ | 0.04 | $*$ | 0.05 | $\mathbf{0 . 0 4}$ |
| 1990 | 0.04 | 0.07 | 0.09 | $*$ | 0.26 | 0.08 | $*$ | 0.07 | $\mathbf{0 . 1 0}$ |
| 1991 | 0.18 | 0.12 | 0.07 | $*$ | 0.36 | 0.12 | $*$ | 0.08 | $\mathbf{0 . 1 6}$ |
| 1992 | 0.02 | 0.11 | 0.14 | 0.04 | 0.37 | 0.12 | $*$ | 0.10 | $\mathbf{0 . 1 3}$ |
| 1993 | 0.09 | 0.14 | 0.11 | 0.07 | 0.37 | 0.12 | 0.17 | 0.11 | $\mathbf{0 . 1 5}$ |
| 1994 | 0.05 | 0.08 | 0.10 | 0.03 | 0.25 | 0.11 | 0.12 | 0.08 | $\mathbf{0 . 1 0}$ |
| 1995 | 0.10 | 0.20 | 0.14 | 0.05 | 0.41 | 0.20 | 0.14 | 0.13 | $\mathbf{0 . 1 7}$ |
| 1996 | 0.20 | 0.14 | 0.12 | 0.08 | 0.18 | 0.17 | 0.31 | 0.16 | $\mathbf{0 . 1 7}$ |
| 1997 | 0.23 | 0.29 | 0.21 | 0.13 | 0.38 | 0.23 | 0.27 | 0.22 | $\mathbf{0 . 2 4}$ |
| 1998 | 0.35 | 0.17 | 0.20 | 0.07 | 0.45 | 0.20 | 0.28 | 0.17 | $\mathbf{0 . 2 4}$ |
| 1999 | 0.08 | 0.29 | 0.24 | 0.09 | 0.28 | 0.32 | 0.15 | 0.14 | $\mathbf{0 . 2 0}$ |
| 2000 | 0.14 | 0.18 | 0.06 | 0.13 | 0.25 | 0.17 | 0.30 | 0.09 | $\mathbf{0 . 1 7}$ |
| 2001 | 0.14 | 0.11 | 0.15 | 0.07 | 0.23 | 0.11 | 0.26 | 0.10 | $\mathbf{0 . 1 5}$ |
| 2002 | 0.11 | 0.23 | 0.12 | 0.08 | 0.29 | 0.10 | 0.24 | 0.08 | $\mathbf{0 . 1 5}$ |
| 2003 | 0.14 | 0.15 | 0.12 | 0.11 | 0.22 | 0.10 | 0.16 | 0.10 | $\mathbf{0 . 1 4}$ |
| 2004 | 0.16 | 0.15 | 0.12 | 0.09 | 0.08 | 0.08 | 0.21 | 0.18 | $\mathbf{0 . 1 3}$ |

* Years when few or no striped bass were tagged and released.
** NYOHS and MA have fall tagging programs, and recapture interval of terminal year (2003) is fall 2002 to fall 2003; NCCOOP is a winter tagging program (Jan./Feb.) with recapture interval of terminal year (2003) from January 2003 to January 2004; others are spring tagging programs with recapture interval of terminal year (2003) from spring 2003 to spring 2004.

Table 20b. R/M estimates of catch rates of 28 inch fish from tagging programs. Catch rate is the proportion of tagged striped bass that were caught, but may have been released (with reporting rate adjustment of 0.43).

| Year | NJDB | NYOHS** | NCCOOP | MA** | VA Rap | MDCB | DE/PA | NYHUD | MEAN |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 | $*$ | $*$ | $*$ | $*$ | $*$ | 0.08 | $*$ | $*$ | $\mathbf{0 . 0 8}$ |
| 1988 | $*$ | 0.26 | 0.22 | $*$ | $*$ | 0.11 | $*$ | 0.15 | $\mathbf{0 . 1 9}$ |
| 1989 | 0.27 | 0.23 | 0.14 | $*$ | $*$ | 0.10 | $*$ | 0.21 | $\mathbf{0 . 1 9}$ |
| 1990 | 0.52 | 0.21 | 0.18 | $*$ | 0.49 | 0.18 | $*$ | 0.27 | $\mathbf{0 . 3 1}$ |
| 1991 | 0.47 | 0.24 | 0.20 | $*$ | 0.58 | 0.28 | $*$ | 0.25 | $\mathbf{0 . 3 4}$ |
| 1992 | 0.20 | 0.33 | 0.27 | 0.11 | 0.58 | 0.24 | $*$ | 0.24 | $\mathbf{0 . 2 8}$ |
| 1993 | 0.20 | 0.27 | 0.28 | 0.13 | 0.57 | 0.21 | 0.24 | 0.27 | $\mathbf{0 . 2 7}$ |
| 1994 | 0.21 | 0.21 | 0.22 | 0.10 | 0.36 | 0.22 | 0.20 | 0.21 | $\mathbf{0 . 2 2}$ |
| 1995 | 0.23 | 0.33 | 0.28 | 0.16 | 0.55 | 0.27 | 0.21 | 0.23 | $\mathbf{0 . 2 8}$ |
| 1996 | 0.33 | 0.31 | 0.16 | 0.18 | 0.21 | 0.25 | 0.41 | 0.27 | $\mathbf{0 . 2 6}$ |
| 1997 | 0.34 | 0.35 | 0.26 | 0.22 | 0.44 | 0.28 | 0.29 | 0.31 | $\mathbf{0 . 3 1}$ |
| 1998 | 0.38 | 0.17 | 0.26 | 0.15 | 0.60 | 0.23 | 0.34 | 0.25 | $\mathbf{0 . 3 0}$ |
| 1999 | 0.21 | 0.34 | 0.27 | 0.16 | 0.37 | 0.38 | 0.19 | 0.22 | $\mathbf{0 . 2 7}$ |
| 2000 | 0.22 | 0.33 | 0.13 | 0.14 | 0.39 | 0.20 | 0.36 | 0.20 | $\mathbf{0 . 2 4}$ |
| 2001 | 0.22 | 0.20 | 0.21 | 0.11 | 0.33 | 0.15 | 0.28 | 0.19 | $\mathbf{0 . 2 1}$ |
| 2002 | 0.18 | 0.38 | 0.15 | 0.16 | 0.36 | 0.13 | 0.24 | 0.18 | $\mathbf{0 . 2 2}$ |
| 2003 | 0.23 | 0.21 | 0.16 | 0.12 | 0.30 | 0.16 | 0.25 | 0.20 | $\mathbf{0 . 2 0}$ |
| 2004 | 0.26 | 0.26 | 0.17 | 0.11 | 0.13 | 0.11 | 0.25 | 0.27 | $\mathbf{0 . 1 9}$ |

[^0]Table 20c. R/M estimates of exploitation rates of 18 inch fish from tagging programs. Exploitation rate is the proportion of tagged fish that were harvested or killed (with reporting rate adjustment of 0.43 , and hooking mortality rate adjustment of 0.08 ).

| Year | NJDB | NYOHS** | NCCOOP | MA** | VA Rap | MDCB | DE/PA | NYHUD | MEAN |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 | $*$ | $*$ | $*$ | $*$ | $*$ | 0.01 | $*$ | $*$ | $\mathbf{0 . 0 1}$ |
| 1988 | $*$ | 0.02 | 0.04 | $*$ | $*$ | 0.01 | $*$ | 0.10 | $\mathbf{0 . 0 5}$ |
| 1989 | 0.03 | 0.03 | 0.03 | $*$ | $*$ | 0.01 | $*$ | 0.07 | $\mathbf{0 . 0 4}$ |
| 1990 | 0.07 | 0.04 | 0.07 | $*$ | 0.17 | 0.07 | $*$ | 0.11 | $\mathbf{0 . 0 9}$ |
| 1991 | 0.03 | 0.06 | 0.08 | $*$ | 0.14 | 0.10 | $*$ | 0.11 | $\mathbf{0 . 0 9}$ |
| 1992 | 0.03 | 0.04 | 0.15 | 0.04 | 0.31 | 0.13 | $*$ | 0.13 | $\mathbf{0 . 1 2}$ |
| 1993 | 0.03 | 0.05 | 0.11 | 0.05 | 0.23 | 0.11 | 0.12 | 0.17 | $\mathbf{0 . 1 1}$ |
| 1994 | 0.03 | 0.04 | 0.10 | 0.03 | 0.25 | 0.12 | 0.12 | 0.12 | $\mathbf{0 . 1 0}$ |
| 1995 | 0.06 | 0.05 | 0.14 | 0.04 | 0.19 | 0.18 | 0.13 | 0.16 | $\mathbf{0 . 1 2}$ |
| 1996 | 0.09 | 0.03 | 0.12 | 0.06 | 0.14 | 0.17 | 0.16 | 0.23 | $\mathbf{0 . 1 2}$ |
| 1997 | 0.08 | 0.04 | 0.17 | 0.09 | 0.20 | 0.20 | 0.09 | 0.29 | $\mathbf{0 . 1 4}$ |
| 1998 | 0.12 | 0.03 | 0.17 | 0.08 | 0.15 | 0.19 | 0.14 | 0.22 | $\mathbf{0 . 1 4}$ |
| 1999 | 0.05 | 0.05 | 0.22 | 0.06 | 0.13 | 0.16 | 0.12 | 0.22 | $\mathbf{0 . 1 3}$ |
| 2000 | 0.07 | 0.03 | 0.10 | 0.08 | 0.13 | 0.13 | 0.14 | 0.13 | $\mathbf{0 . 1 0}$ |
| 2001 | 0.09 | 0.05 | 0.12 | 0.05 | 0.18 | 0.12 | 0.14 | 0.14 | $\mathbf{0 . 1 1}$ |
| 2002 | 0.06 | 0.06 | 0.12 | 0.09 | 0.16 | 0.12 | 0.14 | 0.19 | $\mathbf{0 . 1 2}$ |
| 2003 | 0.07 | 0.04 | 0.11 | 0.08 | 0.15 | 0.13 | 0.13 | 0.14 | $\mathbf{0 . 1 1}$ |
| 2004 | 0.12 | 0.04 | 0.12 | 0.09 | 0.07 | 0.10 | 0.12 | 0.04 | $\mathbf{0 . 0 9}$ |

* Years when few or no striped bass were tagged and released.
** NYOHS and MA have fall tagging programs, and recapture interval of terminal year (2003) is fall 2002 to fall 2003; NCCOOP is a winter tagging program (Jan./Feb.) with recapture interval of terminal year (2003) from January 2003 to January 2004; others are spring tagging programs with recapture interval of terminal year (2003) from spring 2003 to spring 2004.

Table 20d. R/M estimates of catch rates of 18 inch fish from tagging programs. Catch rate is the proportion of tagged striped bass that were caught, but may have been released (with reporting rate adjustment of 0.43 ).

| Year | NJDB | NYOHS** | NCCOOP | MA** | VA Rap | MDCB | DE/PA | NYHUD | MEAN |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | ---: | ---: |
| 1987 | $*$ | $*$ | $*$ | $*$ | $*$ | 0.16 | $*$ | $*$ | $\mathbf{0 . 1 6}$ |
| 1988 | $*$ | 0.17 | 0.21 | $*$ | $*$ | 0.10 | $*$ | 0.21 | $\mathbf{0 . 1 7}$ |
| 1989 | 0.25 | 0.23 | 0.12 | $*$ | $*$ | 0.08 | $*$ | 0.25 | $\mathbf{0 . 1 9}$ |
| 1990 | 0.38 | 0.20 | 0.18 | $*$ | 0.38 | 0.15 | $*$ | 0.32 | $\mathbf{0 . 2 7}$ |
| 1991 | 0.20 | 0.17 | 0.20 | $*$ | 0.28 | 0.19 | $*$ | 0.24 | $\mathbf{0 . 2 1}$ |
| 1992 | 0.18 | 0.19 | 0.28 | 0.13 | 0.54 | 0.25 | $*$ | 0.30 | $\mathbf{0 . 2 6}$ |
| 1993 | 0.17 | 0.14 | 0.22 | 0.11 | 0.40 | 0.18 | 0.23 | 0.34 | $\mathbf{0 . 2 2}$ |
| 1994 | 0.17 | 0.17 | 0.21 | 0.11 | 0.37 | 0.22 | 0.25 | 0.26 | $\mathbf{0 . 2 2}$ |
| 1995 | 0.20 | 0.15 | 0.24 | 0.14 | 0.30 | 0.28 | 0.28 | 0.27 | $\mathbf{0 . 2 3}$ |
| 1996 | 0.24 | 0.17 | 0.16 | 0.18 | 0.25 | 0.27 | 0.26 | 0.32 | $\mathbf{0 . 2 3}$ |
| 1997 | 0.25 | 0.15 | 0.23 | 0.18 | 0.27 | 0.29 | 0.19 | 0.37 | $\mathbf{0 . 2 4}$ |
| 1998 | 0.27 | 0.15 | 0.32 | 0.15 | 0.24 | 0.28 | 0.26 | 0.29 | $\mathbf{0 . 2 5}$ |
| 1999 | 0.17 | 0.14 | 0.27 | 0.11 | 0.23 | 0.23 | 0.20 | 0.30 | $\mathbf{0 . 2 1}$ |
| 2000 | 0.20 | 0.14 | 0.17 | 0.10 | 0.23 | 0.23 | 0.24 | 0.21 | $\mathbf{0 . 1 9}$ |
| 2001 | 0.21 | 0.14 | 0.18 | 0.09 | 0.28 | 0.19 | 0.22 | 0.20 | $\mathbf{0 . 1 9}$ |
| 2002 | 0.13 | 0.18 | 0.18 | 0.16 | 0.25 | 0.17 | 0.19 | 0.24 | $\mathbf{0 . 1 9}$ |
| 2003 | 0.18 | 0.11 | 0.16 | 0.10 | 0.21 | 0.18 | 0.25 | 0.21 | $\mathbf{0 . 1 8}$ |
| 2004 | 0.25 | 0.14 | 0.17 | 0.11 | 0.11 | 0.16 | 0.18 | 0.22 | $\mathbf{0 . 1 7}$ |

[^1]Table 21. Survival (S) and fishing mortality (F) rates of striped bass $\geq 28$ inches, based on the assumption of constant natural mortality, adjusted for the bias due to live release of recaptured fish. Diagnostic statistic c-hat and bootstrap goodness of fit probability are presented for each dataset (see Methods).

## Coast Programs

Massachusetts; C-hat $=1.17$; bootstrap GOF probability $=0.68$ for the full parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1992 | 0.77 | 0.11 | 0.05 | 0.75 | -0.08 | 0.84 | $\mathbf{0 . 0 2}$ | -0.06 | 0.14 |
| 1993 | 0.77 | 0.11 | 0.07 | 0.57 | -0.09 | 0.85 | $\mathbf{0 . 0 1}$ | -0.07 | 0.12 |
| 1994 | 0.78 | 0.10 | 0.06 | 0.52 | -0.07 | 0.83 | $\mathbf{0 . 0 3}$ | -0.05 | 0.15 |
| 1995 | 0.73 | 0.16 | 0.06 | 0.38 | -0.05 | 0.78 | $\mathbf{0 . 1 0}$ | 0.05 | 0.16 |
| 1996 | 0.73 | 0.16 | 0.09 | 0.26 | -0.06 | 0.78 | $\mathbf{0 . 1 0}$ | 0.05 | 0.16 |
| 1997 | 0.73 | 0.16 | 0.07 | 0.22 | -0.04 | 0.76 | $\mathbf{0 . 1 2}$ | 0.07 | 0.19 |
| 1998 | 0.73 | 0.17 | 0.09 | 0.28 | -0.07 | 0.78 | $\mathbf{0 . 1 0}$ | 0.04 | 0.17 |
| 1999 | 0.73 | 0.17 | 0.08 | 0.28 | -0.05 | 0.77 | $\mathbf{0 . 1 1}$ | 0.05 | 0.19 |
| 2000 | 0.74 | 0.15 | 0.07 | 0.21 | -0.04 | 0.77 | $\mathbf{0 . 1 1}$ | 0.04 | 0.22 |
| 2001 | 0.74 | 0.15 | 0.05 | 0.33 | -0.04 | 0.77 | $\mathbf{0 . 1 1}$ | 0.03 | 0.21 |
| 2002 | 0.74 | 0.14 | 0.07 | 0.32 | -0.06 | 0.79 | $\mathbf{0 . 0 9}$ | 0.01 | 0.19 |
| 2003 | 0.75 | 0.14 | 0.05 | 0.18 | -0.02 | 0.77 | $\mathbf{0 . 1 2}$ | 0.03 | 0.23 |
| 2004 | 0.76 | 0.12 | 0.05 | 0.22 | -0.03 | 0.78 | $\mathbf{0 . 1 0}$ | 0.00 | 0.24 |

New York - Ocean Haul Seine
C-hat adjustment $=1.09$; bootstrap GOF probability $=0.19$ for the full parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1988 | 0.81 | 0.06 | 0.12 | 0.90 | -0.24 | 1.06 | $\mathbf{- 0 . 2 1}$ | -0.31 | -0.05 |
| 1989 | 0.81 | 0.06 | 0.10 | 0.86 | -0.19 | 1.01 | $\mathbf{- 0 . 1 6}$ | -0.25 | 0.00 |
| 1990 | 0.63 | 0.31 | 0.09 | 0.66 | -0.14 | 0.73 | $\mathbf{0 . 1 7}$ | 0.12 | 0.22 |
| 1991 | 0.63 | 0.31 | 0.11 | 0.53 | -0.15 | 0.74 | $\mathbf{0 . 1 6}$ | 0.11 | 0.21 |
| 1992 | 0.63 | 0.31 | 0.15 | 0.54 | -0.20 | 0.79 | $\mathbf{0 . 0 9}$ | 0.04 | 0.14 |
| 1993 | 0.63 | 0.31 | 0.11 | 0.43 | -0.12 | 0.71 | $\mathbf{0 . 1 9}$ | 0.14 | 0.25 |
| 1994 | 0.63 | 0.31 | 0.11 | 0.49 | -0.13 | 0.73 | $\mathbf{0 . 1 7}$ | 0.12 | 0.22 |
| 1995 | 0.65 | 0.28 | 0.15 | 0.34 | -0.14 | 0.76 | $\mathbf{0 . 1 3}$ | 0.07 | 0.19 |
| 1996 | 0.65 | 0.28 | 0.13 | 0.30 | -0.11 | 0.73 | $\mathbf{0 . 1 6}$ | 0.11 | 0.23 |
| 1997 | 0.65 | 0.28 | 0.14 | 0.21 | -0.09 | 0.71 | $\mathbf{0 . 1 9}$ | 0.13 | 0.25 |
| 1998 | 0.65 | 0.28 | 0.10 | 0.19 | -0.05 | 0.68 | $\mathbf{0 . 2 3}$ | 0.17 | 0.29 |
| 1999 | 0.65 | 0.28 | 0.14 | 0.10 | -0.04 | 0.68 | $\mathbf{0 . 2 4}$ | 0.18 | 0.30 |
| 2000 | 0.75 | 0.13 | 0.13 | 0.22 | -0.08 | 0.82 | $\mathbf{0 . 0 5}$ | -0.06 | 0.21 |
| 2001 | 0.75 | 0.13 | 0.10 | 0.24 | -0.06 | 0.81 | $\mathbf{0 . 0 7}$ | -0.04 | 0.23 |
| 2002 | 0.75 | 0.13 | 0.11 | 0.40 | -0.12 | 0.85 | $\mathbf{0 . 0 1}$ | -0.10 | 0.17 |
| 2003 | 0.76 | 0.13 | 0.08 | 0.21 | -0.04 | 0.79 | $\mathbf{0 . 0 9}$ | -0.03 | 0.26 |
| 2004 | 0.76 | 0.12 | 0.11 | 0.36 | -0.10 | 0.85 | $\mathbf{0 . 0 2}$ | -0.11 | 0.21 |

Table 21 cont'd.
New Jersey - Delaware Bay
C-hat adjustment $=1.0$ bootstrap GOF probability $=0.772$ for the fully parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1989 | 0.89 | -0.04 | 0.11 | 1.00 | -0.23 | 1.16 | $\mathbf{- 0 . 2 9}$ | -0.37 | -0.11 |
| 1990 | 0.67 | 0.25 | 0.12 | 0.50 | -0.15 | 0.78 | $\mathbf{0 . 1 0}$ | -0.12 | 0.49 |
| 1991 | 0.66 | 0.27 | 0.22 | 0.38 | -0.27 | 0.90 | $\mathbf{- 0 . 0 4}$ | -0.22 | 0.23 |
| 1992 | 0.65 | 0.28 | 0.08 | 1.00 | -0.17 | 0.78 | $\mathbf{0 . 1 0}$ | -0.03 | 0.27 |
| 1993 | 0.64 | 0.29 | 0.10 | 0.77 | -0.17 | 0.78 | $\mathbf{0 . 1 0}$ | 0.00 | 0.23 |
| 1994 | 0.63 | 0.31 | 0.10 | 0.79 | -0.17 | 0.77 | $\mathbf{0 . 1 2}$ | -0.03 | 0.31 |
| 1995 | 0.73 | 0.16 | 0.10 | 0.61 | -0.15 | 0.86 | $\mathbf{0 . 0 0}$ | -0.10 | 0.14 |
| 1996 | 0.70 | 0.21 | 0.13 | 0.42 | -0.14 | 0.81 | $\mathbf{0 . 0 6}$ | -0.01 | 0.15 |
| 1997 | 0.66 | 0.27 | 0.09 | 0.42 | -0.10 | 0.73 | $\mathbf{0 . 1 7}$ | 0.11 | 0.22 |
| 1998 | 0.62 | 0.33 | 0.16 | 0.30 | -0.14 | 0.72 | $\mathbf{0 . 1 8}$ | 0.10 | 0.27 |
| 1999 | 0.58 | 0.39 | 0.10 | 0.30 | -0.08 | 0.63 | $\mathbf{0 . 3 1}$ | 0.16 | 0.49 |
| 2000 | 0.88 | -0.02 | 0.10 | 0.30 | -0.07 | 0.95 | $\mathbf{- 0 . 1 0}$ | -0.17 | 0.04 |
| 2001 | 0.80 | 0.07 | 0.10 | 0.29 | -0.07 | 0.86 | $\mathbf{0 . 0 0}$ | -0.06 | 0.09 |
| 2002 | 0.68 | 0.23 | 0.08 | 0.34 | -0.06 | 0.73 | $\mathbf{0 . 1 7}$ | 0.08 | 0.27 |
| 2003 | 0.53 | 0.49 | 0.09 | 0.35 | -0.08 | 0.58 | $\mathbf{0 . 4 0}$ | 0.17 | 0.72 |
| 2004 | 0.38 | 0.83 | 0.11 | 0.36 | -0.10 | 0.42 | $\mathbf{0 . 7 2}$ | 0.24 | 1.40 |

North Carolina - Cooperative Trawl Cruise
C-hat adjustment $=1.187$; bootstrap GOF probability $=0.032$ for the full parameterized model .

| Year | S(unadj.) | F(unadj.) | $\begin{gathered} \text { Recovery } \\ \text { Rate } \\ \hline \end{gathered}$ | \% Live <br> Release | Bias Live Release | S(adj.) | F(adj.) | $\begin{gathered} 95 \% \mathrm{LCL} \\ \mathrm{~F}(\mathrm{adj}) \\ \hline \end{gathered}$ | $\begin{gathered} 95 \% \mathrm{UCL} \\ \mathrm{~F}(\mathrm{adj}) \\ \hline \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |  |
| 1988 | 0.72 | 0.17 | 0.09 | 0.76 | -0.16 | 0.87 | -0.01 | -0.16 | 0.26 |
| 1989 | 0.69 | 0.23 | 0.06 | 0.73 | -0.10 | 0.76 | 0.12 | -0.01 | 0.31 |
| 1990 | 0.70 | 0.21 | 0.08 | 0.60 | -0.11 | 0.78 | 0.10 | 0.03 | 0.18 |
| 1991 | 0.70 | 0.21 | 0.09 | 0.68 | -0.14 | 0.81 | 0.06 | 0.00 | 0.14 |
| 1992 | 0.71 | 0.19 | 0.11 | 0.45 | -0.12 | 0.81 | 0.07 | -0.01 | 0.16 |
| 1993 | 0.70 | 0.21 | 0.09 | 0.53 | -0.12 | 0.80 | 0.08 | 0.01 | 0.16 |
| 1994 | 0.69 | 0.22 | 0.08 | 0.49 | -0.10 | 0.77 | 0.12 | 0.03 | 0.21 |
| 1995 | 0.67 | 0.25 | 0.11 | 0.35 | -0.09 | 0.74 | 0.15 | 0.05 | 0.27 |
| 1996 | 0.65 | 0.28 | 0.06 | 0.20 | -0.03 | 0.67 | 0.25 | 0.18 | 0.33 |
| 1997 | 0.64 | 0.29 | 0.10 | 0.20 | -0.05 | 0.68 | 0.24 | 0.16 | 0.33 |
| 1998 | 0.65 | 0.28 | 0.11 | 0.26 | -0.08 | 0.70 | 0.20 | 0.12 | 0.30 |
| 1999 | 0.65 | 0.28 | 0.10 | 0.24 | -0.06 | 0.70 | 0.21 | 0.11 | 0.35 |
| 2000 | 0.64 | 0.30 | 0.05 | 0.36 | -0.04 | 0.67 | 0.25 | 0.13 | 0.40 |
| 2001 | 0.64 | 0.29 | 0.09 | 0.23 | -0.05 | 0.68 | 0.24 | 0.14 | 0.36 |
| 2002 | 0.65 | 0.28 | 0.07 | 0.20 | -0.03 | 0.67 | 0.25 | 0.14 | 0.38 |
| 2003 | 0.63 | 0.31 | 0.07 | 0.27 | -0.05 | 0.66 | 0.26 | 0.13 | 0.42 |
| 2004 | 0.64 | 0.30 | 0.07 | 0.23 | -0.04 | 0.66 | 0.26 | 0.13 | 0.43 |

Table 21 cont'd.

## Producer Area Programs

Delaware / Pennsylvania - Delaware River
C-hat $=1.00$; bootstrap GOF probability $=0.384$ for the fully parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1993 | 0.84 | 0.06 | 0.106 | 0.33 | -0.090 | 0.92 | $\mathbf{- 0 . 0 7}$ | -0.34 | 0.30 |
| 1994 | 0.83 | 0.07 | 0.108 | 0.29 | -0.081 | 0.91 | $\mathbf{- 0 . 0 5}$ | -0.33 | 0.34 |
| 1995 | 0.59 | 0.38 | 0.117 | 0.35 | -0.107 | 0.65 | $\mathbf{0 . 2 7}$ | 0.18 | 0.37 |
| 1996 | 0.59 | 0.38 | 0.138 | 0.28 | -0.109 | 0.66 | $\mathbf{0 . 2 6}$ | 0.17 | 0.36 |
| 1997 | 0.59 | 0.38 | 0.108 | 0.28 | -0.079 | 0.64 | $\mathbf{0 . 3 0}$ | 0.22 | 0.39 |
| 1998 | 0.59 | 0.38 | 0.145 | 0.17 | -0.074 | 0.63 | $\mathbf{0 . 3 1}$ | 0.22 | 0.40 |
| 1999 | 0.59 | 0.38 | 0.079 | 0.21 | -0.042 | 0.61 | $\mathbf{0 . 3 4}$ | 0.26 | 0.43 |
| 2000 | 0.59 | 0.38 | 0.136 | 0.17 | -0.068 | 0.63 | $\mathbf{0 . 3 1}$ | 0.23 | 0.41 |
| 2001 | 0.59 | 0.38 | 0.117 | 0.12 | -0.040 | 0.61 | $\mathbf{0 . 3 4}$ | 0.26 | 0.44 |
| 2002 | 0.59 | 0.38 | 0.100 | 0.18 | -0.048 | 0.62 | $\mathbf{0 . 3 4}$ | 0.25 | 0.43 |
| 2003 | 0.59 | 0.39 | 0.108 | 0.32 | -0.090 | 0.64 | $\mathbf{0 . 2 9}$ | 0.18 | 0.41 |
| 2004 | 0.59 | 0.38 | 0.108 | 0.22 | -0.064 | 0.63 | $\mathbf{0 . 3 1}$ | 0.16 | 0.47 |

Maryland - Chesapeake Bay Spring Spawning Stock
C-hat $=1.0$; bootstrap GOF probability $=0.83$ for the fully parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1987 | 0.97 | -0.12 | 0.03 |  | 0.00 | 0.97 | $\mathbf{- 0 . 1 2}$ | -0.14 | -0.04 |
| 1988 | 0.96 | -0.11 | 0.04 | 0.67 | -0.06 | 1.02 | $\mathbf{- 0 . 1 7}$ | -0.19 | -0.14 |
| 1989 | 0.95 | -0.10 | 0.05 | 0.79 | -0.09 | 1.04 | $\mathbf{- 0 . 1 9}$ | -0.21 | -0.16 |
| 1990 | 0.53 | 0.49 | 0.07 | 0.57 | -0.09 | 0.58 | $\mathbf{0 . 3 9}$ | 0.26 | 0.54 |
| 1991 | 0.59 | 0.38 | 0.12 | 0.59 | -0.18 | 0.72 | $\mathbf{0 . 1 8}$ | 0.11 | 0.26 |
| 1992 | 0.65 | 0.29 | 0.11 | 0.51 | -0.14 | 0.75 | $\mathbf{0 . 1 3}$ | 0.08 | 0.19 |
| 1993 | 0.70 | 0.21 | 0.10 | 0.46 | -0.11 | 0.79 | $\mathbf{0 . 0 9}$ | 0.03 | 0.16 |
| 1994 | 0.75 | 0.14 | 0.09 | 0.47 | -0.11 | 0.84 | $\mathbf{0 . 0 3}$ | -0.05 | 0.13 |
| 1995 | 0.65 | 0.29 | 0.12 | 0.26 | -0.08 | 0.70 | $\mathbf{0 . 2 0}$ | 0.12 | 0.29 |
| 1996 | 0.64 | 0.29 | 0.10 | 0.28 | -0.07 | 0.69 | $\mathbf{0 . 2 2}$ | 0.16 | 0.29 |
| 1997 | 0.64 | 0.30 | 0.11 | 0.22 | -0.07 | 0.68 | $\mathbf{0 . 2 4}$ | 0.19 | 0.29 |
| 1998 | 0.63 | 0.31 | 0.10 | 0.19 | -0.05 | 0.66 | $\mathbf{0 . 2 6}$ | 0.22 | 0.31 |
| 1999 | 0.63 | 0.32 | 0.12 | 0.18 | -0.06 | 0.67 | $\mathbf{0 . 2 6}$ | 0.20 | 0.31 |
| 2000 | 0.62 | 0.33 | 0.08 | 0.19 | -0.04 | 0.65 | $\mathbf{0 . 2 9}$ | 0.22 | 0.37 |
| 2001 | 0.61 | 0.34 | 0.07 | 0.25 | -0.05 | 0.64 | $\mathbf{0 . 2 9}$ | 0.20 | 0.39 |
| 2002 | 0.61 | 0.34 | 0.06 | 0.36 | -0.05 | 0.64 | $\mathbf{0 . 2 9}$ | 0.18 | 0.43 |
| 2003 | 0.60 | 0.35 | 0.07 | 0.20 | -0.03 | 0.63 | $\mathbf{0 . 3 2}$ | 0.18 | 0.49 |
| 2004 | 0.60 | 0.36 | 0.05 | 0.17 | -0.02 | 0.61 | $\mathbf{0 . 3 4}$ | 0.19 | 0.55 |

Table 21 cont'd.
Virginia - Rappahannock River
C-hat adjustment $=1.289$; bootstrap GOF probability $=0.17$ for the fully parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1990 | 0.62 | 0.32 | 0.09 | 0.58 | -0.13 | 0.72 | $\mathbf{0 . 1 9}$ | 0.12 | 0.26 |
| 1991 | 0.62 | 0.32 | 0.09 | 0.56 | -0.13 | 0.72 | $\mathbf{0 . 1 8}$ | 0.12 | 0.25 |
| 1992 | 0.62 | 0.32 | 0.12 | 0.53 | -0.17 | 0.75 | $\mathbf{0 . 1 3}$ | 0.07 | 0.20 |
| 1993 | 0.62 | 0.32 | 0.10 | 0.35 | -0.09 | 0.69 | $\mathbf{0 . 2 2}$ | 0.16 | 0.30 |
| 1994 | 0.62 | 0.32 | 0.08 | 0.32 | -0.07 | 0.67 | $\mathbf{0 . 2 5}$ | 0.18 | 0.32 |
| 1995 | 0.61 | 0.34 | 0.13 | 0.20 | -0.08 | 0.66 | $\mathbf{0 . 2 6}$ | 0.19 | 0.35 |
| 1996 | 0.61 | 0.34 | 0.05 | 0.13 | -0.02 | 0.62 | $\mathbf{0 . 3 3}$ | 0.25 | 0.41 |
| 1997 | 0.61 | 0.34 | 0.08 | 0.17 | -0.04 | 0.63 | $\mathbf{0 . 3 1}$ | 0.23 | 0.39 |
| 1998 | 0.61 | 0.34 | 0.13 | 0.22 | -0.09 | 0.67 | $\mathbf{0 . 2 6}$ | 0.18 | 0.34 |
| 1999 | 0.61 | 0.35 | 0.10 | 0.20 | -0.06 | 0.65 | $\mathbf{0 . 2 9}$ | 0.21 | 0.38 |
| 2000 | 0.63 | 0.31 | 0.08 | 0.35 | -0.07 | 0.68 | $\mathbf{0 . 2 4}$ | 0.15 | 0.34 |
| 2001 | 0.63 | 0.31 | 0.07 | 0.30 | -0.05 | 0.67 | $\mathbf{0 . 2 6}$ | 0.17 | 0.36 |
| 2002 | 0.63 | 0.31 | 0.09 | 0.30 | -0.08 | 0.68 | $\mathbf{0 . 2 3}$ | 0.14 | 0.33 |
| 2003 | 0.63 | 0.31 | 0.09 | 0.25 | -0.06 | 0.67 | $\mathbf{0 . 2 5}$ | 0.16 | 0.36 |
| 2004 | 0.64 | 0.30 | 0.06 | 0.29 | -0.04 | 0.67 | $\mathbf{0 . 2 5}$ | 0.15 | 0.38 |

Hudson River
C-hat adjustment $=1.223$; bootstrap GOF probability $=0.206$ for the fully parameterized model .

|  |  | Recovery <br> Year Live |  |  | Bias Live |  | 95\%LCL |  | 95\%UCL |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.72 | 0.18 | 0.09 | 0.56 | -0.12 | 0.82 | $\mathbf{0 . 0 5}$ | -0.06 | 0.24 |
| 1989 | 0.72 | 0.18 | 0.11 | 0.74 | -0.19 | 0.88 | $\mathbf{- 0 . 0 3}$ | -0.14 | 0.17 |
| 1990 | 0.63 | 0.31 | 0.13 | 0.66 | -0.21 | 0.80 | $\mathbf{0 . 0 8}$ | 0.02 | 0.12 |
| 1991 | 0.63 | 0.31 | 0.10 | 0.50 | -0.13 | 0.72 | $\mathbf{0 . 1 7}$ | 0.12 | 0.21 |
| 1992 | 0.63 | 0.31 | 0.13 | 0.58 | -0.19 | 0.78 | $\mathbf{0 . 0 9}$ | 0.05 | 0.14 |
| 1993 | 0.63 | 0.30 | 0.13 | 0.49 | -0.16 | 0.76 | $\mathbf{0 . 1 3}$ | 0.08 | 0.17 |
| 1994 | 0.64 | 0.30 | 0.12 | 0.52 | -0.16 | 0.76 | $\mathbf{0 . 1 3}$ | 0.08 | 0.18 |
| 1995 | 0.65 | 0.28 | 0.11 | 0.38 | -0.11 | 0.73 | $\mathbf{0 . 1 6}$ | 0.13 | 0.21 |
| 1996 | 0.65 | 0.28 | 0.13 | 0.25 | -0.09 | 0.71 | $\mathbf{0 . 1 9}$ | 0.16 | 0.23 |
| 1997 | 0.65 | 0.28 | 0.16 | 0.32 | -0.14 | 0.75 | $\mathbf{0 . 1 3}$ | 0.10 | 0.18 |
| 1998 | 0.65 | 0.29 | 0.13 | 0.23 | -0.08 | 0.70 | $\mathbf{0 . 2 0}$ | 0.16 | 0.24 |
| 1999 | 0.64 | 0.29 | 0.13 | 0.31 | -0.11 | 0.73 | $\mathbf{0 . 1 7}$ | 0.13 | 0.21 |
| 2000 | 0.68 | 0.24 | 0.08 | 0.36 | -0.07 | 0.73 | $\mathbf{0 . 1 6}$ | 0.11 | 0.26 |
| 2001 | 0.66 | 0.26 | 0.07 | 0.26 | -0.05 | 0.70 | $\mathbf{0 . 2 1}$ | 0.14 | 0.29 |
| 2002 | 0.65 | 0.28 | 0.11 | 0.35 | -0.09 | 0.72 | $\mathbf{0 . 1 8}$ | 0.09 | 0.24 |
| 2003 | 0.63 | 0.31 | 0.10 | 0.33 | -0.08 | 0.68 | $\mathbf{0 . 2 3}$ | 0.10 | 0.28 |
| 2004 | 0.61 | 0.35 | 0.12 | 0.25 | -0.08 | 0.66 | $\mathbf{0 . 2 7}$ | 0.10 | 0.31 |

Table 22. Akaike weights based on the statistical fit of the tag-recapture data to the various models in the suite. These weights are used to obtain the weighted model-averaged estimates of survival. Results are for striped bass tagged at $\geq 28$ inches.

| Coast Programs |  |  |  |  |
| :--- | :---: | :---: | :---: | ---: |
| Model | MADFW | NYOHS | NJDEL | NCCOOP |
| $\{\mathrm{S}() .\mathrm{r}()\}$. | 0.0002 | 0.00528 | 0.0023 | 0 |
| $\{\mathrm{~S}() .\mathrm{r}(\mathrm{p})\}$ | 0.0002 | 0.00693 | 0.0004 | $\mathbf{0 . 1 2 5 2}$ |
| $\{\mathrm{~S}() .\mathrm{r}(\mathrm{t})\}$ | $\mathbf{0 . 1 3 3 0}$ | 0.00013 | $\mathbf{0 . 3 7 0 7}$ | 0.01188 |
| $\{\mathrm{~S}(\mathrm{p}) \mathrm{r}(\mathrm{p})\}$ | 0.0329 | $\mathbf{0 . 1 5 0 6 0}$ | 0.0004 | $\mathbf{0 . 3 6 7 3 2}$ |
| $\{\mathrm{~S}(\mathrm{p}) \mathrm{r}(\mathrm{t})\}$ | $\mathbf{0 . 1 9 2 1}$ | 0.01215 | $\mathbf{0 . 1 5 5 9}$ | 0.01759 |
| $\{\mathrm{~S}(\mathrm{~d}) \mathrm{r}(\mathrm{p})\}$ | 0.0361 | $\mathbf{0 . 1 4 3 5 2}$ | 0.0003 | $\mathbf{0 . 1 3 5 8 8}$ |
| $\{\mathrm{~S}(\mathrm{v}) \mathrm{r}(\mathrm{p})\}$ | 0.0188 | $\mathbf{0 . 1 7 7 2 1}$ | 0.0022 | $\mathbf{0 . 1 6 3 2 7}$ |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{t})\}$ | $\mathbf{0 . 1 6 1 0}$ | 0.00387 | 0.0818 | 0.02038 |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp})\}$ | 0.0161 | $\mathbf{0 . 4 0 9 0 6}$ | 0.0387 | 0.07094 |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})\}$ | 0.0165 | 0.06350 | 0.0230 | 0.03688 |
| $\{\mathrm{~S}(\mathrm{t}) \mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 3 9 2 4}$ | 0.00531 | $\mathbf{0 . 3 0 4 9}$ | 0.00103 |
| $\{\mathrm{~S}(\mathrm{t}) \mathrm{r}(\mathrm{t})\}$ | 0.0006 | 0.02244 | 0.0194 | 0.04963 | l

Producer Area Programs

| Model |  | DE/PA | HUDSON | MDCB | VARAP |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\{\mathrm{S}() .\mathrm{r}()$. |  | 0.088 | 0.0000 | 0 | 0.02657 |
| $\{\mathrm{S}() .\mathrm{r}(\mathrm{p})$ \} |  | 0.039 | 0.1842 | 0 | 0.55075 |
| $\{\mathrm{S}() .\mathrm{r}(\mathrm{t})$ \} |  | 0.000 | 0.0008 | 0 | 0.00044 |
| $\{\mathrm{S}(\mathrm{p}) \mathrm{r}(\mathrm{p})$ \} |  | 0.430 | 0.2088 | 0.01861 | 0.23871 |
| $\{\mathrm{S}(\mathrm{p}) \mathrm{r}(\mathrm{t})$ \} |  | 0.000 | 0.0004 | 0.0006 | 0.00011 |
| $\{\mathrm{S}(\mathrm{d}) \mathrm{r}(\mathrm{p})$ \} |  | 0.184 | 0.1251 | 0.0368 | 0.11287 |
| \{S(v)r(p) \} |  | 0.158 | 0.1115 | 0.00733 | 0.04137 |
| S(Va)r(va) | NA |  | NA | NA | 0.00701 |
| \{S(Tp)r(t) \} |  | 0.000 | 0.0009 | 0.72424 | 0.00002 |
| $\{\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp})$ \} |  | 0.039 | 0.0878 | 0.20179 | 0.00183 |
| $\{\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})$ \} |  | 0.061 | 0.2691 | 0.00564 | 0.02027 |
| $\{\mathrm{S}(\mathrm{t}) \mathrm{r}(\mathrm{p})$ \} |  | 0.000 | 0.0115 | 0.00467 | 0.00004 |
| $\{\mathrm{S}(\mathrm{t}) \mathrm{r}(\mathrm{t})\}$ |  | 0.000 | 0.0000 | 0.00031 | 0 |

Model Descriptions

| S(.) r(.) | Constant survival and reporting |
| :--- | :--- |
| S(t) r(t) | Time specific survival and reporting |
| S(.) r(t) | Constant survival and time specific reporting |
| S(p) $\mathrm{r}(\mathrm{t})$ | Regulatory period based survival and time specific reporting |
| $\mathrm{S}(\mathrm{p}) \mathrm{r}(\mathrm{p})$ | Regulatory period based survival and reporting |
| $\mathrm{S}() .\mathrm{r}(\mathrm{p})$ | Constant survival and regulatory period based reporting |
| $\mathrm{S}(\mathrm{t}) \mathrm{r}(\mathrm{p})$ | Time specific survival and regulatory period based reporting |
| $\mathrm{S}(\mathrm{d}) \mathrm{r}(\mathrm{p})$ | Regulatory period survival with terminal year unique and regulatory period reporting |
| $\mathrm{S}(\mathrm{v}) \mathrm{r}(\mathrm{p})$ | Regulatory period survival with 2 terminal years unique and regulatory period reporting |
| $\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp})$ | Linear trend within regulatory period on both survival and reporting |
| $\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})$ | Linear trend within regulatory period survival and regulatory period reporting (no trend) |
| $\mathrm{S}(\mathrm{Tp}) \mathrm{r}(\mathrm{t})$ | Linear trend within regulatory period survival and time specific reporting (no trend) |
| $\mathrm{S}(\mathrm{Va}) \mathrm{r}(\mathrm{Va})$ | Three period model for VA program (90-92, 93-94, 95-04) |

Table 23. Summary table, based on the assumption of constant natural mortality, with estimates of annual instantaneous fishing mortality, F , of striped bass $\geq 28$ inches by individual program, averaged over all coastal programs, with a weighted average over producer areas and with an overall coastwide estimate. Estimates of coastwide abundance of age 7+ striped bass from 19872004 are also provided.

| Year | MADFW | NYOHS | NJDEL | NCCOOP | Unweighted <br> average | lower <br> $95 \%$ CI | upper <br> $95 \%$ CI |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 |  |  |  | -0.01 | $\mathbf{0 . 0 1}$ |  |  |
| 1989 |  |  | -0.29 | 0.12 | $\mathbf{- 0 . 0 9}$ |  |  |
| 1990 |  | 0.17 | 0.10 | 0.10 | $\mathbf{0 . 1 2}$ |  |  |
| 1991 |  | 0.16 | -0.04 | 0.06 | $\mathbf{0 . 0 6}$ |  |  |
| 1992 | 0.02 | 0.09 | 0.10 | 0.07 | $\mathbf{0 . 0 7}$ | 0.00 | 0.17 |
| 1993 | 0.01 | 0.19 | 0.10 | 0.08 | $\mathbf{0 . 1 0}$ | 0.00 | 0.28 |
| 1994 | 0.03 | 0.17 | 0.12 | 0.12 | $\mathbf{0 . 1 1}$ | 0.01 | 0.31 |
| 1995 | 0.10 | 0.13 | 0.00 | 0.15 | $\mathbf{0 . 1 0}$ | 0.05 | 0.21 |
| 1996 | 0.10 | 0.16 | 0.06 | 0.25 | $\mathbf{0 . 1 4}$ | 0.08 | 0.31 |
| 1997 | 0.12 | 0.19 | 0.17 | 0.24 | $\mathbf{0 . 1 8}$ | 0.11 | 0.42 |
| 1998 | 0.10 | 0.23 | 0.18 | 0.20 | $\mathbf{0 . 1 8}$ | 0.01 | 0.44 |
| 1999 | 0.11 | 0.24 | 0.31 | 0.21 | $\mathbf{0 . 2 2}$ | 0.12 | 0.28 |
| 2000 | 0.11 | 0.05 | -0.10 | 0.25 | $\mathbf{0 . 0 8}$ | 0.09 | 0.50 |
| 2001 | 0.11 | 0.07 | 0.00 | 0.24 | $\mathbf{0 . 1 0}$ | 0.04 | 0.41 |
| 2002 | 0.09 | 0.01 | 0.17 | 0.25 | $\mathbf{0 . 1 3}$ | 0.02 | 0.29 |
| 2003 | 0.12 | 0.09 | 0.40 | 0.26 | $\mathbf{0 . 2 2}$ | 0.02 | 0.29 |
| 2004 | 0.10 | 0.02 | 0.72 | 0.26 | $\mathbf{0 . 2 7}$ | 0.03 | 0.33 |

Producer Area Programs

| Year | HUDSON | DE/PA | MDCB | VARAP | Weighted <br> average* | lower <br> $95 \%$ CI | upper <br> $95 \%$ CI |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 |  |  | -0.12 |  | $\mathbf{0 . 1 2}$ |  |  |
| 1988 | 0.05 |  | -0.17 |  | $\mathbf{- 0 . 1 7}$ |  |  |
| 1989 | -0.03 |  | -0.19 |  | $\mathbf{0 . 1 9}$ |  |  |
| 1990 | 0.08 |  | 0.39 | 0.19 | $\mathbf{0 . 2 6}$ |  |  |
| 1991 | 0.17 |  | 0.18 | 0.18 | $\mathbf{0 . 1 7}$ |  |  |
| 1992 | 0.09 |  | 0.13 | 0.13 | $\mathbf{0 . 1 2}$ |  |  |
| 1993 | 0.13 | -0.07 | 0.09 | 0.22 | $\mathbf{0 . 1 1}$ | 0.06 | 0.23 |
| 1994 | 0.13 | -0.05 | 0.03 | 0.25 | $\mathbf{0 . 0 9}$ | 0.03 | 0.22 |
| 1995 | 0.16 | 0.27 | 0.20 | 0.26 | $\mathbf{0 . 2 2}$ | 0.13 | 0.29 |
| 1996 | 0.19 | 0.26 | 0.22 | 0.33 | $\mathbf{0 . 2 5}$ | 0.18 | 0.32 |
| 1997 | 0.13 | 0.30 | 0.24 | 0.31 | $\mathbf{0 . 2 5}$ | 0.19 | 0.32 |
| 1998 | 0.20 | 0.31 | 0.26 | 0.26 | $\mathbf{0 . 2 6}$ | 0.20 | 0.32 |
| 1999 | 0.17 | 0.34 | 0.26 | 0.29 | $\mathbf{0 . 2 6}$ | 0.20 | 0.35 |
| 2000 | 0.16 | 0.31 | 0.29 | 0.24 | $\mathbf{0 . 2 6}$ | 0.20 | 0.38 |
| 2001 | 0.21 | 0.34 | 0.29 | 0.26 | $\mathbf{0 . 2 8}$ | 0.20 | 0.42 |
| 2002 | 0.18 | 0.34 | 0.29 | 0.23 | $\mathbf{0 . 2 7}$ | 0.17 | 0.44 |
| 2003 | 0.23 | 0.29 | 0.32 | 0.25 | $\mathbf{0 . 2 9}$ | 0.17 | 0.50 |
| 2004 | 0.27 | 0.31 | 0.34 | 0.25 | $\mathbf{0 . 3 1}$ | 0.17 | 0.50 |

* Weighting Scheme: Hudson (0.13); Delaware (0.09); Chesapeake Bay ( 0.78 ), where MD ( 0.67 ) and VA ( 0.33 ).

Table 23 cont'd. Constant M coastwide fishing mortality rates and total abundance estimates.

Unweighted average F estimate of coastal and producer area estimates and estimated total abundance of age $7+$ Atlantic striped bass based on $F$ estimates that assume constant natural mortality.

| Age 7+ CAA | Total number of age 7+, |
| :---: | :---: |
| Thousands | Thousands. |
| 45.5 | $\mathbf{- 3 8 3}$ |
| 101.4 | $\mathbf{- 1 1 3 3}$ |
| 95 | $\mathbf{- 6 8 0}$ |
| 222.3 | $\mathbf{1 , 1 6 3}$ |
| 296.4 | $\mathbf{2 , 6 3 6}$ |
| 262.7 | $\mathbf{2 , 8 4 3}$ |
| 380.6 | $\mathbf{3 , 6 2 6}$ |
| 475.9 | $\mathbf{4 , 7 5 8}$ |
| 740 | $\mathbf{4 , 7 1 6}$ |
| 965.3 | $\mathbf{4 , 9 0 8}$ |
| 1371.1 | $\mathbf{6 , 4 5 7}$ |
| 1080.5 | $\mathbf{5 , 0 0 0}$ |
| 1146.8 | $\mathbf{4 , 8 0 0}$ |
| 1471.8 | $\mathbf{8 , 6 5 2}$ |
| 1583.2 | $\mathbf{8 , 3 5 5}$ |
| 2075.4 | $\mathbf{1 0 , 5 3 2}$ |
| 2163.1 | $\mathbf{8 , 5 9 9}$ |
| 2376.2 | $\mathbf{8 , 1 8 7}$ |

Table 24. Estimates of fishing mortality for 28 inch striped bass based on Baranov's catch equation without assuming constant natural mortality, based on the exploitation rates (Table 20) and the bias-adjusted estimates of survival (Table 22). The tables also present annual estimates of instantaneous natural mortality, M. Column headings are S : annual bias-corrected survival rate, Z : total instantaneous mortality, A: annual percentage mortality expressed as a proportion, U : annual exploitation rate, F : instantaneous fishing mortality rate and M : instantaneous natural mortality rate.

## Producer areas

Maryland - Chesapeake Bay Spring Spawning Stock

| Year | $\underline{Z}$ | A | U | F | M | Year | $\underline{Z}$ | A | $\underline{\mathbf{U}}$ | F | M |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | -0.02 | -0.02 | 0.07 | 0.07 | -0.10 | 1988 |  |  |  |  |  |
| 1989 | -0.04 | -0.04 | 0.04 | 0.04 | -0.08 | 1989 |  |  |  |  |  |
| 1990 | 0.54 | 0.42 | 0.08 | 0.11 | 0.43 | 1990 | 0.34 | 0.28 | 0.26 | 0.31 | 0.03 |
| 1991 | 0.33 | 0.28 | 0.12 | 0.15 | 0.19 | 1991 | 0.33 | 0.28 | 0.36 | 0.43 | -0.10 |
| 1992 | 0.28 | 0.25 | 0.12 | 0.14 | 0.14 | 1992 | 0.28 | 0.25 | 0.37 | 0.42 | -0.14 |
| 1993 | 0.24 | 0.21 | 0.12 | 0.13 | 0.10 | 1993 | 0.37 | 0.31 | 0.37 | 0.44 | -0.07 |
| 1994 | 0.18 | 0.16 | 0.11 | 0.12 | 0.06 | 1994 | 0.40 | 0.33 | 0.25 | 0.31 | 0.09 |
| 1995 | 0.35 | 0.30 | 0.20 | 0.24 | 0.11 | 1995 | 0.41 | 0.34 | 0.41 | 0.50 | -0.09 |
| 1996 | 0.37 | 0.31 | 0.17 | 0.20 | 0.17 | 1996 | 0.48 | 0.38 | 0.18 | 0.22 | 0.26 |
| 1997 | 0.39 | 0.32 | 0.23 | 0.28 | 0.10 | 1997 | 0.46 | 0.37 | 0.38 | 0.47 | -0.01 |
| 1998 | 0.41 | 0.34 | 0.20 | 0.24 | 0.17 | 1998 | 0.41 | 0.33 | 0.45 | 0.55 | -0.15 |
| 1999 | 0.41 | 0.33 | 0.32 | 0.39 | 0.01 | 1999 | 0.44 | 0.35 | 0.28 | 0.35 | 0.09 |
| 2000 | 0.44 | 0.35 | 0.17 | 0.21 | 0.23 | 2000 | 0.39 | 0.32 | 0.25 | 0.31 | 0.08 |
| 2001 | 0.44 | 0.36 | 0.11 | 0.13 | 0.31 | 2001 | 0.41 | 0.33 | 0.23 | 0.27 | 0.13 |
| 2002 | 0.44 | 0.36 | 0.10 | 0.12 | 0.32 | 2002 | 0.38 | 0.32 | 0.29 | 0.35 | 0.03 |
| 2003 | 0.47 | 0.37 | 0.10 | 0.13 | 0.34 | 2003 | 0.40 | 0.33 | 0.22 | 0.26 | 0.14 |
| 2004 | 0.49 | 0.39 | 0.08 | 0.11 | 0.39 | 2004 | 0.40 | 0.33 | 0.08 | 0.09 | 0.31 |
| Average | 0.34 | 0.28 | 0.14 | 0.17 | 0.17 | Average | 0.39 | 0.32 | 0.29 | 0.35 | 0.04 |

Delaware River, Delaware, New Jersey, Pennsylvania Spring Spawning Stock

| Year | $\underline{Z}$ | A | $\underline{\mathbf{U}}$ | F | M | Year | $\underline{Z}$ | A | $\underline{\mathbf{U}}$ | F | M |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 |  |  |  |  |  | 1988 | 0.20 | 0.18 | 0.10 | 0.11 | 0.09 |
| 1989 |  |  |  |  |  | 1989 | 0.12 | 0.11 | 0.07 | 0.07 | 0.05 |
| 1990 |  |  |  |  |  | 1990 | 0.23 | 0.20 | 0.11 | 0.12 | 0.11 |
| 1991 |  |  |  |  |  | 1991 | 0.32 | 0.27 | 0.11 | 0.13 | 0.19 |
| 1992 |  |  |  |  |  | 1992 | 0.24 | 0.21 | 0.13 | 0.15 | 0.09 |
| 1993 | 0.08 | 0.11 | 0.17 | 0.18 | -0.06 | 1993 | 0.27 | 0.24 | 0.17 | 0.19 | 0.08 |
| 1994 | 0.10 | 0.12 | 0.12 | 0.12 | 0.01 | 1994 | 0.28 | 0.24 | 0.12 | 0.13 | 0.14 |
| 1995 | 0.42 | 0.34 | 0.14 | 0.17 | 0.25 | 1995 | 0.31 | 0.26 | 0.16 | 0.18 | 0.13 |
| 1996 | 0.41 | 0.33 | 0.31 | 0.38 | 0.02 | 1996 | 0.34 | 0.29 | 0.23 | 0.27 | 0.07 |
| 1997 | 0.45 | 0.36 | 0.27 | 0.33 | 0.12 | 1997 | 0.28 | 0.25 | 0.29 | 0.33 | -0.05 |
| 1998 | 0.46 | 0.37 | 0.28 | 0.34 | 0.12 | 1998 | 0.35 | 0.29 | 0.22 | 0.26 | 0.09 |
| 1999 | 0.49 | 0.39 | 0.15 | 0.19 | 0.30 | 1999 | 0.32 | 0.27 | 0.22 | 0.25 | 0.07 |
| 2000 | 0.46 | 0.37 | 0.30 | 0.37 | 0.09 | 2000 | 0.31 | 0.27 | 0.13 | 0.15 | 0.16 |
| 2001 | 0.49 | 0.39 | 0.26 | 0.33 | 0.16 | 2001 | 0.36 | 0.30 | 0.14 | 0.16 | 0.20 |
| 2002 | 0.48 | 0.38 | 0.24 | 0.30 | 0.18 | 2002 | 0.33 | 0.28 | 0.19 | 0.23 | 0.11 |
| 2003 | 0.44 | 0.36 | 0.16 | 0.20 | 0.25 | 2003 | 0.38 | 0.32 | 0.13 | 0.16 | 0.22 |
| 2004 | 0.46 | 0.37 | 0.21 | 0.26 | 0.20 | 2004 | 0.42 | 0.34 | 0.18 | 0.22 | 0.19 |
| Average | 0.40 | 0.33 | 0.22 | 0.26 | 0.14 | Average | 0.30 | 0.26 | 0.16 | 0.18 | 0.11 |

Table 24 cont'd.

## Coastal Programs

| Massachusetts Fall Tagging |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\frac{\text { Year }}{1988}$ | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| 1989 |  |  |  |  |  |
| 1990 |  |  |  |  |  |
| 1991 |  |  |  |  |  |
| 1992 | 0.17 | 0.16 | 0.04 | $\mathbf{0 . 0 4}$ | 0.14 |
| 1993 | 0.16 | 0.15 | 0.07 | $\mathbf{0 . 0 7}$ | 0.09 |
| 1994 | 0.18 | 0.17 | 0.03 | $\mathbf{0 . 0 3}$ | 0.15 |
| 1995 | 0.25 | 0.22 | 0.05 | $\mathbf{0 . 0 6}$ | 0.19 |
| 1996 | 0.25 | 0.22 | 0.08 | $\mathbf{0 . 1 0}$ | 0.16 |
| 1997 | 0.27 | 0.24 | 0.13 | $\mathbf{0 . 1 4}$ | 0.13 |
| 1998 | 0.25 | 0.22 | 0.07 | $\mathbf{0 . 0 8}$ | 0.17 |
| 1999 | 0.26 | 0.23 | 0.09 | $\mathbf{0 . 1 0}$ | 0.17 |
| 2000 | 0.26 | 0.23 | 0.13 | $\mathbf{0 . 1 5}$ | 0.12 |
| 2001 | 0.26 | 0.23 | 0.07 | $\mathbf{0 . 0 8}$ | 0.18 |
| 2002 | 0.24 | 0.21 | 0.08 | $\mathbf{0 . 0 9}$ | 0.15 |
| 2003 | 0.27 | 0.23 | 0.11 | $\mathbf{0 . 1 2}$ | 0.14 |
| 2004 | 0.25 | 0.22 | 0.09 | $\mathbf{0 . 1 0}$ | 0.14 |
|  |  |  |  |  |  |
| Average | 0.24 | 0.21 | 0.08 | $\mathbf{0 . 0 9}$ | 0.15 |

New Jersey Delaware Bay February-April

| $\underline{\text { Year }}$ | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 |  |  |  |  |  |
| 1989 | -0.14 | -0.16 | 0.02 | $\mathbf{0 . 0 2}$ | -0.16 |
| 1990 | 0.25 | 0.22 | 0.04 | $\mathbf{0 . 0 5}$ | 0.20 |
| 1991 | 0.11 | 0.10 | 0.18 | $\mathbf{0 . 1 9}$ | -0.08 |
| 1992 | 0.25 | 0.22 | 0.02 | $\mathbf{0 . 0 2}$ | 0.23 |
| 1993 | 0.25 | 0.22 | 0.09 | $\mathbf{0 . 1 0}$ | 0.15 |
| 1994 | 0.27 | 0.23 | 0.05 | $\mathbf{0 . 0 6}$ | 0.21 |
| 1995 | 0.15 | 0.14 | 0.10 | $\mathbf{0 . 1 1}$ | 0.04 |
| 1996 | 0.21 | 0.19 | 0.20 | $\mathbf{0 . 2 2}$ | -0.01 |
| 1997 | 0.32 | 0.27 | 0.23 | $\mathbf{0 . 2 7}$ | 0.05 |
| 1998 | 0.33 | 0.28 | 0.35 | $\mathbf{0 . 4 1}$ | -0.08 |
| 1999 | 0.46 | 0.37 | 0.08 | $\mathbf{0 . 1 0}$ | 0.36 |
| 2000 | 0.05 | 0.05 | 0.14 | $\mathbf{0 . 1 4}$ | -0.09 |
| 2001 | 0.15 | 0.14 | 0.14 | $\mathbf{0 . 1 5}$ | 0.00 |
| 2002 | 0.32 | 0.27 | 0.11 | $\mathbf{0 . 1 3}$ | 0.19 |
| 2003 | 0.55 | 0.42 | 0.14 | $\mathbf{0 . 1 8}$ | 0.37 |
| 2004 | 0.87 | 0.58 | 0.16 | $\mathbf{0 . 2 3}$ | 0.63 |
|  |  |  |  |  |  |
| Average | 0.27 | 0.22 | 0.13 | $\mathbf{0 . 1 5}$ | 0.13 |

New York Ocean Haul Seine Fall Tagging

| Year | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | -0.06 | 0.19 | 0.05 | $\mathbf{- 0 . 0 2}$ | -0.04 |
| 1989 | -0.01 | 0.19 | 0.04 | $\mathbf{0 . 0 0}$ | -0.01 |
| 1990 | 0.32 | 0.37 | 0.07 | $\mathbf{0 . 0 6}$ | 0.26 |
| 1991 | 0.31 | 0.37 | 0.12 | $\mathbf{0 . 1 0}$ | 0.21 |
| 1992 | 0.24 | 0.37 | 0.11 | $\mathbf{0 . 0 7}$ | 0.17 |
| 1993 | 0.34 | 0.37 | 0.14 | $\mathbf{0 . 1 2}$ | 0.22 |
| 1994 | 0.32 | 0.37 | 0.08 | $\mathbf{0 . 0 7}$ | 0.25 |
| 1995 | 0.28 | 0.35 | 0.20 | $\mathbf{0 . 1 6}$ | 0.12 |
| 1996 | 0.31 | 0.35 | 0.14 | $\mathbf{0 . 1 2}$ | 0.19 |
| 1997 | 0.34 | 0.35 | 0.29 | $\mathbf{0 . 2 8}$ | 0.05 |
| 1998 | 0.38 | 0.35 | 0.17 | $\mathbf{0 . 1 8}$ | 0.20 |
| 1999 | 0.39 | 0.35 | 0.29 | $\mathbf{0 . 3 3}$ | 0.06 |
| 2000 | 0.20 | 0.25 | 0.18 | $\mathbf{0 . 1 4}$ | 0.06 |
| 2001 | 0.22 | 0.25 | 0.11 | $\mathbf{0 . 0 9}$ | 0.12 |
| 2002 | 0.16 | 0.25 | 0.23 | $\mathbf{0 . 1 4}$ | 0.01 |
| 2003 | 0.24 | 0.24 | 0.15 | $\mathbf{0 . 1 4}$ | 0.09 |
| 2004 | 0.17 | 0.24 | 0.15 | $\mathbf{0 . 1 0}$ | 0.07 |
|  |  |  |  |  |  |
| Average | 0.24 | 0.31 | 0.15 | $\mathbf{0 . 1 2}$ | 0.12 |

North Carolina Co-op Winter Cruise

| Year | $\underline{Z}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.14 | 0.13 | 0.06 | $\mathbf{0 . 0 7}$ | 0.08 |
| 1989 | 0.27 | 0.24 | 0.05 | $\mathbf{0 . 0 5}$ | 0.22 |
| 1990 | 0.25 | 0.22 | 0.09 | $\mathbf{0 . 1 0}$ | 0.15 |
| 1991 | 0.21 | 0.19 | 0.07 | $\mathbf{0 . 0 8}$ | 0.13 |
| 1992 | 0.22 | 0.19 | 0.14 | $\mathbf{0 . 1 6}$ | 0.06 |
| 1993 | 0.23 | 0.20 | 0.11 | $\mathbf{0 . 1 3}$ | 0.10 |
| 1994 | 0.27 | 0.23 | 0.10 | $\mathbf{0 . 1 2}$ | 0.15 |
| 1995 | 0.30 | 0.26 | 0.14 | $\mathbf{0 . 1 7}$ | 0.14 |
| 1996 | 0.40 | 0.33 | 0.12 | $\mathbf{0 . 1 5}$ | 0.25 |
| 1997 | 0.39 | 0.32 | 0.21 | $\mathbf{0 . 2 5}$ | 0.14 |
| 1998 | 0.35 | 0.30 | 0.20 | $\mathbf{0 . 2 4}$ | 0.11 |
| 1999 | 0.36 | 0.30 | 0.24 | $\mathbf{0 . 2 8}$ | 0.08 |
| 2000 | 0.40 | 0.33 | 0.06 | $\mathbf{0 . 0 8}$ | 0.33 |
| 2001 | 0.39 | 0.32 | 0.15 | $\mathbf{0 . 1 8}$ | 0.21 |
| 2002 | 0.40 | 0.33 | 0.12 | $\mathbf{0 . 1 4}$ | 0.26 |
| 2003 | 0.41 | 0.34 | 0.12 | $\mathbf{0 . 1 5}$ | 0.26 |
| 2004 | 0.41 | 0.34 | 0.12 | $\mathbf{0 . 1 5}$ | 0.26 |
|  |  |  |  |  |  |
| Average | 0.32 | 0.27 | 0.12 | $\mathbf{0 . 1 5}$ | 0.17 |

Table 25. Summary table of F estimates based on Baranov's catch equation without assuming constant natural mortality of striped bass $\geq 28$ inches by individual program, averaged over all coastal programs, with a weighted average over producter areas and with an overall coastwide estimate. Estimates of coastwide abundance of age 7+ striped bass from 1987-2004 are also provided.

## Coast Programs

| Year | MADFW | NYOHS | NJDEL | NCCOOP | Unweighted <br> average |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 |  | -0.02 |  | 0.07 | $\mathbf{0 . 0 3}$ |
| 1989 |  | 0.00 | 0.02 | 0.05 | $\mathbf{0 . 0 2}$ |
| 1990 |  | 0.06 | 0.05 | 0.10 | $\mathbf{0 . 0 7}$ |
| 1991 |  | 0.10 | 0.19 | 0.08 | $\mathbf{0 . 1 2}$ |
| 1992 | 0.02 | 0.07 | 0.02 | 0.16 | $\mathbf{0 . 0 7}$ |
| 1993 | 0.01 | 0.12 | 0.10 | 0.13 | $\mathbf{0 . 0 9}$ |
| 1994 | 0.03 | 0.07 | 0.06 | 0.12 | $\mathbf{0 . 0 7}$ |
| 1995 | 0.10 | 0.16 | 0.11 | 0.17 | $\mathbf{0 . 1 3}$ |
| 1996 | 0.10 | 0.12 | 0.22 | 0.15 | $\mathbf{0 . 1 5}$ |
| 1997 | 0.12 | 0.28 | 0.27 | 0.25 | $\mathbf{0 . 2 3}$ |
| 1998 | 0.10 | 0.18 | 0.41 | 0.24 | $\mathbf{0 . 2 3}$ |
| 1999 | 0.11 | 0.33 | 0.10 | 0.28 | $\mathbf{0 . 2 1}$ |
| 2000 | 0.11 | 0.14 | 0.14 | 0.08 | $\mathbf{0 . 1 2}$ |
| 2001 | 0.11 | 0.09 | 0.15 | 0.18 | $\mathbf{0 . 1 3}$ |
| 2002 | 0.09 | 0.14 | 0.13 | 0.14 | $\mathbf{0 . 1 3}$ |
| 2003 | 0.12 | 0.14 | 0.18 | 0.15 | $\mathbf{0 . 1 5}$ |
| 2004 | 0.10 | 0.10 | 0.23 | 0.15 | $\mathbf{0 . 1 5}$ |

Producer Area Programs

| Year | HUDSON | DE/PA | MDCB | VARAP | Weighted <br> average* |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 |  |  |  |  |  |
| 1988 | 0.05 |  | 0.07 |  | $\mathbf{0 . 0 5}$ |
| 1989 | 0.05 |  | 0.04 |  | $\mathbf{0 . 0 3}$ |
| 1990 | 0.08 |  | 0.11 | 0.31 | $\mathbf{0 . 1 5}$ |
| 1991 | 0.10 |  | 0.15 | 0.43 | $\mathbf{0 . 2 0}$ |
| 1992 | 0.11 |  | 0.14 | 0.42 | $\mathbf{0 . 2 0}$ |
| 1993 | 0.13 | 0.18 | 0.13 | 0.44 | $\mathbf{0 . 2 2}$ |
| 1994 | 0.10 | 0.12 | 0.12 | 0.31 | $\mathbf{0 . 1 7}$ |
| 1995 | 0.15 | 0.17 | 0.24 | 0.50 | $\mathbf{0 . 2 9}$ |
| 1996 | 0.19 | 0.38 | 0.20 | 0.22 | $\mathbf{0 . 2 2}$ |
| 1997 | 0.25 | 0.33 | 0.28 | 0.47 | $\mathbf{0 . 3 3}$ |
| 1998 | 0.21 | 0.34 | 0.24 | 0.55 | $\mathbf{0 . 3 3}$ |
| 1999 | 0.17 | 0.19 | 0.39 | 0.35 | $\mathbf{0 . 3 3}$ |
| 2000 | 0.11 | 0.37 | 0.21 | 0.31 | $\mathbf{0 . 2 4}$ |
| 2001 | 0.12 | 0.33 | 0.13 | 0.27 | $\mathbf{0 . 1 9}$ |
| 2002 | 0.10 | 0.30 | 0.12 | 0.35 | $\mathbf{0 . 1 9}$ |
| 2003 | 0.12 | 0.20 | 0.13 | 0.26 | $\mathbf{0 . 1 7}$ |
| 2004 | 0.22 | 0.26 | 0.11 | 0.09 | $\mathbf{0 . 1 3}$ |

* Weighting Scheme: Hudson (0.13); Delaware (0.09); Chesapeake Bay (0.78), where MD (0.67) and VA (0.33).

Table 25 cont'd.

Coastwide Average Fishing Mortality Rate - Unweighted average of producer and coastal program mean fishing mortality rates obtained via the catch equation. Estimated total abundance of Atlantic coast striped bass ages 7+ obtained via the equation, Total Kill $=\mathrm{F} *$ Total Abundance, solving for total abundance.

| Year | Fishing Mortality | Age 7+CAA, <br> Thousands | Total Abundance ages 7+, <br> Thousands |
| :---: | :---: | :---: | :---: |
| 1988 | $\mathbf{0 . 0 4}$ | 101.4 | 2,849 |
| 1989 | $\mathbf{0 . 0 2}$ | 95 | 3,827 |
| 1990 | $\mathbf{0 . 1 1}$ | 222.3 | 2,075 |
| 1991 | $\mathbf{0 . 1 6}$ | 296.4 | 1,844 |
| 1992 | $\mathbf{0 . 1 3}$ | 262.7 | 1,994 |
| 1993 | $\mathbf{0 . 1 5}$ | 380.6 | 2,486 |
| 1994 | $\mathbf{0 . 1 2}$ | 475.9 | 4,027 |
| 1995 | $\mathbf{0 . 2 1}$ | 740 | 3,486 |
| 1996 | $\mathbf{0 . 1 9}$ | 965.3 | 5,201 |
| 1997 | $\mathbf{0 . 2 8}$ | 1371.1 | 4,893 |
| 1998 | $\mathbf{0 . 2 8}$ | 1080.5 | 3,877 |
| 1999 | $\mathbf{0 . 2 7}$ | 1146.8 | 4,256 |
| 2000 | $\mathbf{0 . 1 8}$ | 1471.8 | 8,280 |
| 2001 | $\mathbf{0 . 1 6}$ | 1583.2 | 9,907 |
| 2002 | $\mathbf{0 . 1 6}$ | 2075.4 | 13,066 |
| 2003 | $\mathbf{0 . 1 6}$ | 2163.1 | 13,672 |
| 2004 | $\mathbf{0 . 1 4}$ | 2376.2 | 17,099 |

Table 26. Survival (S) and fishing mortality (F) rates of striped bass $\geq 18$ inches assuming constant natural mortality, adjusted bias due to live release of recaptured fish. Diagnostic statistics c-hat and bootstrap goodness-of-fit are provided for each program.

## Producer Area Programs

Hudson River
C-hat adjustment $=1.236$; bootstrap GOF probability $=0.302$ for the full parameterized model.

| Year | S(unadj.) | F(unadj.) | Recovery Rate | \% Live <br> Release | Bias Live Release | S(adi | F(adj.) | $95 \% \mathrm{LCL}$ <br> F(adj) | $55 \% \mathrm{UCL}$ <br> F(adj) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.74 | 0.15 | 0.07 | 0.75 | -0.11 | 0.83 | 0.04 | -0.06 | 0.15 |
| 1989 | 0.72 | 0.18 | 0.09 | 0.79 | -0.16 | 0.85 | 0.01 | -0.08 | 0.17 |
| 1990 | 0.62 | 0.32 | 0.11 | 0.73 | -0.19 | 0.77 | 0.11 | 0.04 | 0.16 |
| 1991 | 0.64 | 0.30 | 0.10 | 0.62 | -0.15 | 0.75 | 0.14 | 0.08 | 0.18 |
| 1992 | 0.64 | 0.30 | 0.10 | 0.65 | -0.16 | 0.76 | 0.13 | 0.08 | 0.17 |
| 1993 | 0.64 | 0.29 | 0.11 | 0.57 | -0.14 | 0.75 | 0.14 | 0.10 | 0.19 |
| 1994 | 0.65 | 0.28 | 0.10 | 0.60 | -0.13 | 0.75 | 0.14 | 0.09 | 0.20 |
| 1995 | 0.67 | 0.25 | 0.09 | 0.44 | -0.10 | 0.75 | 0.14 | 0.09 | 0.21 |
| 1996 | 0.66 | 0.26 | 0.11 | 0.34 | -0.10 | 0.73 | 0.16 | 0.13 | 0.21 |
| 1997 | 0.65 | 0.27 | 0.13 | 0.38 | -0.12 | 0.75 | 0.14 | 0.09 | 0.21 |
| 1998 | 0.65 | 0.27 | 0.11 | 0.28 | -0.08 | 0.72 | 0.19 | 0.14 | 0.23 |
| 1999 | 0.65 | 0.27 | 0.10 | 0.35 | -0.09 | 0.72 | 0.18 | 0.12 | 0.22 |
| 2000 | 0.68 | 0.23 | 0.08 | 0.46 | -0.09 | 0.75 | 0.14 | 0.09 | 0.22 |
| 2001 | 0.67 | 0.25 | 0.07 | 0.37 | -0.06 | 0.71 | 0.19 | 0.13 | 0.25 |
| 2002 | 0.65 | 0.29 | 0.07 | 0.43 | -0.08 | 0.70 | 0.21 | 0.10 | 0.31 |
| 2003 | 0.63 | 0.30 | 0.09 | 0.46 | -0.10 | 0.70 | 0.20 | 0.08 | 0.28 |
| 2004 | 0.61 | 0.34 | 0.09 | 0.38 | -0.09 | 0.67 | 0.25 | -0.01 | 0.56 |

Delaware River; C-hat $=1.00$; bootstrap GOF probability $=0.76$ for the full parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live | 95\%LCL |  |  | 95\%UCL |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |  |
| 1993 | 0.64 | 0.30 | 0.099 | 0.39 | -0.0909 | 0.71 | $\mathbf{0 . 2 0}$ | 0.05 | 0.37 |  |
| 1994 | 0.63 | 0.31 | 0.106 | 0.55 | -0.142 | 0.73 | $\mathbf{0 . 1 6}$ | 0.03 | 0.31 |  |
| 1995 | 0.61 | 0.34 | 0.118 | 0.50 | -0.148 | 0.72 | $\mathbf{0 . 1 8}$ | 0.12 | 0.25 |  |
| 1996 | 0.61 | 0.35 | 0.121 | 0.44 | -0.137 | 0.70 | $\mathbf{0 . 2 0}$ | 0.14 | 0.27 |  |
| 1997 | 0.61 | 0.34 | 0.078 | 0.52 | -0.096 | 0.68 | $\mathbf{0 . 2 4}$ | 0.17 | 0.30 |  |
| 1998 | 0.62 | 0.36 | 0.104 | 0.48 | -0.12 | 0.69 | $\mathbf{0 . 2 3}$ | 0.16 | 0.30 |  |
| 1999 | 0.61 | 0.354 | 0.087 | 0.47 | -0.10 | 0.68 | $\mathbf{0 . 2 4}$ | 0.18 | 0.31 |  |
| 2000 | 0.60 | 0.37 | 0.098 | 0.46 | -0.11 | 0.67 | $\mathbf{0 . 2 5}$ | 0.18 | 0.33 |  |
| 2001 | 0.60 | 0.37 | 0.072 | 0.56 | -0.09 | 0.66 | $\mathbf{0 . 2 7}$ | 0.20 | 0.34 |  |
| 2002 | 0.60 | 0.36 | 0.080 | 0.35 | -0.07 | 0.65 | $\mathbf{0 . 2 9}$ | 0.22 | 0.37 |  |
| 2003 | 0.59 | 0.38 | 0.107 | 0.46 | -0.12 | 0.67 | $\mathbf{0 . 2 5}$ | 0.14 | 0.37 |  |
| 2004 | 0.60 | 0.36 | 0.077 | 0.38 | -0.02 | 0.65 | $\mathbf{0 . 2 9}$ | 0.20 | 0.39 |  |

Table 26, cont'd.
Maryland - Chesapeake Bay Spring Spawning Stock
C-hat adjustment $=1.157$; bootstrap GOF probability $=0.05$ for the fully parameterized model.

| Year | S(unadj.) | F(unadj.) | Recovery Rate | \% Live <br> Release | Bias Live Release | S(adj.) | F(adj.) | $\begin{gathered} 95 \% \mathrm{LCL} \\ \mathrm{~F}(\mathrm{adj}) \\ \hline \end{gathered}$ | $\begin{gathered} 95 \% \mathrm{UCL} \\ \mathrm{~F}(\mathrm{adj}) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 | 0.81 | 0.06 | 0.07 | 0.95 | -0.15 | 0.95 | -0.10 | 0.05 | -0.18 |
| 1988 | 0.84 | 0.02 | 0.04 | 0.84 | -0.08 | 0.91 | -0.06 | 0.00 | -0.10 |
| 1989 | 0.87 | -0.01 | 0.03 | 0.93 | -0.07 | 0.94 | -0.08 | 0.03 | -0.15 |
| 1990 | 0.64 | 0.30 | 0.06 | 0.58 | -0.07 | 0.69 | 0.22 | 0.29 | 0.16 |
| 1991 | 0.64 | 0.30 | 0.08 | 0.45 | -0.09 | 0.70 | 0.21 | 0.26 | 0.17 |
| 1992 | 0.63 | 0.31 | 0.11 | 0.43 | -0.12 | 0.71 | 0.19 | 0.22 | 0.15 |
| 1993 | 0.62 | 0.32 | 0.09 | 0.38 | -0.08 | 0.68 | 0.23 | 0.28 | 0.19 |
| 1994 | 0.62 | 0.33 | 0.10 | 0.43 | -0.11 | 0.69 | 0.22 | 0.29 | 0.15 |
| 1995 | 0.63 | 0.31 | 0.12 | 0.32 | -0.10 | 0.70 | 0.20 | 0.32 | 0.10 |
| 1996 | 0.61 | 0.35 | 0.11 | 0.35 | -0.10 | 0.67 | 0.25 | 0.31 | 0.18 |
| 1997 | 0.58 | 0.40 | 0.11 | 0.27 | -0.08 | 0.63 | 0.31 | 0.36 | 0.26 |
| 1998 | 0.55 | 0.46 | 0.11 | 0.25 | -0.07 | 0.59 | 0.38 | 0.48 | 0.29 |
| 1999 | 0.52 | 0.51 | 0.11 | 0.21 | -0.06 | 0.55 | 0.44 | 0.62 | 0.29 |
| 2000 | 0.48 | 0.59 | 0.10 | 0.36 | -0.09 | 0.52 | 0.50 | 0.69 | 0.34 |
| 2001 | 0.50 | 0.55 | 0.08 | 0.33 | -0.06 | 0.53 | 0.48 | 0.59 | 0.38 |
| 2002 | 0.53 | 0.49 | 0.07 | 0.32 | -0.06 | 0.56 | 0.43 | 0.59 | 0.29 |
| 2003 | 0.55 | 0.45 | 0.08 | 0.24 | -0.05 | 0.58 | 0.40 | 0.67 | 0.19 |
| 2004 | 0.58 | 0.40 | 0.07 | 0.25 | -0.04 | 0.60 | 0.36 | 0.75 | 0.11 |

Virginia - Rappahannock River
C-hat adjustment $=1.49$; bootstrap GOF probability $=0.092$ for the full parameterized model .

|  | Recovery <br> Year |  |  |  | S(unadj.) | F(unadj.) | Rate | Release | Bias Live |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Release | S(adj.) | F(adj.) | F(adj) | F(adj) |  |  |  |  |  |
| 1990 | 0.80 | 0.07 | 0.111 | 0.481 | -0.14 | 0.93 | $\mathbf{- 0 . 0 8}$ | -0.231 | 0.296 |
| 1991 | 0.30 | 1.06 | 0.063 | 0.524 | -0.08 | 0.32 | $\mathbf{0 . 9 8}$ | 0.560 | 1.489 |
| 1992 | 0.79 | 0.09 | 0.124 | 0.408 | -0.14 | 0.92 | $\mathbf{- 0 . 0 6}$ | -0.267 | 0.791 |
| 1993 | 0.60 | 0.36 | 0.088 | 0.456 | -0.11 | 0.67 | $\mathbf{0 . 2 5}$ | -0.048 | 0.799 |
| 1994 | 0.57 | 0.41 | 0.086 | 0.381 | -0.09 | 0.62 | $\mathbf{0 . 3 2}$ | 0.006 | 0.874 |
| 1995 | 0.69 | 0.22 | 0.077 | 0.262 | -0.05 | 0.73 | $\mathbf{0 . 1 7}$ | -0.080 | 0.716 |
| 1996 | 0.62 | 0.32 | 0.056 | 0.274 | -0.04 | 0.65 | $\mathbf{0 . 2 8}$ | 0.004 | 0.807 |
| 1997 | 0.56 | 0.43 | 0.068 | 0.330 | -0.06 | 0.59 | $\mathbf{0 . 3 7}$ | 0.091 | 0.809 |
| 1998 | 0.41 | 0.74 | 0.064 | 0.362 | -0.06 | 0.44 | $\mathbf{0 . 6 8}$ | 0.362 | 1.086 |
| 1999 | 0.37 | 0.84 | 0.078 | 0.286 | -0.06 | 0.40 | $\mathbf{0 . 7 8}$ | 0.472 | 1.154 |
| 2000 | 0.42 | 0.72 | 0.067 | 0.436 | -0.07 | 0.45 | $\mathbf{0 . 6 4}$ | 0.371 | 0.972 |
| 2001 | 0.46 | 0.63 | 0.072 | 0.367 | -0.07 | 0.49 | $\mathbf{0 . 5 5}$ | 0.211 | 1.044 |
| 2002 | 0.64 | 0.29 | 0.067 | 0.368 | -0.06 | 0.69 | $\mathbf{0 . 2 2}$ | -0.068 | 0.847 |
| 2003 | 0.72 | 0.18 | 0.068 | 0.271 | -0.05 | 0.76 | $\mathbf{0 . 1 3}$ | -0.116 | 0.797 |
| 2004 | 0.78 | 0.10 | 0.054 | 0.268 | -0.04 | 0.81 | $\mathbf{0 . 0 6}$ | -0.100 | 0.437 |

Table 26, cont'd.

## Coastal Programs

North Carolina - Cooperative Trawl Cruise
C-hat adjustment $=2.214$; bootstrap GOF probability $<0.001$ for the full parameterized model.

|  |  | Recovery |  |  | \% Live | Bias Live |  | 95\%LCL |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1988 | 0.907 | -0.259 | 0.095 | 0.875 | -0.186 | 1.115 | $\mathbf{- 0 . 2 6}$ | -0.281 | -0.230 |
| 1989 | 0.605 | 0.262 | 0.046 | 0.858 | -0.086 | 0.662 | $\mathbf{0 . 2 6}$ | 0.089 | 0.499 |
| 1990 | 0.564 | 0.311 | 0.070 | 0.665 | -0.106 | 0.631 | $\mathbf{0 . 3 1}$ | 0.130 | 0.548 |
| 1991 | 0.611 | 0.210 | 0.091 | 0.574 | -0.125 | 0.698 | $\mathbf{0 . 2 1}$ | 0.053 | 0.418 |
| 1992 | 0.770 | -0.017 | 0.107 | 0.453 | -0.121 | 0.875 | $\mathbf{- 0 . 0 2}$ | -0.186 | 0.373 |
| 1993 | 0.765 | 0.008 | 0.093 | 0.451 | -0.103 | 0.853 | $\mathbf{0 . 0 1}$ | -0.154 | 0.357 |
| 1994 | 0.560 | 0.316 | 0.083 | 0.544 | -0.107 | 0.627 | $\mathbf{0 . 3 2}$ | 0.039 | 0.748 |
| 1995 | 0.846 | -0.089 | 0.100 | 0.402 | -0.101 | 0.941 | $\mathbf{- 0 . 0 9}$ | -0.214 | 0.315 |
| 1996 | 0.575 | 0.369 | 0.057 | 0.254 | -0.035 | 0.595 | $\mathbf{0 . 3 7}$ | 0.177 | 0.628 |
| 1997 | 0.522 | 0.446 | 0.087 | 0.240 | -0.053 | 0.551 | $\mathbf{0 . 4 5}$ | 0.182 | 0.818 |
| 1998 | 0.621 | 0.241 | 0.110 | 0.283 | -0.082 | 0.676 | $\mathbf{0 . 2 4}$ | 0.023 | 0.582 |
| 1999 | 0.865 | -0.076 | 0.097 | 0.274 | -0.069 | 0.929 | $\mathbf{- 0 . 0 8}$ | -0.174 | 0.187 |
| 2000 | 0.343 | 0.855 | 0.066 | 0.413 | -0.064 | 0.366 | $\mathbf{0 . 8 5}$ | 0.630 | 1.110 |
| 2001 | 0.514 | 0.446 | 0.078 | 0.354 | -0.067 | 0.551 | $\mathbf{0 . 4 5}$ | 0.248 | 0.696 |
| 2002 | 0.592 | 0.315 | 0.074 | 0.317 | -0.057 | 0.628 | $\mathbf{0 . 3 2}$ | 0.108 | 0.612 |
| 2003 | 0.494 | 0.506 | 0.071 | 0.272 | -0.047 | 0.519 | $\mathbf{0 . 5 1}$ | 0.205 | 0.936 |
| 2004 | 0.685 | 0.177 | 0.074 | 0.280 | -0.051 | 0.721 | $\mathbf{0 . 1 8}$ | 0.001 | 0.461 |

New Jersey - Delaware Bay
C-hat adjustment $=1.00$; bootstrap GOF probability $=0.496$ for the fully parameterized model.

| Year | S(unadj.) | F(unadj.) | Recovery | \% Released | bias | S(adj.) | F(adj.) | LCLM (F) UCLM (F) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1989 | 0.88 | -0.02 | 0.11 | 0.92 | -0.22 | 1.12 | $\mathbf{- 0 . 2 7}$ | -0.24 | 0.64 |
| 1990 | 0.83 | 0.04 | 0.11 | 0.83 | -0.21 | 1.04 | $\mathbf{- 0 . 1 9}$ | -0.21 | 0.72 |
| 1991 | 0.58 | 0.39 | 0.08 | 0.77 | -0.14 | 0.68 | $\mathbf{0 . 2 4}$ | 0.17 | 0.56 |
| 1992 | 0.64 | 0.30 | 0.07 | 0.88 | -0.13 | 0.74 | $\mathbf{0 . 1 6}$ | 0.17 | 0.34 |
| 1993 | 0.55 | 0.45 | 0.08 | 0.84 | -0.15 | 0.64 | $\mathbf{0 . 2 9}$ | 0.32 | 0.43 |
| 1994 | 0.68 | 0.23 | 0.08 | 0.86 | -0.15 | 0.80 | $\mathbf{0 . 0 7}$ | 0.13 | 0.18 |
| 1995 | 0.78 | 0.10 | 0.09 | 0.66 | -0.13 | 0.90 | $\mathbf{- 0 . 0 5}$ | 0.02 | 0.07 |
| 1996 | 0.75 | 0.14 | 0.11 | 0.60 | -0.16 | 0.90 | $\mathbf{- 0 . 0 4}$ | -0.02 | 0.18 |
| 1997 | 0.53 | 0.48 | 0.09 | 0.50 | -0.11 | 0.59 | $\mathbf{0 . 3 7}$ | 0.33 | 0.60 |
| 1998 | 0.72 | 0.18 | 0.12 | 0.47 | -0.14 | 0.84 | $\mathbf{0 . 0 2}$ | 0.05 | 0.20 |
| 1999 | 0.65 | 0.27 | 0.08 | 0.50 | -0.09 | 0.72 | $\mathbf{0 . 1 8}$ | 0.22 | 0.31 |
| 2000 | 0.70 | 0.20 | 0.09 | 0.50 | -0.10 | 0.78 | $\mathbf{0 . 0 9}$ | 0.14 | 0.23 |
| 2001 | 0.78 | 0.10 | 0.09 | 0.46 | -0.10 | 0.87 | $\mathbf{- 0 . 0 1}$ | 0.02 | 0.19 |
| 2002 | 0.57 | 0.41 | 0.06 | 0.42 | -0.06 | 0.61 | $\mathbf{0 . 3 4}$ | 0.35 | 0.52 |
| 2003 | 0.48 | 0.59 | 0.09 | 0.48 | -0.10 | 0.53 | $\mathbf{0 . 4 8}$ | 0.46 | 0.68 |
| 2004 | 0.39 | 0.80 | 0.11 | 0.43 | -0.11 | 0.44 | $\mathbf{0 . 6 8}$ | 0.58 | 0.97 |

Table 26, cont'd.

Massachusetts fall tagging program
C-hat $=1.13$, bootstrap GOF probability $=0.62$

|  |  | Recovery |  |  |  | \% Live | Bias Live | 95\%LCL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | S(unadj.) | F(unadj.) | Rate | Release | Release | S(adj.) | F(adj.) | F(adj) | F(adj) |
| 1992 | 0.77 | 0.11 | 0.05 | 0.76 | -0.09 | 0.85 | $\mathbf{0 . 0 2}$ | -0.04 | 0.09 |
| 1993 | 0.77 | 0.11 | 0.06 | 0.59 | -0.08 | 0.84 | $\mathbf{0 . 0 3}$ | -0.03 | 0.10 |
| 1994 | 0.77 | 0.11 | 0.05 | 0.58 | -0.07 | 0.83 | $\mathbf{0 . 0 4}$ | -0.02 | 0.11 |
| 1995 | 0.75 | 0.14 | 0.06 | 0.47 | -0.06 | 0.80 | $\mathbf{0 . 0 8}$ | 0.03 | 0.12 |
| 1996 | 0.75 | 0.14 | 0.09 | 0.43 | -0.09 | 0.83 | $\mathbf{0 . 0 4}$ | 0.00 | 0.09 |
| 1997 | 0.75 | 0.14 | 0.06 | 0.28 | -0.04 | 0.78 | $\mathbf{0 . 1 0}$ | 0.06 | 0.15 |
| 1998 | 0.75 | 0.14 | 0.08 | 0.33 | -0.07 | 0.80 | $\mathbf{0 . 0 7}$ | 0.03 | 0.12 |
| 1999 | 0.75 | 0.14 | 0.06 | 0.32 | -0.04 | 0.78 | $\mathbf{0 . 1 0}$ | 0.06 | 0.15 |
| 2000 | 0.76 | 0.13 | 0.05 | 0.24 | -0.03 | 0.78 | $\mathbf{0 . 1 0}$ | 0.05 | 0.16 |
| 2001 | 0.76 | 0.13 | 0.04 | 0.35 | -0.03 | 0.78 | $\mathbf{0 . 0 9}$ | 0.04 | 0.16 |
| 2002 | 0.76 | 0.13 | 0.06 | 0.29 | -0.04 | 0.79 | $\mathbf{0 . 0 8}$ | 0.03 | 0.15 |
| 2003 | 0.76 | 0.13 | 0.04 | 0.23 | -0.02 | 0.78 | $\mathbf{0 . 1 0}$ | 0.04 | 0.18 |
| 2004 | 0.76 | 0.12 | 0.05 | 0.22 | -0.02 | 0.78 | $\mathbf{0 . 1 0}$ | 0.03 | 0.18 |

New York Ocean Haul Seine
C-hat adjustment $=1.82$; bootstrap GOF probability $=0$ for the fully parameterized model.

| Year | S(unadj.) | F(unadj.) | Recovery | \% Released | bias | S(adj.) | F(adj.) | LCLM (F) | UCLM (F) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.550 | 0.45 | 0.077 | 0.94 | -0.16 | 0.653 | $\mathbf{0 . 2 8}$ | 0.123 | 0.464 |
| 1989 | 0.908 | -0.05 | 0.092 | 0.93 | -0.19 | 1.120 | $\mathbf{- 0 . 2 6}$ | -0.279 | -0.245 |
| 1990 | 0.551 | 0.45 | 0.073 | 0.82 | -0.13 | 0.636 | $\mathbf{0 . 3 0}$ | 0.137 | 0.510 |
| 1991 | 0.756 | 0.13 | 0.080 | 0.69 | -0.13 | 0.866 | $\mathbf{- 0 . 0 1}$ | -0.149 | 0.250 |
| 1992 | 0.932 | -0.08 | 0.069 | 0.72 | -0.11 | 1.050 | $\mathbf{- 0 . 2 0}$ | -0.212 | -0.183 |
| 1993 | 0.492 | 0.56 | 0.055 | 0.62 | -0.08 | 0.533 | $\mathbf{0 . 4 8}$ | 0.311 | 0.680 |
| 1994 | 0.680 | 0.24 | 0.062 | 0.71 | -0.10 | 0.755 | $\mathbf{0 . 1 3}$ | -0.008 | 0.329 |
| 1995 | 0.938 | -0.09 | 0.063 | 0.55 | -0.08 | 1.020 | $\mathbf{- 0 . 1 7}$ | -0.180 | -0.156 |
| 1996 | 0.792 | 0.08 | 0.058 | 0.61 | -0.08 | 0.861 | $\mathbf{0 . 0 0}$ | -0.142 | 0.307 |
| 1997 | 0.600 | 0.36 | 0.051 | 0.56 | -0.07 | 0.642 | $\mathbf{0 . 2 9}$ | 0.087 | 0.594 |
| 1998 | 0.489 | 0.57 | 0.054 | 0.57 | -0.07 | 0.526 | $\mathbf{0 . 4 9}$ | 0.267 | 0.782 |
| 1999 | 0.676 | 0.24 | 0.056 | 0.49 | -0.06 | 0.722 | $\mathbf{0 . 1 8}$ | -0.027 | 0.529 |
| 2000 | 0.588 | 0.38 | 0.047 | 0.59 | -0.06 | 0.626 | $\mathbf{0 . 3 2}$ | 0.091 | 0.653 |
| 2001 | 0.593 | 0.37 | 0.053 | 0.51 | -0.06 | 0.632 | $\mathbf{0 . 3 1}$ | 0.072 | 0.672 |
| 2002 | 0.890 | -0.03 | 0.065 | 0.52 | -0.08 | 0.965 | $\mathbf{- 0 . 1 1}$ | -0.227 | 1.326 |
| 2003 | 0.428 | 0.70 | 0.044 | 0.43 | -0.04 | 0.447 | $\mathbf{0 . 6 5}$ | 0.248 | 1.243 |
| 2004 | 0.595 | 0.37 | 0.061 | 0.48 | -0.07 | 0.639 | $\mathbf{0 . 3 0}$ | 0.147 | 0.491 |

Table 27. Akaike weights based on the statistical fit of the tag-recapture data to the various models in the suite. These weights are used to obtain the weighted model-averaged estimates of survival. Results are for striped bass tagged at $\geq 18$ inches.

Producer Area Programs

| Model | HUDSON | DE/PA | MDCB | VARAP |
| :--- | :---: | :---: | ---: | ---: |
|  |  |  |  |  |
| $\{\mathrm{S}() .\mathrm{r}()\}$. | 0.0000 | 0 | 0 | 0 |
| $\{\mathrm{~S}() .\mathrm{r}(\mathrm{p})\}$ | 0.0323 | 0.000 | 0 | 0.000 |
| $\{\mathrm{~S}() .\mathrm{r}(\mathrm{t})\}$ | 0.0028 | $\mathbf{0 . 5 8 1}$ | 0 | 0 |
| $\{\mathrm{~S}(\mathrm{p}) \mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 1 6 7 1}$ | 0.000 | 0 | 0.000 |
| $\{\mathrm{~S}(\mathrm{p}) \mathrm{r}(\mathrm{t})\}$ | 0.0013 | $\mathbf{0 . 2 7 3}$ | $\mathbf{0 . 1 9 6 6 1}$ | 0 |
| $\{\mathrm{~S}(\mathrm{~d}) \mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 1 3 6 4}$ | 0.000 | 0 | 0.000 |
| $\{\mathrm{~S}(\mathrm{v}) \mathrm{r}(\mathrm{p})\}$ | 0.0978 | 0.002 | 0 | 0.000 |
| $\mathrm{~S}(\mathrm{Va}) \mathrm{r}(\mathrm{va})$ | NA | NA |  | 0 |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{t})\}$ | 0.0086 | 0.059 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp})\}$ | 0.0891 | 0.000 | 0.7902 | 0.064 |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 3 0 5 4}$ | 0.000 | 0 | 0.005 |
| $\{\mathrm{~S}(\mathrm{t}) \mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 1 5 8 2}$ | 0.083 | 0.00034 | 0.000 |
| $\{\mathrm{~S}(\mathrm{t}) \mathrm{r}(\mathrm{t})\}$ | 0.0011 | 0.001 | 0.02403 | $\mathbf{0 . 9 3 0}$ |

Coastal Programs

| Model | MA FALL |  | NY LI OHS | NJ DB FEB-APR | NC COOP |
| :--- | ---: | :---: | :---: | :---: | :---: |
| $\{\mathrm{S}() .\mathrm{r}()\}$. | 0 | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}() .\mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 5 5 8}$ | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}() .\mathrm{r}(\mathrm{t})\}$ | 0.053 | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{p}) \mathrm{r}(\mathrm{p})\}$ | $\mathbf{0 . 1 7 4}$ | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{p}) \mathrm{r}(\mathrm{t})\}$ | 0.026 | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{~d}) \mathrm{r}(\mathrm{p})\}$ | 0.080 | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{v}) \mathrm{r}(\mathrm{p})\}$ | 0.077 | 0.00 | 0.000 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{t})\}$ | 0.003 | $* * *$ | 0.0469 | 0.014 |  |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{Tp})\}$ | 0.002 | $* * *$ | 0.00 | $\mathbf{0 . 1 5 7}$ |  |
| $\{\mathrm{~S}(\mathrm{Tp}) \mathrm{r}(\mathrm{p})\}$ | 0.014 | 0.00 | 0.00 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{t}) \mathrm{r}(\mathrm{p})\}$ | 0.013 | $* * *$ | 0.00 | 0.000 |  |
| $\{\mathrm{~S}(\mathrm{t}) \mathrm{r}(\mathrm{t})\}$ | 0.000 | $\mathbf{1 . 0 0}$ | $\mathbf{0 . 9 5}$ | $\mathbf{0 . 8 2 9}$ |  |

Table 28. Summary table, based on the assumption of constant natural mortality, with estimates of annual instantaneous fishing mortality, F , of striped bass $\geq 18$ inches by individual program, averaged over all coastal programs, with a weighted average over producer areas and with an overall coastwide estimate. Estimates of coastwide abundance of age 3+ striped bass from 19872004 are also provided.

Producer Area Programs*
Weighted

| Year | HUDSON | DE/PA | MDCB | VARAP | Weighted <br> Average |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 |  |  | -0.10 |  | $\mathbf{- 0 . 0 5}$ |
| 1988 | 0.04 |  | -0.06 |  | $\mathbf{- 0 . 0 2}$ |
| 1989 | 0.01 |  | -0.08 |  | $\mathbf{- 0 . 0 4}$ |
| 1990 | 0.11 |  | 0.22 | -0.08 | $\mathbf{0 . 1 1}$ |
| 1991 | 0.14 |  | 0.21 | 0.98 | $\mathbf{0 . 3 8}$ |
| 1992 | 0.13 |  | 0.19 | -0.06 | $\mathbf{0 . 1 0}$ |
| 1993 | 0.14 | 0.20 | 0.23 | 0.25 | $\mathbf{0 . 2 2}$ |
| 1994 | 0.14 | 0.16 | 0.22 | 0.32 | $\mathbf{0 . 2 3}$ |
| 1995 | 0.14 | 0.18 | 0.20 | 0.17 | $\mathbf{0 . 1 8}$ |
| 1996 | 0.16 | 0.20 | 0.25 | 0.28 | $\mathbf{0 . 2 4}$ |
| 1997 | 0.14 | 0.24 | 0.31 | 0.37 | $\mathbf{0 . 3 0}$ |
| 1998 | 0.19 | 0.23 | 0.38 | 0.68 | $\mathbf{0 . 4 2}$ |
| 1999 | 0.18 | 0.24 | 0.44 | 0.78 | $\mathbf{0 . 4 8}$ |
| 2000 | 0.14 | 0.25 | 0.50 | 0.64 | $\mathbf{0 . 4 7}$ |
| 2001 | 0.19 | 0.27 | 0.48 | 0.55 | $\mathbf{0 . 4 4}$ |
| 2002 | 0.21 | 0.29 | 0.43 | 0.22 | $\mathbf{0 . 3 4}$ |
| 2003 | 0.20 | 0.25 | 0.40 | 0.13 | $\mathbf{0 . 2 9}$ |
| 2004 | 0.25 | 0.29 | 0.36 | 0.06 | $\mathbf{0 . 2 6}$ |

Weighting Scheme: Hudson (0.13); Delaware (0.09);
Chesapeake Bay (0.78), where MD (0.67) and VA (0.33).

Coast Programs

| Year | MADFW | NYOHS | NJDEL | NCCOOP | Unweighted <br> average |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 |  |  |  |  |  |
| 1988 |  | 0.28 |  | -0.26 | $\mathbf{0 . 0 1}$ |
| 1989 |  | -0.26 | -0.27 | 0.26 | $\mathbf{- 0 . 0 9}$ |
| 1990 |  | 0.30 | -0.19 | 0.31 | $\mathbf{0 . 1 4}$ |
| 1991 |  | -0.01 | 0.24 | 0.21 | $\mathbf{0 . 1 5}$ |
| 1992 | 0.02 | -0.20 | 0.16 | -0.02 | $\mathbf{- 0 . 0 1}$ |
| 1993 | 0.03 | 0.48 | 0.29 | 0.01 | $\mathbf{0 . 2 0}$ |
| 1994 | 0.04 | 0.13 | 0.07 | 0.32 | $\mathbf{0 . 1 4}$ |
| 1995 | 0.08 | -0.17 | -0.05 | -0.09 | $\mathbf{- 0 . 0 6}$ |
| 1996 | 0.04 | 0.00 | -0.04 | 0.37 | $\mathbf{0 . 0 9}$ |
| 1997 | 0.10 | 0.29 | 0.37 | 0.45 | $\mathbf{0 . 3 0}$ |
| 1998 | 0.07 | 0.49 | 0.02 | 0.24 | $\mathbf{0 . 2 1}$ |
| 1999 | 0.10 | 0.18 | 0.18 | -0.08 | $\mathbf{0 . 0 9}$ |
| 2000 | 0.10 | 0.32 | 0.09 | 0.85 | $\mathbf{0 . 3 4}$ |
| 2001 | 0.09 | 0.31 | -0.01 | 0.45 | $\mathbf{0 . 2 1}$ |
| 2002 | 0.08 | -0.11 | 0.34 | 0.32 | $\mathbf{0 . 1 6}$ |


| 2003 | 0.10 | 0.65 | 0.48 | 0.51 | $\mathbf{0 . 4 4}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2004 | 0.10 | 0.30 | 0.68 | 0.18 | $\mathbf{0 . 3 1}$ |

Table 28 cont'd.

Unweighted average F estimate of coastal and producer area estimates and estimated total abundance of age 3+ Atlantic striped bass based on F estimates that assume constant natural mortality.

| Year | Fishing Mortality | Total Kill <br> inc. discards | Total Stock Size (3 + yrs. old) <br> Thousands |
| ---: | :---: | :---: | :---: |
| 1987 | $\mathbf{0 . 0 5}$ |  |  |
| 1988 | $\mathbf{- 0 . 0 1}$ |  |  |
| 1989 | $\mathbf{- 0 . 0 7}$ | 921 | 7,358 |
| 1990 | $\mathbf{0 . 1 3}$ | 988 | 3,747 |
| 1991 | $\mathbf{0 . 2 6}$ | 987 | 22,865 |
| 1992 | $\mathbf{0 . 0 4}$ | 1,437 | 6,780 |
| 1993 | $\mathbf{0 . 2 1}$ | 1,867 | 10,161 |
| 1994 | $\mathbf{0 . 1 8}$ | 3,000 | 47,743 |
| 1995 | $\mathbf{0 . 0 6}$ | 3,376 | 20,346 |
| 1996 | $\mathbf{0 . 1 7}$ | 4,580 | 15,251 |
| 1997 | $\mathbf{0 . 3 0}$ | 4,118 | 13,188 |
| 1998 | $\mathbf{0 . 3 1}$ | 3,704 | 12,981 |
| 1999 | $\mathbf{0 . 2 9}$ | 5,044 | 12,466 |
| 2000 | $\mathbf{0 . 4 0}$ | 4,344 | 13,317 |
| 2001 | $\mathbf{0 . 3 3}$ | 3,890 | 15,793 |
| 2002 | $\mathbf{0 . 2 5}$ | 4,836 | 13,339 |
| 2003 | $\mathbf{0 . 3 6}$ | 5,185 | 18,037 |

Table 29. Estimates of fishing mortality for 18 inch plus striped bass based on Baranov's catch equation without assuming constant natural mortality, based on the exploitation rates (Table 20) and the bias-adjusted estimates of survival (Table 22). The tables also present annual estimates of instantaneous natural mortality, M. Column headings are S : annual bias-corrected survival rate, Z: total instantaneous mortality, A: annual percentage mortality expressed as a proportion, U : annual exploitation rate, F : instantaneous fishing mortality rate and M : instantaneous natural mortality rate.

Producer Areas
Maryland Chesapeake Bay Spring Spawning Stock Virginia Rappahanock River Spring Spawning Stock Survey

| Year | $\underline{z}$ | A | $\underline{\text { u }}$ | F | M | Year | $\underline{Z}$ | A | $\underline{\mathbf{u}}$ | F | M |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 | 0.05 | 0.05 | 0.01 | 0.01 | 0.05 |  |  |  |  |  |  |
| 1988 | 0.09 | 0.09 | 0.01 | 0.01 | 0.08 | 1988 |  |  |  |  |  |
| 1989 | 0.07 | 0.06 | 0.01 | 0.01 | 0.06 | 1989 |  |  |  |  |  |
| 1990 | 0.37 | 0.31 | 0.07 | 0.08 | 0.29 | 1990 | 0.07 | 0.07 | 0.17 | 0.18 | -0.11 |
| 1991 | 0.36 | 0.30 | 0.10 | 0.12 | 0.24 | 1991 | 1.13 | 0.68 | 0.14 | 0.23 | 0.90 |
| 1992 | 0.34 | 0.29 | 0.13 | 0.15 | 0.18 | 1992 | 0.09 | 0.08 | 0.31 | 0.32 | -0.23 |
| 1993 | 0.38 | 0.32 | 0.11 | 0.13 | 0.25 | 1993 | 0.40 | 0.33 | 0.23 | 0.28 | 0.12 |
| 1994 | 0.37 | 0.31 | 0.12 | 0.14 | 0.23 | 1994 | 0.47 | 0.38 | 0.25 | 0.31 | 0.16 |
| 1995 | 0.35 | 0.30 | 0.18 | 0.22 | 0.13 | 1995 | 0.32 | 0.27 | 0.19 | 0.22 | 0.09 |
| 1996 | 0.40 | 0.33 | 0.17 | 0.20 | 0.20 | 1996 | 0.43 | 0.35 | 0.14 | 0.17 | 0.26 |
| 1997 | 0.46 | 0.37 | 0.20 | 0.25 | 0.21 | 1997 | 0.52 | 0.41 | 0.20 | 0.25 | 0.27 |
| 1998 | 0.53 | 0.41 | 0.19 | 0.25 | 0.28 | 1998 | 0.83 | 0.56 | 0.15 | 0.23 | 0.60 |
| 1999 | 0.59 | 0.45 | 0.16 | 0.22 | 0.38 | 1999 | 0.93 | 0.60 | 0.13 | 0.20 | 0.73 |
| 2000 | 0.65 | 0.48 | 0.13 | 0.18 | 0.47 | 2000 | 0.79 | 0.55 | 0.13 | 0.18 | 0.60 |
| 2001 | 0.63 | 0.47 | 0.12 | 0.16 | 0.47 | 2001 | 0.70 | 0.51 | 0.18 | 0.25 | 0.46 |
| 2002 | 0.58 | 0.44 | 0.12 | 0.15 | 0.43 | 2002 | 0.37 | 0.31 | 0.16 | 0.19 | 0.19 |
| 2003 | 0.55 | 0.42 | 0.13 | 0.17 | 0.38 | 2003 | 0.28 | 0.24 | 0.15 | 0.17 | 0.11 |
| 2004 | 0.51 | 0.40 | 0.10 | 0.13 | 0.38 | 2004 | 0.21 | 0.19 | 0.07 | 0.08 | 0.13 |
| Average | 0.40 | 0.32 | 0.11 | 0.14 | 0.28 | Average | 0.50 | 0.37 | 0.17 | 0.22 | 0.29 |

Delaware River: Delaware, Pennsylvania, New Jersey. Spring Spawning Stock

| Year | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ | $\underline{\text { Year }}$ | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 |  |  |  |  |  | 1988 | 0.19 | 0.17 | 0.05 | $\mathbf{0 . 0 5}$ | 0.13 |
| 1989 |  |  |  |  |  | 1989 | 0.16 | 0.15 | 0.05 | $\mathbf{0 . 0 5}$ | 0.11 |
| 1990 |  |  |  |  |  | 1990 | 0.26 | 0.23 | 0.07 | $\mathbf{0 . 0 8}$ | 0.17 |
| 1991 |  |  |  |  |  | 1991 | 0.29 | 0.25 | 0.08 | $\mathbf{0 . 1 0}$ | 0.19 |
| 1992 |  |  |  |  |  | 1992 | 0.27 | 0.24 | 0.10 | $\mathbf{0 . 1 1}$ | 0.16 |
| 1993 | 0.35 | 0.29 | 0.12 | $\mathbf{0 . 1 5}$ | 0.20 | 1993 | 0.29 | 0.25 | 0.11 | $\mathbf{0 . 1 3}$ | 0.16 |
| 1994 | 0.31 | 0.27 | 0.12 | $\mathbf{0 . 1 4}$ | 0.17 | 1994 | 0.29 | 0.25 | 0.08 | $\mathbf{0 . 1 0}$ | 0.19 |
| 1995 | 0.33 | 0.28 | 0.13 | $\mathbf{0 . 1 6}$ | 0.17 | 1995 | 0.29 | 0.25 | 0.13 | $\mathbf{0 . 1 5}$ | 0.14 |
| 1996 | 0.36 | 0.30 | 0.16 | $\mathbf{0 . 1 9}$ | 0.17 | 1996 | 0.31 | 0.27 | 0.16 | $\mathbf{0 . 1 9}$ | 0.12 |
| 1997 | 0.39 | 0.32 | 0.09 | $\mathbf{0 . 1 1}$ | 0.27 | 1997 | 0.29 | 0.25 | 0.22 | $\mathbf{0 . 2 5}$ | 0.04 |
| 1998 | 0.37 | 0.31 | 0.14 | $\mathbf{0 . 1 7}$ | 0.20 | 1998 | 0.33 | 0.28 | 0.17 | $\mathbf{0 . 2 0}$ | 0.13 |
| 1999 | 0.39 | 0.32 | 0.12 | $\mathbf{0 . 1 4}$ | 0.25 | 1999 | 0.33 | 0.28 | 0.14 | $\mathbf{0 . 1 7}$ | 0.16 |
| 2000 | 0.40 | 0.33 | 0.14 | $\mathbf{0 . 1 7}$ | 0.23 | 2000 | 0.29 | 0.25 | 0.09 | $\mathbf{0 . 1 1}$ | 0.18 |
| 2001 | 0.42 | 0.34 | 0.14 | $\mathbf{0 . 1 7}$ | 0.25 | 2001 | 0.34 | 0.29 | 0.10 | $\mathbf{0 . 1 2}$ | 0.22 |
| 2002 | 0.43 | 0.35 | 0.14 | $\mathbf{0 . 1 8}$ | 0.25 | 2002 | 0.36 | 0.30 | 0.08 | $\mathbf{0 . 1 0}$ | 0.26 |
| 2003 | 0.40 | 0.33 | 0.13 | $\mathbf{0 . 1 6}$ | 0.24 | 2003 | 0.35 | 0.30 | 0.10 | $\mathbf{0 . 1 2}$ | 0.24 |
| 2004 | 0.43 | 0.35 | 0.12 | $\mathbf{0 . 1 5}$ | 0.28 | 2004 | 0.40 | 0.33 | 0.04 | $\mathbf{0 . 0 5}$ | 0.35 |
| Average | $\mathbf{0 . 3 8}$ | $\mathbf{0 . 3 2}$ | $\mathbf{0 . 1 3}$ | $\mathbf{0 . 1 6}$ | $\mathbf{0 . 2 2}$ |  | Average | $\mathbf{0 . 3 0}$ | $\mathbf{0 . 2 6}$ | $\mathbf{0 . 1 5}$ | $\mathbf{0 . 1 7}$ |
| $\mathbf{0 . 1 2}$ |  |  |  |  |  |  |  |  |  |  |  |

Table 29 cont'd.

## Coastal Areas

Massachusetts Fall Tagging

| Year | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1992 | 0.17 | 0.15 | 0.04 | $\mathbf{0 . 0 4}$ | 0.13 |
| 1993 | 0.18 | 0.16 | 0.05 | $\mathbf{0 . 0 6}$ | 0.12 |
| 1994 | 0.19 | 0.17 | 0.03 | $\mathbf{0 . 0 4}$ | 0.15 |
| 1995 | 0.23 | 0.20 | 0.04 | $\mathbf{0 . 0 4}$ | 0.18 |
| 1996 | 0.19 | 0.17 | 0.06 | $\mathbf{0 . 0 7}$ | 0.12 |
| 1997 | 0.25 | 0.22 | 0.09 | $\mathbf{0 . 1 0}$ | 0.15 |
| 1998 | 0.22 | 0.20 | 0.08 | $\mathbf{0 . 0 9}$ | 0.13 |
| 1999 | 0.25 | 0.22 | 0.06 | $\mathbf{0 . 0 7}$ | 0.18 |
| 2000 | 0.25 | 0.22 | 0.08 | $\mathbf{0 . 1 0}$ | 0.15 |
| 2001 | 0.24 | 0.22 | 0.05 | $\mathbf{0 . 0 6}$ | 0.19 |
| 2002 | 0.23 | 0.21 | 0.09 | $\mathbf{0 . 1 0}$ | 0.13 |
| 2003 | 0.25 | 0.22 | 0.08 | $\mathbf{0 . 0 9}$ | 0.16 |
| 2004 | 0.25 | 0.22 | 0.09 | $\mathbf{0 . 1 0}$ | 0.15 |
|  |  |  |  |  |  |
| Average | 0.22 | 0.20 | 0.06 | $\mathbf{0 . 0 7}$ | 0.15 |

North Carolina Co-operative Winter Cruise

| Year | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\mathbf{F}$ | $\underline{\mathbf{M}}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| 1988 | -0.11 | -0.11 | 0.04 | $\mathbf{0 . 0 4}$ | -0.15 |
| 1989 | 0.41 | 0.34 | 0.03 | $\mathbf{0 . 0 4}$ | 0.37 |
| 1990 | 0.46 | 0.37 | 0.07 | $\mathbf{0 . 0 9}$ | 0.37 |
| 1991 | 0.36 | 0.30 | 0.08 | $\mathbf{0 . 1 0}$ | 0.26 |
| 1992 | 0.13 | 0.12 | 0.15 | $\mathbf{0 . 1 6}$ | -0.03 |
| 1993 | 0.16 | 0.15 | 0.11 | $\mathbf{0 . 1 2}$ | 0.04 |
| 1994 | 0.47 | 0.37 | 0.10 | $\mathbf{0 . 1 2}$ | 0.35 |
| 1995 | 0.06 | 0.06 | 0.14 | $\mathbf{0 . 1 5}$ | -0.09 |
| 1996 | 0.52 | 0.40 | 0.12 | $\mathbf{0 . 1 5}$ | 0.37 |
| 1997 | 0.60 | 0.45 | 0.17 | $\mathbf{0 . 2 3}$ | 0.37 |
| 1998 | 0.39 | 0.32 | 0.17 | $\mathbf{0 . 2 0}$ | 0.19 |
| 1999 | 0.07 | 0.07 | 0.22 | $\mathbf{0 . 2 3}$ | -0.15 |
| 2000 | 1.00 | 0.63 | 0.10 | $\mathbf{0 . 1 6}$ | 0.85 |
| 2001 | 0.60 | 0.45 | 0.12 | $\mathbf{0 . 1 5}$ | 0.44 |
| 2002 | 0.47 | 0.37 | 0.12 | $\mathbf{0 . 1 5}$ | 0.31 |
| 2003 | 0.66 | 0.48 | 0.11 | $\mathbf{0 . 1 6}$ | 0.50 |
| 2004 | 0.33 | 0.28 | 0.12 | $\mathbf{0 . 1 4}$ | 0.19 |

New York Ocean Haul Seine Fall Tagging

| Year | $\underline{\mathbf{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.43 | 0.35 | 0.02 | $\mathbf{0 . 0 3}$ | 0.40 |
| 1989 | -0.11 | -0.12 | 0.03 | $\mathbf{0 . 0 3}$ | -0.14 |
| 1990 | 0.45 | 0.36 | 0.04 | $\mathbf{0 . 0 5}$ | 0.40 |
| 1991 | 0.14 | 0.13 | 0.06 | $\mathbf{0 . 0 6}$ | 0.08 |
| 1992 | -0.05 | -0.05 | 0.04 | $\mathbf{0 . 0 4}$ | -0.09 |
| 1993 | 0.63 | 0.47 | 0.05 | $\mathbf{0 . 0 6}$ | 0.56 |
| 1994 | 0.28 | 0.25 | 0.04 | $\mathbf{0 . 0 4}$ | 0.24 |
| 1995 | -0.02 | -0.02 | 0.05 | $\mathbf{0 . 0 5}$ | -0.07 |
| 1996 | 0.15 | 0.14 | 0.03 | $\mathbf{0 . 0 3}$ | 0.12 |
| 1997 | 0.44 | 0.36 | 0.04 | $\mathbf{0 . 0 4}$ | 0.40 |
| 1998 | 0.64 | 0.47 | 0.03 | $\mathbf{0 . 0 4}$ | 0.60 |
| 1999 | 0.33 | 0.28 | 0.05 | $\mathbf{0 . 0 6}$ | 0.27 |
| 2000 | 0.47 | 0.37 | 0.03 | $\mathbf{0 . 0 4}$ | 0.43 |
| 2001 | 0.46 | 0.37 | 0.05 | $\mathbf{0 . 0 6}$ | 0.40 |
| 2002 | 0.04 | 0.04 | 0.06 | $\mathbf{0 . 0 7}$ | -0.03 |
| 2003 | 0.80 | 0.55 | 0.04 | $\mathbf{0 . 0 5}$ | 0.75 |
| 2004 | 0.45 | 0.36 | 0.04 | $\mathbf{0 . 0 5}$ | 0.40 |
|  |  |  |  |  |  |
| Average | 0.33 | 0.25 | 0.04 | $\mathbf{0 . 0 5}$ | 0.28 |

New Jersey Delaware Bay February-April

| Year | $\underline{\underline{Z}}$ | $\underline{\mathbf{A}}$ | $\underline{\mathbf{U}}$ | $\underline{\mathbf{F}}$ | $\underline{\mathbf{M}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.43 | 0.35 | 0.02 | $\mathbf{0 . 0 3}$ | 0.40 |
| 1989 | -0.11 | -0.12 | 0.03 | $\mathbf{0 . 0 3}$ | -0.14 |
| 1990 | 0.45 | 0.36 | 0.04 | $\mathbf{0 . 0 5}$ | 0.40 |
| 1991 | 0.14 | 0.13 | 0.06 | $\mathbf{0 . 0 6}$ | 0.08 |
| 1992 | -0.05 | -0.05 | 0.04 | $\mathbf{0 . 0 4}$ | -0.09 |
| 1993 | 0.63 | 0.47 | 0.05 | $\mathbf{0 . 0 6}$ | 0.56 |
| 1994 | 0.28 | 0.25 | 0.04 | $\mathbf{0 . 0 4}$ | 0.24 |
| 1995 | -0.02 | -0.02 | 0.05 | $\mathbf{0 . 0 5}$ | -0.07 |
| 1996 | 0.15 | 0.14 | 0.03 | $\mathbf{0 . 0 3}$ | 0.12 |
| 1997 | 0.44 | 0.36 | 0.04 | $\mathbf{0 . 0 4}$ | 0.40 |
| 1998 | 0.64 | 0.47 | 0.03 | $\mathbf{0 . 0 4}$ | 0.60 |
| 1999 | 0.33 | 0.28 | 0.05 | $\mathbf{0 . 0 6}$ | 0.27 |
| 2000 | 0.47 | 0.37 | 0.03 | $\mathbf{0 . 0 4}$ | 0.43 |
| 2001 | 0.46 | 0.37 | 0.05 | $\mathbf{0 . 0 6}$ | 0.40 |
| 2002 | 0.04 | 0.04 | 0.06 | $\mathbf{0 . 0 7}$ | -0.03 |
| 2003 | 0.80 | 0.55 | 0.04 | $\mathbf{0 . 0 5}$ | 0.75 |
| 2004 | 0.45 | 0.36 | 0.04 | $\mathbf{0 . 0 5}$ | 0.40 |

Table 30. Summaries of tag-based estimates of annual instantaneous fishing mortality if striped bass $\geq 18$ inches based on the catch equation without assuming constant natural mortality. The table also provides estimates of stock size coastwide of striped bass ages 3+ by solving the equation Kill $=\mathrm{F}^{*}$ (average stock size). Estimates are adjusted for bias caused by live release of recaptured fish.

## Producer Area Programs*

| Year | HUDSON | DE/PA | MDCB | VARAP | Average |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 |  |  |  |  |  |
| 1988 | 0.05 |  | 0.01 |  | $\mathbf{0 . 0 1}$ |
| 1989 | 0.05 |  | 0.01 |  | $\mathbf{0 . 0 1}$ |
| 1990 | 0.08 |  | 0.01 | 0.18 | $\mathbf{0 . 0 6}$ |
| 1991 | 0.10 |  | 0.08 | 0.23 | $\mathbf{0 . 1 1}$ |
| 1992 | 0.11 |  | 0.12 | 0.32 | $\mathbf{0 . 1 6}$ |
| 1993 | 0.13 | 0.15 | 0.15 | 0.28 | $\mathbf{0 . 1 8}$ |
| 1994 | 0.10 | 0.14 | 0.13 | 0.31 | $\mathbf{0 . 1 7}$ |
| 1995 | 0.15 | 0.16 | 0.14 | 0.22 | $\mathbf{0 . 1 6}$ |
| 1996 | 0.19 | 0.19 | 0.22 | 0.17 | $\mathbf{0 . 2 0}$ |
| 1997 | 0.25 | 0.11 | 0.20 | 0.25 | $\mathbf{0 . 2 1}$ |
| 1998 | 0.20 | 0.17 | 0.25 | 0.23 | $\mathbf{0 . 2 3}$ |
| 1999 | 0.17 | 0.14 | 0.25 | 0.20 | $\mathbf{0 . 2 2}$ |
| 2000 | 0.11 | 0.17 | 0.22 | 0.18 | $\mathbf{0 . 1 9}$ |
| 2001 | 0.12 | 0.17 | 0.18 | 0.25 | $\mathbf{0 . 1 9}$ |
| 2002 | 0.10 | 0.18 | 0.16 | 0.19 | $\mathbf{0 . 1 6}$ |
| 2003 | 0.12 | 0.16 | 0.15 | 0.17 | $\mathbf{0 . 1 5}$ |
| 2004 | 0.05 | 0.15 | 0.17 | 0.08 | $\mathbf{0 . 1 3}$ |

* Weighting Scheme: Hudson (0.13); Delaware (0.09);

Chesapeake Bay (0.78), where MD (0.67) and VA (0.33).

## Coast Programs

| Year | MADFW | NYOHS | NJDEL | NCCOOP | Unweighted <br> average |
| :--- | :---: | :---: | :---: | :---: | :---: |
| 1987 |  |  |  |  |  |
| 1988 |  | 0.03 |  | 0.04 | $\mathbf{0 . 0 3}$ |
| 1989 |  | 0.03 | 0.03 | 0.04 | $\mathbf{0 . 0 3}$ |
| 199 |  | 0.05 | 0.05 | 0.09 | $\mathbf{0 . 0 6}$ |
| 1991 |  | 0.06 | 0.06 | 0.10 | $\mathbf{0 . 0 7}$ |
| 1992 | 0.04 | 0.04 | 0.04 | 0.16 | $\mathbf{0 . 0 7}$ |
| 1993 | 0.06 | 0.06 | 0.06 | 0.12 | $\mathbf{0 . 0 8}$ |
| 1994 | 0.04 | 0.04 | 0.04 | 0.12 | $\mathbf{0 . 0 6}$ |
| 1995 | 0.04 | 0.05 | 0.05 | 0.15 | $\mathbf{0 . 0 7}$ |
| 1996 | 0.07 | 0.03 | 0.03 | 0.15 | $\mathbf{0 . 0 7}$ |
| 1997 | 0.10 | 0.04 | 0.04 | 0.23 | $\mathbf{0 . 1 0}$ |
| 1998 | 0.09 | 0.04 | 0.04 | 0.20 | $\mathbf{0 . 0 9}$ |
| 1999 | 0.07 | 0.06 | 0.06 | 0.23 | $\mathbf{0 . 1 0}$ |
| 2000 | 0.10 | 0.04 | 0.04 | 0.16 | $\mathbf{0 . 0 8}$ |
| 2001 | 0.06 | 0.06 | 0.06 | 0.15 | $\mathbf{0 . 0 8}$ |
| 2002 | 0.10 | 0.07 | 0.07 | 0.15 | $\mathbf{0 . 1 0}$ |
| 2003 | 0.09 | 0.05 | 0.05 | 0.16 | $\mathbf{0 . 0 9}$ |
| 2004 | 0.10 | 0.05 | 0.05 | 0.14 | $\mathbf{0 . 0 8}$ |

Table 30 cont'd.
Coastwide Fishing Mortality Rate obtained using the catch equation and obtaining an unweighted average of Coastal and producer area estimates for all fish. Coastwide stock size estimates obtained using Kill = F * (average stock size).

Total Loss Total Stock

| Year Fishing Mortality  includes discards | Size, Thousands <br> 1987 |  |  |
| :--- | :---: | :---: | :---: |
| 1988 | $\mathbf{0 . 0 2}$ | 445 | 19,189 |
| 1989 | $\mathbf{0 . 0 2}$ | 480 | 20,071 |
| 1990 | $\mathbf{0 . 0 6}$ | 921 | 14,978 |
| 1991 | $\mathbf{0 . 0 9}$ | 988 | 10,566 |
| 1992 | $\mathbf{0 . 1 2}$ | 987 | 8,561 |
| 1993 | $\mathbf{0 . 1 3}$ | 1,437 | 11,114 |
| 1994 | $\mathbf{0 . 1 2}$ | 1,867 | 15,944 |
| 1995 | $\mathbf{0 . 1 2}$ | 3,000 | 25,345 |
| 1996 | $\mathbf{0 . 1 4}$ | 3,376 | 24,945 |
| 1997 | $\mathbf{0 . 1 6}$ | 4,580 | 28,873 |
| 1998 | $\mathbf{0 . 1 6}$ | 4,118 | 25,607 |
| 1999 | $\mathbf{0 . 1 6}$ | 3,704 | 23,423 |
| 2000 | $\mathbf{0 . 1 4}$ | 5,044 | 36,980 |
| 2001 | $\mathbf{0 . 1 4}$ | 4,344 | 32,154 |
| 2002 | $\mathbf{0 . 1 3}$ | 3,890 | 30,305 |
| 2003 | $\mathbf{0 . 1 2}$ | 4,836 | 40,036 |
| 2004 | $\mathbf{0 . 1 1}$ | 5,185 | 48,485 |

Table 31. Total length frequencies of fish tagged in 2004 by program.

| Coast Programs |  |  |  | Producer Area Programs |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| TL | MADFW | NYOHS | NJDEP | NCCOOP | DE/PA | MDCB | VARAP | HUDSON |
| 199 |  |  |  |  |  |  |  |  |
| 249 |  | 0 | 0 |  |  |  |  |  |
| 299 |  | 0 | 0 |  |  |  |  |  |
| 349 |  | 0 | 0 | 20 |  | 23 |  |  |
| 399 |  | 14 | 0 | 180 |  | 47 |  |  |
| 449 |  | 219 | 0 | 340 | 63 | 122 |  | 1 |
| 499 |  | 342 | 7 | 505 | 51 | 113 | 155 | 88 |
| 549 |  | 351 | 50 | 408 | 63 | 63 | 212 | 119 |
| 599 |  | 251 | 129 | 242 | 47 | 65 | 220 | 118 |
| 649 | 3 | 150 | 279 | 178 | 38 | 39 | 153 | 150 |
| 699 | 26 | 42 | 449 | 195 | 21 | 43 | 46 | 120 |
| 749 | 93 | 24 | 433 | 262 | 16 | 38 | 43 | 109 |
| 799 | 167 | 17 | 281 | 196 | 22 | 49 | 179 | 136 |
| 849 | 153 | 15 | 169 | 103 | 29 | 86 | 198 | 124 |
| 899 | 98 | 5 | 69 | 43 | 40 | 76 | 109 | 115 |
| 949 | 54 | 6 | 15 | 24 | 26 | 60 | 82 | 60 |
| 999 | 24 | 2 | 3 | 6 | 16 | 34 | 41 | 65 |
| 1049 | 15 | 2 | 1 | 2 | 8 | 15 | 22 | 30 |
| 1099 | 15 | 1 | 0 | 1 | 8 | 13 | 13 | 17 |
| >1099 | 7 | 2 | 0 | 2 | 4 | 7 | 4 | 5 |
| Total | 655 | 1443 | 1885 | 2707 | 452 | 893 | 1477 | 1257 |

Table 32. Age frequencies of tagged fish recaptured in 2004 by program.

| Coast Programs |  |  |  | Producer Area Programs |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| AGE | MADFW | NYOHS | NJDEP | DE/PA | MDCB | VARAP | HUDSON |
| 1 |  |  |  |  |  |  |  |
| 2 |  | 1 |  |  | 0 |  |  |
| 3 |  | 16 | 3 | 4 | 0 |  |  |
| 4 |  | 15 | 53 | 4 | 0 | 13 |  |
| 5 | 1 | 14 | 186 | 4 | 0 | 24 |  |
| 6 | 1 | 13 | 151 | 7 | 0 | 18 |  |
| 7 | 7 | 29 | 76 | 6 | 1 | 13 |  |
| 8 | 8 | 16 | 33 | 13 | 1 | 18 |  |
| 9 | 7 | 11 | 12 | 14 | 0 | 10 |  |
| 10 | 4 | 19 |  | 16 | 2 | 18 |  |
| 11 | 6 | 4 | 1 | 19 | 4 | 6 |  |
| 12 | 7 | 2 |  | 7 | 0 | 8 |  |
| 13 | 3 | 2 |  | 3 | 3 | 7 |  |
| 14 | 2 | 3 |  | 2 | 3 | 2 |  |
| 15 | 2 | 3 |  | 5 | 6 |  |  |
| 16 | 1 | 2 |  | 1 | 0 |  |  |
| 17 |  | 2 |  |  | 1 | 2 |  |
| 18 |  | 1 |  |  | 0 |  |  |
| 19 |  | 3 |  |  | 1 |  |  |
| 20 | 1 |  |  |  | 0 |  |  |
| 21 |  |  |  |  |  |  |  |
| 22 |  |  |  |  |  |  |  |
| Total | 50 | 156 | 515 | 105 | 22 | 139 |  |

Table 33. Distribution of tag recaptures by state (program) and month.

## Coast Programs

Massachusetts (recaptures in 2004 from fish tagged and released during 1992-2004)

| State | Jan. | Feb. | March |  | April | May |  | June |  | July |  | Aug. |  | Sept. |  | Oct. |  | Nov. | Dec. |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ME |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| NH |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| MA |  |  |  |  |  |  |  |  | 4 |  | 10 |  | 7 |  | 5 |  |  |  |  |  | 26 |
| RI |  |  |  |  |  |  |  |  | 1 |  |  |  | 1 |  |  |  | 1 |  |  |  | 3 |
| CT |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| NY |  |  |  |  |  |  | 3 |  | 4 |  |  |  |  |  |  |  | 2 | 2 | 2 | 1 | 12 |
| NJ |  |  |  | 1 |  | 2 | 2 |  | 2 |  |  |  |  |  | 1 |  | 1 | 7 | 7 |  | 16 |
| DE |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| MD |  |  |  | 1 |  | 2 | 3 |  |  |  |  |  |  |  |  |  |  | 1 | 1 |  | 7 |
| VA | 2 | 2 |  | 2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 3 | 11 |
| NC | 1 | 1 | 2 |  |  | 1 |  |  |  |  |  |  |  |  |  |  |  |  |  | 1 | 6 |
| Total | 3 | 3 | 2 | 4 |  | 5 | 8 | - | 11 |  | 10 |  | 8 |  | 6 |  | 4 | 15 |  | 5 | 81 |

New York - Ocean Haul Seine (recaptures in 2004 from fish tagged/release during 1988-2004)


Table 33 cont'd.
New Jersey - Delaware Bay (recaptures in 2004 from fish tagged/release during 1989-2004)

| State | Jan. | Feb. | March | April | May | June | July | Aug. | Sept. | Oct. | Nov. | Dec. | Total |
| :--- | :--- | :--- | :--- | :--- | :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| ME |  |  |  |  |  | 4 | 7 | 3 | 3 |  |  | 1 | 18 |
| NH |  |  |  |  |  | 2 | 1 | 3 |  |  |  |  | 6 |
| MA |  |  |  |  | 11 | 41 | 44 | 34 | 17 | 4 |  |  | 151 |
| RI |  |  |  |  | 6 | 12 | 3 | 6 | 2 | 2 |  |  | 31 |
| CT |  |  |  |  | 7 | 4 | 5 | 1 | 3 | 3 |  |  | 23 |
| NY |  |  |  | 1 | 18 | 15 | 11 | 13 | 9 | 21 | 19 | 2 | 109 |
| NJ |  | 1 |  | 9 | 34 | 6 | 5 |  | 1 | 10 | 31 | 7 | 104 |
| PA |  |  |  |  | 1 |  |  |  |  |  |  |  | 1 |
| DE |  |  |  | 2 | 1 |  |  |  |  | 2 | 2 | 1 | 8 |
| MD |  |  |  | 8 | 7 |  |  |  |  | 1 | 2 |  | 18 |
| VA | 3 | 4 | 1 | 1 |  |  |  |  |  |  | 2 | 3 | 14 |
| NC | 6 | 2 |  |  | 1 |  |  |  |  |  | 1 | 6 | 16 |
| Total | 9 | 7 | 1 | 21 | 86 | 84 | 76 | 60 | 35 | 43 | 57 | 20 | 499 |



Table 33 cont'd.

## Producer Area Programs

Delaware / Pennsylvania - Delaware River
(recaptures in 2004 from fish tagged and released during 1992-2004)

| State | Jan. | Feb. | March | April | May |  | June |  | July |  | Aug. |  | Sept. | Oct. |  | Nov. | Dec. | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ME |  |  |  |  |  | 1 |  |  |  | 1 |  |  |  |  |  |  |  | 2 |
| NH |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| MA |  |  |  |  |  |  |  | 4 |  | 4 |  | 6 |  | 3 |  |  |  | 17 |
| RI |  |  |  |  |  |  |  |  |  |  |  |  |  | 1 | 1 |  |  | 2 |
| CT |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| NY |  |  |  |  |  | 1 |  | 3 |  | 1 |  | 1 |  |  |  |  |  | 7 |
| NJ |  |  |  | 2 |  | 7 |  | 11 |  | 5 |  | 1 |  | 1 | 7 | 12 | 1 | 47 |
| PA |  |  |  |  |  | 2 |  |  |  |  |  |  |  |  |  |  |  | 2 |
| DE |  |  |  |  | 1 | 2 |  |  |  |  |  |  |  | 1 | 3 |  |  | 11 |
| MD |  |  |  |  | 4 |  |  | 3 |  |  |  |  |  | 2 |  |  |  | 13 |
| VA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 2 | 2 |
| NC | 4 | 4 | 1 | 1 |  | 1 |  |  |  |  |  |  |  |  |  |  |  | 7 |
| Total | 4 | 4 | 1 | 3 | 5 | 14 | - | 21 |  | 11 |  | 8 |  | 8 | 11 | 21 | 3 | 110 |

Maryland - Chesapeake Bay Spring Spawning Stock
(recaptures in 2004 from fish tagged and released during 1992-2004)

| State | Jan. | Feb. | March | April |  | May |  | June |  | July |  | Aug. |  | Sept. |  | Oct. |  | Nov. | Dec. | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ME |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| NH |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| MA |  |  |  |  |  |  | 1 |  | 2 |  | 7 |  | 6 |  | 3 |  | 1 |  |  | 20 |
| RI |  |  |  |  |  |  |  |  |  |  |  |  | 2 |  |  |  |  |  |  | 2 |
| CT |  |  |  |  |  |  |  |  | 2 |  | 1 |  |  |  |  |  |  |  |  | 3 |
| NY |  |  |  |  |  |  |  |  | 2 |  |  |  | 3 |  |  |  | 1 |  |  | 6 |
| NJ |  |  |  |  | 1 |  | 1 |  |  |  |  |  |  |  |  |  | 1 |  |  | 4 |
| PA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| DE |  |  |  |  |  |  |  |  | 1 |  |  |  | 1 |  |  |  |  |  |  | 3 |
| MD |  | 1 | 1 | 1 | 1 |  | 8 |  | 12 |  | 13 |  | 6 |  | 9 |  | 10 |  | 1 | 69 |
| VA |  |  | 1 | 2 |  |  | 1 |  | 2 |  | 1 |  | 1 |  | 1 |  | 5 |  | 7 | 26 |
| NC |  | 5 |  | 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 1 | 7 |
| Total |  | 6 | 2 | 4 | 2 | , | 11 |  | 21 | - | 22 |  | 19 | - | 13 | 析 | 18 | 13 | 9 | 140 |

Table 33 cont'd.
Virginia - Rappahannock River (recaptures in 2004 from fish tagged and released during 1992-2004)

| State | Jan. | Feb. | March | April |  | May |  | June | Ju |  | Aug. |  | Sept. | Oct |  | ov. | Dec | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| NH |  |  |  |  |  |  |  |  | 1 |  |  |  |  |  |  |  |  | 1 |
| MA |  |  |  |  |  |  | 2 |  | 3 | 8 |  | 2 |  | 1 |  |  |  | 16 |
| RI |  |  |  |  |  |  |  |  | 1 | 2 |  | 1 |  | 1 |  |  |  | 5 |
| CT |  |  |  |  |  |  |  |  |  | 1 |  |  |  | 1 |  |  |  | 2 |
| NY |  |  |  |  |  |  | 4 |  | 4 | 2 |  | 1 |  | 2 | 4 |  |  | 19 |
| NJ |  |  |  |  |  |  | 2 |  | 1 | 1 |  |  |  |  |  |  |  | 6 |
| DE |  |  |  |  |  |  |  |  | 2 |  |  |  |  |  | 1 |  |  | 4 |
| MD |  |  |  |  | 1 |  | 5 |  | 12 | 6 |  |  |  |  | 2 |  |  | 26 |
| VA |  |  |  |  | 11 |  | 9 |  | 9 | 3 |  | 1 |  |  | 7 |  |  | 856 |
| NC |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 22 |
| Total | $0$ | 0 | 0 | 0 | 12 |  | 22 |  | 33 | 23 |  | 5 |  | 5 | 14 | 1 |  | 0 137 |

Hudson River
(recaptures in 2004 from fish tagged and released during 1992-2004)

| State | Jan. | Feb. |  | March | April |  | May |  | June |  | July |  | Aug. | Sep |  | Oct. |  | ov. | Dec |  | tal |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| ME |  |  |  |  |  |  |  |  |  |  |  |  |  | 1 | 2 |  |  |  | 1 |  | 4 |
| NH |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| MA |  |  |  |  |  |  |  | 1 |  | 7 |  | 13 |  | 4 | 4 |  | 1 |  | 1 |  | 41 |
| RI |  |  |  |  |  |  |  |  |  | 6 |  | 5 |  | 3 |  |  | 1 |  | 1 |  | 19 |
| CT |  |  |  |  |  | 1 |  | 1 |  | 7 |  | 6 |  | 4 |  |  | 2 |  | 1 |  | 24 |
| NY |  |  |  |  |  | 5 |  | 37 |  | 15 |  | 16 |  | 0 | 11 |  | 19 | 10 |  | 1 | 124 |
| NJ |  |  |  |  |  |  |  | 2 |  | 13 |  | 3 |  |  |  |  | 4 | 2 |  | 4 | 46 |
| PA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| DE |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 1 |  | 1 |
| MD |  |  |  |  |  |  |  |  |  | 1 |  |  |  |  |  |  | 1 |  |  |  | 2 |
| VA |  | 1 | 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5 | 7 |
| NC |  | 3 |  |  |  |  |  |  |  |  |  |  |  | 1 |  |  |  |  |  | 5 | 9 |
| Total |  | 4 | 1 |  | 0 | 6 |  | 41 |  | 49 |  | 43 |  | 33 | 22 |  | 28 | 3 | 5 | 15 | 277 |

## Appendix A. Criteria to evaluate the VPA indices

(Approved and accepted by SB TC)
The Workshop participants developed a list of evaluation steps that should be applied to each index. The state agencies should use the evaluation list for each state survey. Each program should be analyzed to determine if the survey is conducted at the appropriate time of year, i.e. bracketing the correct spawning period. Similarly, the survey design should be reviewed by the state to determine if the sampling area is correct. If the state determines there is a lot of noise in the data, the state should attempt to refine the data. For instance, if some of the stations catch striped bass consistently and others do not, can something be done to refine these data? The states should identify if the indices are sex-specific indices or age-specific due to survey design. Because a self-evaluation by each state could be subjective, the Technical Committee should evaluate the state's program evaluation and make a recommendation to the Striped Bass Stock Assessment Subcommittee.

1. Evaluate design and best method to evaluate uncertainty of index.
2. Assess the index and/or improve the index to get the best signal.
3. Validate the index before use in the VPA.
a. Sensitivity of the VPA results to the influence each index.
b. Validate an index to a JAI, where possible.
c. Longitudinal catch curves, to determine the cohort trends.
d. Plots of age specific index v. year to see if cohorts are moving in a specific direction.
4. Evaluation by the agency conducting the survey
a. Rank (weight) index
b. Criticisms/Supporting Evidence
5. Evaluate by the Striped Bass Technical Committee
a. Evaluate index based on survey design, precision, and ability to track cohorts or portion of the stock targeted.
b. Provide recommendations to the Striped Bass Stock Assessment Subcommittee on which indices should be used in the assessment.

# Appendix B. Estimation of Wave-1 Harvest in North Carolina and Virginia 

DT: 7/11/2005
TO: ASMFC Striped Bass Technical Committee
FR: Joseph Grist, ASMFC
RE: MRFSS North Carolina Wave-1 2004 harvest

## Introduction

During the March 2005 Striped Bass Technical Committee (STB TC) meeting, the results for the 2004 wave-1 North Carolina (NC) harvest were reported. This was the first time wave-1 was directly sampled by the Marine Recreational Fisheries Statistics Survey (MRFSS), and the results were both predictable and a cause for concern. A total of 177,288 striped bass (equivalent to $3,615,670 \mathrm{lb}$ ) were harvested during wave-1 in North Carolina.

Anecdotal knowledge has suggested that North Carolina, Virginia, and possibly other states had a sizeable wave-1 fishery. The 2004 wave-1 harvest values for North Carolina and the wave-1 tag return data (Figure 1) for North Carolina and Virginia support this suggestion. However, information is still lacking on what the previous annual harvest rates were, as well as the level of exploitation in Virginia and elsewhere during wave-1. The STB TC requested an examination of the data that included suggestions for how to incorporate these data efficiently into the coastwide STB assessment.

The goal of this analysis is to determine if tag return data during wave-6 and wave-2 are correlated with the reported total catch and, if so, if a proxy ratio may be utilized to backcalculate wave-1 data for North Carolina and Virginia.

## Data

Striped bass tag return data from North Carolina and Virginia were provided by the U.S. Fish and Wildlife Service (USFWS). Data were queried from the MRFSS website (http://www.st.nmfs.gov/st1/recreational/queries/effort/effort time_series.html) on July 11, 2005 for North Carolina and Virginia, having selected variables by harvest (A+B1), all oceans combined, and all modes combined.

## Methods

Tag return and MRFSS data were merged by wave and by year and were analyzed for each state. SAS 9.1 was utilized to calculate Pearson's correlation coefficient (PROC CORR), generate linear regressions, and conduct ANOVA or analysis of variance (PROC REG) to test for similarities between tag return and total catch data by wave. Only wave-6 (November and December) and Wave-2 (March and April) data were analyzed.
Results

## North Carolina

Tag returns were positively correlated with total catch (0.5828) during wave-6 (Figure 2). ANOVA indicated significant evidence $(p$-value $=0.0366)$ that total catch could explain the proportion of tag returns during wave-6.

Tag returns were positively correlated with total catch (0.9518) during wave-2 (Figure 3). ANOVA indicated significant evidence ( $p$-value $<0.0001$ ) that total catch could explain the proportion of tag returns during wave- 2 .

## Virginia

Tag returns were positively correlated with total catch (0.5827) during wave-6 (Figure 4). Although ANOVA did not indicate statistically significant evidence $(p$-value $=0.0599)$ that total catch could explain the proportion of tag returns during wave 6 , the given $p$-value indicates suggestive, but inconclusive, evidence that the null hypothesis is false, possibly representing biological significance.

Tag returns were slightly negatively correlated with total catch (-0.4007) during wave-2 (Figure 5). ANOVA did not indicate significant evidence $(p$-value $=0.4311)$ that total catch could explain the proportion of tag returns during wave-2. However, the tag return data were not consistent from year to year and a negative correlation was expected.

## Summary

The 2004 wave- 1 total catch for North Carolina corresponds with observed recreational effort that begins during wave- 6 and continues into wave- 1 throughout the coastal waters of northeastern North Carolina and southeastern Virginia (Sara Winslow, NCDMF, personal communication).

Analysis indicates that tag return data can be used to explain total catch in wave-6 and wave-2 in North Carolina. If the assumption that wave- 1 follows a similar trend is acceptable by the STB TC, then wave-1 data before 2004 could be back-calculated for North Carolina striped bass harvest. There are two possible methods for back-calculation (Figure 6). One would be using the direct 2004 ratio of tag returns to reported total catch. The other would be to use the combined ratio of tag returns to total catch for both wave- 6 and wave- 2 .

Correlation analysis for Virginia did indicate total catch could be explained by tag returns, although ANOVA did not provide strong evidence for or against the reported correlation. However, tag return evidence does show a wave-1 striped bass fishery is occurring in Virginia (Figure 1), and using the wave-6 mean ratio of tag returns to reported total catch for 1996-2004 could be utilized to back-calculate the wave-1 striped bass recreational fishery (Figure 7).

## Wave-1 Tag Returns



Figure 1. Wave-1 tag returns for Virginia and North Carolina.


Figure 2. Wave-6 tag returns versus total catch for North Carolina.

Wave 2: North Carolina


Figure 3. Wave-2 tag returns versus total catch for North Carolina.


Figure 4. Wave-6 tag returns versus total catch for Virginia.

Wave 2: STB


Figure 5. Wave-2 tag returns versus total catch for Virginia.

Catch Projection: North Carolina Wave-1


Figure 6. Comparison of catch projections for North Carolina wave-1.

## Catch Projection: Virginia Wave-1



Figure 7. Catch projection for Virginia wave-1.

# Appendix C. Wave-1 Total Recreational Catch for North Carolina and Virginia: Final Calculations 

DT: August 10, 2005
TO: Striped Bass Stock Assessment Sub-Committee
FR: Joseph Grist, ASMFC
RE: Wave-1 total recreational catch for NC and VA final calculations
Based on the report presented to the Striped Bass TC on July 11, 2005 concerning the North Carolina and Virginia MRFSS wave-1 recreational catch, Table 1 contains calculations for total catch for each state.

North Carolina: Wave-1 total catch for 1996-2003 is based on the NC specific 2004 wave-1 ratio of tag returns to MRFSS total catch numbers. There were 47 tags returned during the wave1 fishery period for the ocean fishery. The MRFSS reported catch (A+B1) was 177,288 striped bass during the same period. This resulted in a 2004 ratio tags to catch of 0.000265 . This ratio was applied to the wave- 1 tag returns for the NC ocean fishery to provide a back-calculated total catch for wave-1 in NC.

Virginia: Unlike NC, a 2004 wave-1 total catch was not reported. However, analysis of the tag returns suggested that a winter fishery similar to that of North Carolina occurred off VA during 2004. The July $11^{\text {th }}$ report to the TC did indicate that VA wave- 6 tag returns were positively correlated to catch and implied biological significance, though wave-2 analysis did not. Personal communication with Sara Winslow (NCDMF) confirmed that the winter fishery begins in the latter half of wave- 6 and continues into wave- 1 in northeastern NC, and similar trends would be expected for southeastern VA. Anecdotally, this suggested that wave-6 and wave-1 catch would show some level of correlation in fishing activity. Using known wave-1 tag returns, a mean ratio (0.000167) of tag returns to catch for VA wave-6, 1996-2004, was utilized to back-calculate the total wave-1 catch.

Table 1. Wave-1 catch values for North Carolina and Virginia, 1996-2004.

| Year | Total catch values (projected) |  |
| :---: | :---: | :---: |
|  | NC | VA |
| 1996 | 18,860 | 5,985 |
| 1997 | 49,037 | 83,793 |
| 1998 | 15,088 | 89,778 |
| 1999 | 18,860 | 107,734 |
| 2000 | 7,544 | 53,867 |
| 2001 | 18,860 | 53,867 |
| 2002 | 75,442 | 89,778 |
| 2003 | 79,214 | 53,867 |
| 2004 | 177,288* | 155,616 |
| *actual ca |  |  |

# Appendix D. Analysis and Discussion of the 1998-2002 Striped Bass Coastwide Weight-at-Age 

Analysis and Discussion of the<br>1998-2002 Striped Bass Coastwide Weight-at-Age

Prepared for the
Striped Bass Stock Assessment Sub-Committee Meeting
August 9 - 11, 2005

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## Introduction

A crucial element of the yearly catch-age based virtual population analyses (VPA) of Atlantic striped bass is the calculation of biomass of the mixed coastal stock. This calculation requires coastwide weight-at-age (WAA). The coastwide WAA has consistently been calculated as a weighted mean:

$$
\begin{align*}
& \text { State WAA }=\Sigma \text { (state WAA * \% state CAA by numbers) }  \tag{Eqn. 1}\\
& \text { Coastwide WAA }=\Sigma(\text { State WAA * state } \% \text { coastwide CAA }) \tag{Eqn. 2}
\end{align*}
$$

The current VPA analysis uses a time series dating back to 1982. The yearly values were not calculated on a yearly basis, however. In 1997, the values for 1982-1997 were developed. These values were developed using data from all states, subdividing each year into quarterly time periods to account for growth, and weighting by numbers of fish. (Details of developing weights at age for 1982 to 1996 can be found in NEFSC Lab Ref. 98-03.) Coastwide WAA was not re-calculated in 1998 or 1999. Instead, the 1997 values were used as these years' values. The 2000, 2001 and 2002 coastwide WAA were developed at the Stock Assessment Subcommittee Workshops, weighted by total weight of fish, using readily available data sets. Therefore, the methodology and data sets used for these calculations were not consistent, either with the methodology used for the 1982-1997 WAA or with each other. The 2000-2002 values showed an apparent decline in WAA, but it was impossible to determine if this apparent trend was due to the change in method or a true change in WAA.

In 2004, a standardized report format was developed that calculated WAA as part of the CAA calculations. The 2003 coastwide WAA was developed using all states' data:

- Maine and New Hampshire recreational harvest and discards,
- Massachusetts recreational and commercial catch,
- Rhode Island recreational and commercial catch,
- Connecticut recreational catch,
- New York recreational catch and commercial landings,
- New Jersey recreational catch,
- Delaware recreational and commercial catch,
- Maryland recreational and commercial catch,
- Virginia recreational and commercial catch, and
- North Carolina recreational and commercial catch.

An apparent decline was observed between the 2001and 2002 coastwide WAA - only 2 of 13 age-classes of harvested fish did not show a reduction in WAA (Table 1). Due to concerns about this apparent decrease in coastwide WAA and the inability to compare 1998-2002 with the rest of the time series, the subcommittee decided to re-calculate these coastwide WAA values.

## Methods: Recalculation of the 1998-2002 values.

All states were requested to provide the 1998-2002 time series of WAA, landings and discards. Because information was not received from all states, it was decided to develop the coastwide WAA from information for states with greatest catch. For 1998-2001, the coastwide WAA was calculated using the 5 major harvester states (MA, NY, NJ, MD, VA), NH and CT (Table 2). For 2002, data were available to include RI and DE (Table 3). WAA was calculated as the weighted mean, weighted by numbers for commercial harvest, recreational harvest, and recreational discard. Annual state removals were taken from the time series tables for commercial harvest, recreational harvest and recreational discard numbers in the 2004 coastwide compliance report summary prepared by Gary Sheppard if not provided by state. WAA for the nearest neighboring state was used if that state's WAA was not available. The oldest age group was designated "13+", and 1982-1997 "13+" values were recalculated as the arithmetic averages of 13- to 15 -year-old age class values. A constraint imposed by the 1998-2002 data was that an annual time frame was used for all calculations, as opposed to the finer time frame used in the 1982-1997 and 2003 calculations. The time series matrix of WAA including re-calculated values is presented in Table 4.

## Discussion

The apparent decrease in WAA from 2000-2002 within the "old" WAA time series. Most age classes showed a decrease between 2000 and 2002 ( 14 of 15 age-classes) (Table 2). However, examination of the development of the WAA revealed that this decrease was due to differences in the development of the values. Because average WAA greater for coastal than Chesapeake Bay states for all harvested age classes, calculations are skewed if the harvest proportion is not used in the WAA calculations.

Evaluation of the apparent decline between 2001-2002 values. The 1982-1997 coastwide WAA time series was developed using all states' data. In contrast, the 2001 coastwide WAA was developed without data from RI, CT, MD and NC. Due to comparatively low harvest, RI, CT and NC do not contribute strongly to the coastwide WAA. However, the exclusion of MD data from the 2001 calculation had a major influence on the coastwide value. Without the MD numbers factoring in to the average, the coastwide WAA was disproportionately weighted by MA (Figure 1, Table 5). This is significant because MD is a Chesapeake Bay harvest state and MA is a coastal harvest state. Based on data from 1982-1997, the majority of fish harvested in Chesapeake Bay (ages 3-11) were, on average, 2.6 kg ( 5.7 lb ) smaller than coastal fish (Table 6). The unnaturally strong contribution of MA in the 2001 WAA, followed by the strong contribution of MD fish in the 2002 WAA, certainly contributed to the observed decline in the
coastwide WAA.
Patterns in WAA from 2000 - 2003 within the recalculated WAA time series. Coastwide WAA values for 2000 to 2002 were recalculated using a consistent method that was considered functionally equivalent to the method used for earlier calculations. Although a subset of states was used, these states constitute the majority of the harvest and therefore maintained the overall harvest proportion throughout the WAA calculations. In contrast to the earlier values, these values showed a consistent increase across the 2000-2003 time frame (Table 4). Between 2000 and 2001, 11 of the 13 age classes showed an increase in WAA, between 2002 and 2003, 12 of the 13 age classes showed an increase in WAA. The 2003 WAA was developed from information provided by all states for the 2003 stock assessment. Comparison of the 2003 WAA against the mean values for 2000-2002 showed an increase in 11 of 13 age classes.

Comparison of "old" vs. recalculated WAA values from 2000 - 2002. Although the recalculated WAA values showed an increase across the 2000-2003 time frame, these values were lower than the mean of the 1982-1996 time series (Table 7).

## Future Work.

Future years' WAA will be calculated from information provided in stock assessment "Compliance Report Template", and will therefore include all states' data. No recommendations are suggested to improve calculation methodology for future years.

It would be useful to determine if there truly was a decrease between the 1982-96 WAA and the 1998-2003 WAA. However, data are not available to recalculate 1982-2002 WAA using the current method, nor are data available to recalculate 2000-03 using the earlier method.

Figure 1. Composition of Striped Bass Coastwide WAA by State.
1982-1997 coastwide WAA shows a fairly even distribution from the 5 major harvest (by numbers) states (MA, NY, NJ, MD, VA). 2001 WAA is dominated by MA. 2002 WAA shows a strong contribution from MD and VA (Chesapeake Bay harvest states).


Table 1. Striped Bass Coastwide WAA (kg) Time Series Used for the 2002 Stock Assessment.
1997-1999 values are identical. Note the apparent decline in WAA between 2001-2002.

| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | $\begin{gathered} \text { Age } \\ 8 \end{gathered}$ | 9 | 10 | 11 | 12 | 13 | 14 | 15 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1982 | 0.13 | 0.64 | 1.09 | 1.54 | 2.42 | 3.75 | 4.83 | 5.79 | 6.20 | 8.68 | 10.80 | 11.20 | 12.97 | 13.26 | 15.91 |
| 1983 | 0.20 | 0.55 | 0.94 | 1.37 | 2.37 | 3.29 | 3.77 | 5.36 | 6.01 | 8.10 | 9.57 | 10.39 | 11.11 | 11.10 | 11.12 |
| 1984 | 0.24 | 0.60 | 1.69 | 1.62 | 2.67 | 3.39 | 5.07 | 5.65 | 6.76 | 7.76 | 8.41 | 12.65 | 10.65 | 11.75 | 14.75 |
| 1985 | 0.06 | 0.61 | 1.07 | 1.66 | 2.19 | 3.59 | 4.91 | 5.46 | 6.77 | 7.45 | 9.00 | 10.69 | 11.42 | 14.34 | 15.98 |
| 1986 | 0.14 | 0.57 | 1.27 | 2.40 | 2.44 | 3.12 | 3.95 | 5.05 | 5.44 | 6.09 | 7.75 | 9.16 | 10.97 | 11.55 | 15.83 |
| 1987 | 0.20 | 0.77 | 1.41 | 2.11 | 2.50 | 2.91 | 3.61 | 4.74 | 5.52 | 6.49 | 7.77 | 9.78 | 11.38 | 11.62 | 16.46 |
| 1988 | 0.31 | 0.91 | 1.10 | 1.98 | 3.12 | 4.02 | 4.38 | 4.70 | 5.24 | 5.62 | 8.58 | 10.40 | 11.50 | 11.31 | 17.00 |
| 1989 | 0.16 | 0.83 | 1.22 | 2.23 | 3.06 | 4.53 | 5.37 | 6.23 | 6.04 | 8.68 | 8.94 | 9.74 | 13.04 | 9.93 | 17.11 |
| 1990 | 0.08 | 0.89 | 1.14 | 2.05 | 2.35 | 3.83 | 4.91 | 5.96 | 5.70 | 5.97 | 7.44 | 9.08 | 9.36 | 10.80 | 17.65 |
| 1991 | 0.21 | 0.92 | 1.29 | 2.17 | 2.62 | 3.17 | 4.81 | 5.64 | 6.46 | 6.24 | 9.46 | 8.30 | 9.62 | 15.96 | 17.09 |
| 1992 | 0.10 | 0.69 | 1.31 | 1.93 | 2.81 | 3.67 | 4.90 | 5.79 | 6.96 | 8.15 | 9.77 | 12.44 | 13.10 | 11.15 | 17.65 |
| 1993 | 0.07 | 0.76 | 1.31 | 1.99 | 2.77 | 3.58 | 4.80 | 6.11 | 7.03 | 8.01 | 9.53 | 10.76 | 14.45 | 13.85 | 15.36 |
| 1994 | 0.24 | 1.05 | 1.69 | 2.21 | 2.85 | 3.50 | 4.94 | 6.20 | 6.80 | 7.53 | 9.73 | 10.69 | 11.38 | 9.06 | 17.75 |
| 1995 | 0.28 | 0.70 | 1.35 | 2.18 | 2.77 | 3.65 | 5.38 | 6.16 | 7.27 | 8.86 | 7.57 | 9.73 | 13.97 | 15.65 | 20.37 |
| 1996 | 0.14 | 1.05 | 1.47 | 2.32 | 3.23 | 4.52 | 6.39 | 7.11 | 7.81 | 9.20 | 9.31 | 10.10 | 11.36 | 12.45 | 17.30 |
| 1997 | 0.13 | 0.62 | 1.18 | 2.46 | 2.81 | 3.64 | 4.51 | 5.07 | 6.73 | 9.17 | 9.94 | 10.24 | 11.94 | 14.49 | 17.92 |
| 1998 | 0.13 | 0.62 | 1.18 | 2.46 | 2.81 | 3.64 | 4.51 | 5.07 | 6.73 | 9.17 | 9.94 | 10.24 | 11.94 | 14.49 | 17.92 |
| 1999 | 0.13 | 0.62 | 1.18 | 2.46 | 2.81 | 3.64 | 4.51 | 5.07 | 6.73 | 9.17 | 9.94 | 10.24 | 11.94 | 14.49 | 17.92 |
| 2000 | 0.14 | 1.05 | 1.47 | 2.32 | 3.23 | 4.52 | 6.39 | 7.11 | 7.81 | 9.20 | 9.31 | 10.10 | 11.36 | 12.45 | 17.30 |
| 2001 | 0.13 | 0.62 | 1.17 | 2.46 | 2.81 | 3.63 | 4.51 | 5.07 | 6.73 | 9.17 | 9.94 | 10.24 | 11.94 | 14.49 | 17.92 |
| 2002 | 0.82 | 0.81 | 1.25 | 1.75 | 2.47 | 3.30 | 4.16 | 5.48 | 6.36 | 7.45 | 8.75 | 8.89 | 9.99 | 11.03 | 13.95 |

Table 2. Revised Time Series of Striped Bass Coastwide WAA (kg).

| Year | Age |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ | $\mathbf{6}$ | $\mathbf{7}$ | $\mathbf{8}$ | $\mathbf{9}$ | $\mathbf{1 0}$ | $\mathbf{1 1}$ | $\mathbf{1 2}$ | $\mathbf{1 3 +}$ |
| 1982 | 0.1 | 0.6 | 1.1 | 1.5 | 2.4 | 3.7 | 4.8 | 5.8 | 6.2 | 8.7 | 10.8 | 11.2 | 14.0 |
| 1983 | 0.2 | 0.6 | 0.9 | 1.4 | 2.4 | 3.3 | 3.8 | 5.4 | 6.0 | 8.1 | 9.6 | 10.4 | 11.1 |
| 1984 | 0.2 | 0.6 | 1.7 | 1.6 | 2.7 | 3.4 | 5.1 | 5.7 | 6.8 | 7.8 | 8.4 | 12.7 | 12.4 |
| 1985 | 0.1 | 0.6 | 1.1 | 1.7 | 2.2 | 3.6 | 4.9 | 5.5 | 6.8 | 7.4 | 9.0 | 10.7 | 13.9 |
| 1986 | 0.1 | 0.6 | 1.3 | 2.4 | 2.4 | 3.1 | 4.0 | 5.0 | 5.4 | 6.1 | 7.8 | 9.2 | 12.8 |
| 1987 | 0.2 | 0.8 | 1.4 | 2.1 | 2.5 | 2.9 | 3.6 | 4.7 | 5.5 | 6.5 | 7.8 | 9.8 | 13.2 |
| 1988 | 0.3 | 0.9 | 1.1 | 2.0 | 3.1 | 4.0 | 4.4 | 4.7 | 5.2 | 5.6 | 8.6 | 10.4 | 13.3 |
| 1989 | 0.2 | 0.8 | 1.2 | 2.2 | 3.1 | 4.5 | 5.4 | 6.2 | 6.0 | 8.7 | 8.9 | 9.7 | 13.4 |
| 1990 | 0.1 | 0.9 | 1.1 | 2.1 | 2.4 | 3.8 | 4.9 | 6.0 | 5.7 | 6.0 | 7.4 | 9.1 | 12.6 |
| 1991 | 0.2 | 0.9 | 1.3 | 2.2 | 2.6 | 3.2 | 4.8 | 5.6 | 6.5 | 6.2 | 9.5 | 8.3 | 14.2 |
| 1992 | 0.1 | 0.7 | 1.3 | 1.9 | 2.8 | 3.7 | 4.9 | 5.8 | 7.0 | 8.2 | 9.8 | 12.4 | 14.0 |
| 1993 | 0.1 | 0.8 | 1.3 | 2.0 | 2.8 | 3.6 | 4.8 | 6.1 | 7.0 | 8.0 | 9.5 | 10.8 | 14.6 |
| 1994 | 0.2 | 1.1 | 1.7 | 2.2 | 2.9 | 3.5 | 4.9 | 6.2 | 6.8 | 7.5 | 9.7 | 10.7 | 12.7 |
| 1995 | 0.3 | 0.7 | 1.3 | 2.2 | 2.8 | 3.7 | 5.4 | 6.2 | 7.3 | 8.9 | 7.6 | 9.7 | 16.7 |
| 1996 | 0.1 | 1.0 | 1.5 | 2.3 | 3.2 | 4.5 | 6.4 | 7.1 | 7.8 | 9.2 | 9.3 | 10.1 | 13.7 |
| 1997 | 0.1 | 0.6 | 1.2 | 2.5 | 2.8 | 3.6 | 4.5 | 5.1 | 6.7 | 9.2 | 9.9 | 10.2 | 14.8 |
| 1998 | 0.4 | 0.8 | 1.2 | 1.6 | 2.2 | 2.9 | 4.7 | 5.7 | 6.8 | 7.0 | 7.8 | 9.9 | 11.9 |
| 1999 | 0.6 | 0.9 | 1.1 | 1.4 | 1.9 | 2.5 | 3.4 | 5.0 | 6.6 | 7.8 | 8.7 | 9.8 | 12.0 |
| 2000 | 0.4 | 0.6 | 1.1 | 1.5 | 2.0 | 2.8 | 3.9 | 5.1 | 7.1 | 7.4 | 9.7 | 10.7 | 13.6 |
| 2001 | 0.2 | 0.4 | 1.1 | 1.8 | 2.2 | 3.2 | 4.1 | 5.0 | 6.4 | 7.8 | 8.6 | 8.3 | 10.9 |
| 2002 | 0.1 | 0.3 | 1.1 | 1.5 | 2.2 | 3.2 | 4.2 | 5.5 | 6.0 | 7.6 | 9.1 | 9.7 | 11.5 |

Table 3. Comparison of 2001\& 2002 Data Used to Develop Striped Bass Coastwide WAA.

| STATE | 2001 |  |  | 2002 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | SURVEYS | \% WAA | \% HARVEST | SURVEYS | \% WAA | \% HARVEST |
| ME | COMM (harv, discards) | 1 | 1 | X | 0 | 2 |
| NH | COMM (harv, discards) | 3 | 1 | REC | 1 | 1 |
| MA | COMBINED | 74 | 16 | COMBINED | 32 | 20 |
| RI | X | 0 | 5 | $X$ | 0 | 5 |
| CT | X | 0 | 3 | X | 0 | 3 |
| NY | COMM \& REC | 6 | 13 | COMM \& REC | 11 | 13 |
| NJ | REC | 10 | 23 | REC | 17 | 19 |
| DE | COMM | $<1$ | 2 | $X$ | 0 | 1 |
| MD | $X$ | 0 | 17 | COMM (C.BAY) | 22 | 15 |
| VA | COMM \& REC | 6 | 17 | COMM \& REC | 17 | 19 |
| NC | X | 0 | 3 | $X$ | 0 | 3 |

Table 4. Comparison of Average Striped Bass WAA (lb) for "Coastal" (MA, NY, NJ) and "Chesapeake Bay" (MD and VA) States, based 1982-1997 Values.

| Age | Coastal | CBay | $\boldsymbol{\Delta}$ |
| :---: | :---: | :---: | :---: |
| $\mathbf{1}$ | 1.8 |  |  |
| $\mathbf{2}$ | 1.9 | 2.3 | -0.4 |
| $\mathbf{3}$ | 3.3 | 2.4 | 0.9 |
| $\mathbf{4}$ | 4.7 | 2.7 | 2.0 |
| $\mathbf{5}$ | 6.7 | 3.5 | 3.2 |
| $\mathbf{6}$ | 8.3 | 5.5 | 2.8 |
| $\mathbf{7}$ | 10.1 | 7.4 | 2.8 |
| $\mathbf{8}$ | 12.9 | 10.4 | 2.5 |
| $\mathbf{9}$ | 14.9 | 12.3 | 2.6 |
| $\mathbf{1 0}$ | 17.4 | 14.1 | 3.4 |
| $\mathbf{1 1}$ | 20.4 | 17.3 | 3.0 |
| $\mathbf{1 2}$ | 22.8 | 14.9 | 7.8 |
| $\mathbf{1 3}$ | 24.9 | 17.7 | 7.2 |
| $\mathbf{1 4}$ | 27.9 | 19.4 | 8.5 |
| $\mathbf{1 5}$ | 35.1 | 15.8 | 19.4 |

Table 5. Information Used to Calculate 1998-2002 Striped Bass Coastwide WAA.

| REMOVAL | YEARS | HARVEST-AT-AGE | Pre-calculated WAA |
| :---: | :---: | :---: | :---: |
| NH Rec landings NH Rec discards | $\begin{aligned} & \hline 98-02 \\ & 98-02 \end{aligned}$ | supplied <br> supplied | used MA used MA |
| MA Rec landings MA Rec discards MA Com landings MA Com discards | $\begin{aligned} & 98-02 \\ & 98-02 \\ & 98-02 \\ & 98-02 \\ & \hline \end{aligned}$ | supplied <br> supplied <br> supplied <br> supplied | supplied <br> supplied <br> supplied <br> supplied |
| RI Com landings RI Rec landings RI Rec discards | $\begin{aligned} & 2002 \\ & 2002 \\ & 2002 \\ & \hline \end{aligned}$ | supplied supplied supplied | used MA used MA used MA |
| CT Rec landings CT Rec discards | $\begin{gathered} 98-02 \\ 98-00,02 \\ \hline \end{gathered}$ | GaryN CAA ${ }^{3}$ <br> GaryN CAA ${ }^{3}$ | used MA <br> used MA |
| NY all <br> NY Com landings NY Rec landings NY Rec discards | $\begin{aligned} & 98-00 \\ & 01-02 \\ & 01-02 \\ & 01-02 \\ & \hline \end{aligned}$ | 01,02 Ann. Rpts. 01,02 Ann. Rpts. 01,02 Ann. Rpts. | 01,02 Ann. Rpts. 01,02 Ann. Rpts. 01,02 Ann. Rpts. |
| NJ Rec landings NJ Rec discards NJ ALL | $\begin{aligned} & 98-01 \\ & 98-01 \\ & 2002 \\ & \hline \end{aligned}$ | $\%$ of harvest \#s ${ }^{1}$ supplied | $\%$ of harvest WAA ${ }^{2}$ supplied |
| Del Com landings Del Rec landings | $\begin{aligned} & 2002 \\ & 2002 \end{aligned}$ | GaryN CAA ${ }^{3}$ <br> GaryN CAA ${ }^{3}$ | used NY used NJ |
| MD Com landings MD Rec landings MD Rec discards | $\begin{aligned} & 98-02 \\ & 98-02 \\ & 98-02 \\ & \hline \end{aligned}$ | supplied | supplied |
| VA Com landings VA Rec landings VA Rec discards VA ALL | $\left\|\begin{array}{c} 98-00,02 \\ 98-00,02 \\ 98-00,02 \\ 2001 \end{array}\right\|$ | $\begin{aligned} & \text { GaryN CAA }{ }^{3} \\ & \text { GaryN CAA } \\ & \text { GaryN CAA } \\ & \text { GaryN CAA }{ }^{3} \\ & \hline \end{aligned}$ | used MD <br> used MD <br> used MD <br> used MD |

${ }^{1}$ (rec harvest-at-age)*(rec discard-at-age)/(total harvest)
${ }^{2}$ Ages 2-5: discard WAA $=0.8^{*}$ harvest WAA, Ages $6+$ : discard WAA $=0.9^{*}$ harvest WAA
${ }^{3}$ Coastwide summary CAA document supplied by Gary Nelson

Table 6. Removals Used to Calculate 1998-2002 Striped Bass Coastwide WAA.

| 1998 | 1999 | 2000 | 2001 | 2002 |
| :---: | :---: | :---: | :---: | :---: |
| NH Rec landings NH Rec discards | NH Rec landings NH Rec discards | NH Rec landings NH Rec discards | NH Rec landings NH Rec discards | NH Rec landings NH Rec discards |
| MA Rec landings MA Rec discards MA Com landings MA Com discards | MA Rec landings MA Rec discards MA Com landings MA Com discards | MA Rec landings MA Rec discards MA Com landings MA Com discards | MA Rec landings MA Rec discards MA Com landings MA Com discards | MA Rec landings MA Rec discards MA Com landings MA Com discards |
|  |  |  |  | RI Com landings RI Rec landings RI Rec discards |
| CT Rec landings CT Rec discards | CT Rec landings CT Rec discards | CT Rec landings CT Rec discards | CT Rec landings | CT Rec landings CT Rec discards |
| NY all | NY all | NY ALL | NY Com landings NY Rec landings NY Rec discards | NY Com landings NY Rec landings NY Rec discards |
| NJ Rec landings NJ Rec discards | NJ Rec landings NJ Rec discards | NJ Rec landings NJ Rec discards | NJ Rec landings NJ Rec discards | NJ ALL |
|  |  |  |  | Del Com landings Del Rec landings |
| MD Com landings MD Rec landings MD Rec discards | MD Com landings <br> MD Rec landings MD Rec discards | MD Com landings MD Rec landings MD Rec discards | MD Com landings MD Rec landings MD Rec discards | MD Com landings MD Rec landings MD Rec discards |
| VA Com landings VA Rec landings VA Rec discards | VA Com landings VA Rec landings VA Rec discards | VA Com landings VA Rec landings VA Rec discards | VA ALL | VA Com landings VA Rec landings VA Rec discards |

${ }^{1}$ (rec harvest-at-age)*(rec discard-at-age)/(total harvest)
${ }^{2}$ Ages 2-5: discard WAA $=0.8 *$ harvest WAA, Ages $6+$ : discard WAA $=0.9 *$ harvest WAA
${ }^{3}$ Coastwide summary CAA document supplied by Gary Nelson

Table 7. Comparison of "Old" and "New", or Recalculated Striped Bass Coastwide WAA (kg) for 2000-2003.

|  | YEAR AGE | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13/13+ | 14 | 15 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| OLD | 2000 | 0.14 | 1.05 | 1.47 | 2.32 | 3.23 | 4.52 | 6.39 | 7.11 | 7.81 | 9.2 | 9.31 | 10.1 | 11.36 | 12.45 | 17.3 |
|  | 2001 | 0.13 | 0.62 | 1.17 | 2.46 | 2.81 | 3.63 | 4.51 | 5.07 | 6.73 | 9.17 | 9.94 | 10.24 | 11.94 | 14.49 | 17.92 |
|  | 2002 | 0.82 | 0.81 | 1.25 | 1.75 | 2.47 | 3.3 | 4.16 | 5.48 | 6.36 | 7.45 | 8.75 | 8.89 | 9.99 | 11.03 | 13.95 |
|  | MEAN 00-02 | 0.36 | 0.83 | 1.30 | 2.18 | 2.84 | 3.82 | 5.02 | 5.89 | 6.97 | 8.61 | 9.33 | 9.74 | 11.10 | 12.66 | 16.39 |
|  | $\triangle$ 2002-2001 | 0.69 | 0.19 | 0.08 | -0.71 | -0.34 | -0.33 | -0.35 | 0.41 | -0.37 | -1.72 | -1.19 | -1.35 | -1.95 | -3.46 | -3.97 |
|  | $\Delta$ 2002-2000 | 0.68 | -0.24 | -0.22 | -0.57 | -0.76 | -1.22 | -2.23 | -1.63 | -1.45 | -1.75 | -0.56 | -1.21 | -1.37 | -1.42 | -3.35 |
| NEW | 2000 | 0.2 | 0.6 | 0.9 | 1.4 | 1.9 | 2.8 | 4 | 4.9 | 6.1 | 6 | 8.8 | 9.8 | 12.8 |  |  |
|  | 2001 | 0.1 | 0.4 | 0.8 | 1.7 | 2.2 | 3.2 | 4 | 5 | 5.9 | 7.2 | 8.1 | 7.4 | 10.6 |  |  |
|  | 2002 | 0.1 | 0.3 | 1.1 | 1.5 | 2.2 | 3.2 | 4.2 | 5.5 | 6.0 | 7.6 | 9.1 | 9.7 | 11.5 |  |  |
|  | 2003 | 0.1 | 0.6 | 1.0 | 1.4 | 2.2 | 3.2 | 4.1 | 5.2 | 6.1 | 7.2 | 8.5 | 9.4 | 11 |  |  |
| $\begin{gathered} \text { NEW } \\ \text { VS. } \\ \text { OLD } \end{gathered}$ | $\Delta$ 2000(N) - 2000(O) | 0.06 | -0.45 | -0.57 | -0.92 | -1.33 | -1.72 | -2.39 | -2.21 | -1.71 | -3.2 | -0.51 | -0.3 | 1.44 |  |  |
|  | $\Delta$ 2001(N) - 2001(O) | -0.03 | -0.22 | -0.37 | -0.76 | -0.61 | -0.43 | -0.51 | -0.07 | -0.83 | -1.97 | -1.84 | -2.84 | -1.34 |  |  |
|  | $\Delta$ 2002(N) - 2002(O) | -0.72 | -0.51 | -0.15 | -0.25 | -0.27 | -0.10 | 0.04 | 0.02 | -0.36 | 0.15 | 0.35 | 0.81 | -0.16 |  |  |
|  | MEAN 82-96 | 0.2 | 0.8 | 1.3 | 2.0 | 2.7 | 3.6 | 4.8 | 5.7 | 6.4 | 7.5 | 8.9 | 10.3 | 13.5 |  |  |
|  | $\Delta 2003$ - MEAN 82-96 | -0.07 | -0.17 | -0.29 | -0.58 | -0.48 | -0.43 | -0.7 | -0.53 | -0.3 | -0.32 | -0.41 | -0.94 | -2.5 |  |  |

Negative values emphasized by italics.

## Appendix E. A Coastwide MRFSS Index

## Methods

Generalized linear modelling (McCullagh and Nelder, 1989) was used to derive annual mean catch-per-hour estimates by adjusting the number of caught fish per trip for the classification variables of state, year, two-month sampling wave, number of days fished in the past 12 months (as a measure of avidity), and number of hours fished. In the analyses, I used only data from anglers who said they targeted striped bass to insure methods used among anglers are as consistent as possible and to identify those targeting anglers that did not catch striped bass (zero catches). Also, only data from private boats fishing in the Ocean during waves 3-6 from 1988 to 2004 were used.

A delta-lognormal model (Lo et al., 1992) was selected as the best approach to estimate year effects after examination of model dispersion (Terceiro, 2003) and standardized residual deviance versus linear predictor plots (McCullagh and Nelder, 1989). In the delta-lognormal model, catch data is decomposed into catch success/failure and positive catch components. Each component is analyzed separately using appropriate statistical techniques and then the statistical models are recombined to obtain estimates of the variable of interest. The catch success/failure was modelled as a binary response to the categorical variables using multiple logistic regression.

$$
\log i t(p)=\log (p / 1-p)=\alpha+\sum_{i=1}^{n} \beta_{i} X_{i}+\varepsilon
$$

where p is the probability of catching a fish, $\alpha$ is the intercept, $\beta_{\mathrm{i}}$ is the slope coefficient of the $i$ th factor, $X_{i}$ is the $i$ th categorical variable (coded as 0 or 1 ), and $\varepsilon$ is the error term. PROC LOGISTIC (SAS, 2000) was used to estimate parameters, and goodness-of-fit was assessed using concordance measures and the Hosmer-Lemeshow test (SAS, 2000).

Positive catches, transformed using the natural logarithm, were modelled assuming a normal error distribution using PROC GLM.

$$
\log (y)=\alpha+\sum_{i=1}^{n} \beta_{i} X_{i}+\varepsilon
$$

where $y$ is the observed positive catch, $\beta_{i}$, and $X_{i}$ are the same symbols as defined earlier, and $\varepsilon$ is the normal error term. Any variable not significant at $\alpha=0.05$ with type-III (partial) sum of squares was dropped from the initial GLM model and the analysis was repeated. First-order interactions were considered in the initial analyses but it was not always possible to generate annual means by the least-square methods with some interactions included (see Searle et al., 1980); therefore, only main effects were considered.

The annual index of striped bass releases was estimated by combining the two component models. The estimate in year $i$ from the models is given by

$$
\hat{\mathbf{I}}_{\mathbf{i}}=\hat{\mathbf{p}}_{\mathbf{i}} * \hat{\mathbf{y}}_{\mathbf{i}}
$$

where $p_{i}$ and $y_{i}$ are the predicted annual responses from the logistic and GLM. $p_{i}$ is

$$
\hat{\mathbf{p}}_{\mathbf{i}}=\frac{\exp \left(\hat{\alpha}+\hat{\beta}_{i}\right)}{1+\exp \left(\hat{\alpha}+\hat{\beta}_{i}\right)}
$$

calculated by
and $y_{i}$ is calculated by

$$
\hat{\mathbf{y}}_{i}=\exp \left(\mathrm{LSM}_{i}+\sigma^{2} / 2\right)
$$

where $\mathrm{LSM}_{i}$ is the least squares mean for year $i$ and $\sigma^{2}$ is the mean square error.

## Results

See Table 1 and 2 for the logistic and GLM outputs. Figure 1 is the index.

Gary Nelson
August 2005

Table 1. Logistic regression output .



Table 2. GLM output


Dependent Variable: I ogtot

| $\mathrm{Pr}>$ | Source | DF | Sum of Squares | Mean Square | F Value |
| :---: | :---: | :---: | :---: | :---: | :---: |
| <. 0001 | Model | 44 | 1090.02228 | 24.77323 | 32.74 |
|  | Error | 12037 | 9108.70262 | 0.75673 |  |
|  | Corrected Total | 12081 | 10198.72490 |  |  |
|  | R-Square | Coeff Var | Root MSE | Iogtot Mean |  |
|  | 0.106878 | 83.19786 | 0.869900 | 1. 045579 |  |
| Pr > F | Source | DF | Type I SS | Mean Square | F Value |
|  | ST | 10 | 258.0235556 | 25.8023556 | 34.10 |
| <. 0001 | YEAR | 16 | 155.8872713 | 9. 7429545 | 12.88 |
| <. 0001 | WAVE | 3 | 76.5444416 | 25.5148139 | 33.72 |
| <. 0001 |  |  |  |  |  |
| <. 0001 | FFDAYS12 | 14 | 288.3477562 | 20.5962683 | 27.22 |
| <. 0001 | NUM_HRSF | 1 | 311.2192537 | 311.2192537 | 411.27 |
| $\mathrm{Pr}>\mathrm{F}$ | Source | DF | Type III SS | Mean Square | F Value |
| <. 0001 | ST | 10 | 221.1009308 | 22.1100931 | 29.22 |
|  | YEAR | 16 | 150.8842719 | 9.4302670 | 12.46 |
| <. 0001 | WAVE | 3 | 73.8469641 | 24.6156547 | 32.53 |
| <. 0001 | FFDAYS12 | 14 | 291.5233657 | 20.8230976 | 27.52 |
| <. 0001 | NUM HRSF | 1 | 311.2192537 | 311.2192537 | 411.27 |
| <. 0001 |  |  |  |  |  |


|  | The GLM Procedure Least Squares Means |  |  |
| :---: | :---: | :---: | :---: |
| \| t | YEAR | logtot LSMEAN | $\begin{aligned} & \text { Standard } \\ & \text { Error } \quad \operatorname{Pr}> \end{aligned}$ |
|  | 1988 | 0.87115273 | 0.12364562 |
| <. 0001 | 1989 | 0.64504701 | 0.11751929 |
| <. 0001 | 1990 | 0.73978321 | 0.09831961 |
| <. 0001 |  |  |  |
|  | 1991 | 0.92486174 | 0.07810480 |
| <. 0001 | 1992 | 1. 22925291 | 0.07121558 |
| <. 0001 |  |  |  |
| <. 0001 | 1993 | 0.87196348 | 0.06831894 |
|  | 1994 | 0.99212787 | 0.06668443 |
| <. 0001 | 1995 | 1.10485998 | 0.06469854 |
| <. 0001 |  |  |  |
| <. 0001 | 1996 | 1.11781190 | 0.06316001 |
|  | 1997 | 1.15689303 | 0.06223858 |
|  | 1998 | 1.23720682 | 0.06133543 |
| <. 0001 | 1999 | 1.18550815 | 0.06266735 |
| <. 0001 |  |  |  |
| <. 0001 | 2000 | 1.13429729 | 0.06397751 |
| < 0001 | 2001 | 1. 07562958 | 0.06191667 |
| <. 0001 | 2002 | 0.94729993 | 0.06493991 |
| <. 0001 |  |  |  |
| <. 0001 | 2003 | 0.97453148 | 0.06407868 |
| 0 | 2004 | 1. 05500478 | 0.06371743 |



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McCullagh, P. and J. A. Nelder. 1989. Generalized linear models, 511 p. Chapman and Hall, London.

Terceiro, M. 2003. The statistical propoerties of recreational catch rate data for some fish stocks off the northeast US coast. Fish. Bull. 101: 653-672.


[^0]:    * Years when few or no striped bass were tagged and released.
    ** See footnote in Table 11.

[^1]:    * Years when few or no striped bass were tagged and released.
    ** See footnote in Table 13.

